

---

# Supplementary Material of “On the Effectiveness of Sampling for Evolutionary Optimization in Noisy Environments”

Chao Qian<sup>1,2</sup>

Yang Yu<sup>2</sup>

Ke Tang<sup>1</sup>

Yaochu Jin<sup>3</sup>

Xin Yao<sup>1,4</sup>

Zhi-Hua Zhou<sup>2\*</sup>

chaoqian@ustc.edu.cn

yuy@nju.edu.cn

ketang@ustc.edu.cn

yaochu.jin@surrey.ac.uk

x.yao@cs.bham.ac.uk

zhouzh@nju.edu.cn

<sup>1</sup>USTC-Birmingham Joint Research Institute in Intelligent Computation and Its Applications, School of Computer Science and Technology, University of Science and Technology of China, Hefei, 230027, China

<sup>2</sup>National Key Laboratory for Novel Software Technology, Nanjing University, Nanjing, 210023, China

<sup>3</sup>Department of Computer Science, University of Surrey, Guildford, GU2 7XH, UK

<sup>4</sup>Center of Excellence for Research in Computational Intelligence and Applications, School of Computer Science, University of Birmingham, Birmingham, B15 2TT, UK

---

## 1 Detailed Proofs

This document aims to provide the detailed proofs of Theorems 2 and 10, which are omitted in our original paper due to space limitations.

**Proof of Theorem 2.** We use Lemma 4 to prove this theorem. We first analyze  $p_{i,i+d}$  as that analyzed in the proof of Theorem 1. Note that for a solution  $x$ , the fitness value output by sampling with  $k = 2$  is  $\hat{f}(x) = (f_1^n(x) + f_2^n(x))/2$ , where  $f_1^n(x)$  and  $f_2^n(x)$  are noisy fitness values output by two independent fitness evaluations.

(1) When  $d \geq 3$ ,  $\hat{f}(x') \leq n - i - d + 1 \leq n - i - 2 < \hat{f}(x)$ . Thus, the offspring  $x'$  will be discarded, then we have  $\forall d \geq 3 : p_{i,i+d} = 0$ .

(2) When  $d = 2$ , the offspring solution  $x'$  will be accepted if and only if  $\hat{f}(x') = n - i - 1 = \hat{f}(x)$ . The probability of  $\hat{f}(x') = n - i - 1$  is  $(\frac{i+2}{n})^2$ , since it needs to always flip one 0-bit of  $x'$  in two noisy fitness evaluations. The probability of  $\hat{f}(x) = n - i - 1$  is  $(\frac{n-i}{n})^2$ , since it needs to always flip one 1-bit of  $x$ . Thus,  $p_{i,i+2} = P_2 \cdot (\frac{i+2}{n})^2 (\frac{n-i}{n})^2$ .

(3) When  $d = 1$ , there are three possible cases for the acceptance of  $x'$ :  $\hat{f}(x') = n - i \wedge \hat{f}(x) = n - i - 1$ ,  $\hat{f}(x') = n - i \wedge \hat{f}(x) = n - i$  and  $\hat{f}(x') = n - i - 1 \wedge \hat{f}(x) = n - i - 1$ . The probability of  $\hat{f}(x') = n - i$  is  $(\frac{i+1}{n})^2$ , since it needs to always flip one 0-bit of  $x$  in two noisy evaluations. The probability of  $\hat{f}(x') = n - i - 1$  is  $2 \frac{i+1}{n} \frac{n-i-1}{n}$ , since it needs to flip one 0-bit of  $x$  in one noisy evaluation and flip one 1-bit in the other

---

\* Corresponding author

noisy evaluation. Similarly, we can derive that the probabilities of  $\hat{f}(x) = n - i - 1$  and  $\hat{f}(x) = n - i$  are  $\left(\frac{n-i}{n}\right)^2$  and  $2\frac{n-i}{n}\frac{i}{n}$ , respectively. Thus,  $p_{i,i+1} = P_1 \cdot \left(\left(\frac{i+1}{n}\right)^2\left(\frac{n-i}{n}\right)^2 + 2\frac{n-i}{n}\frac{i}{n}\right) + 2\frac{i+1}{n}\frac{n-i-1}{n}\left(\frac{n-i}{n}\right)^2$ .

(4) When  $d = -1$ ,  $x'$  will be rejected if and only if  $\hat{f}(x') = n - i \wedge \hat{f}(x) = n - i + 1$ . The probability of  $\hat{f}(x') = n - i$  is  $\left(\frac{n-i+1}{n}\right)^2$ , since it needs to always flip one 1-bit of  $x'$  in two noisy evaluations. The probability of  $\hat{f}(x) = n - i + 1$  is  $\left(\frac{i}{n}\right)^2$ , since it needs to always flip one 0-bit of  $x$ . Thus,  $p_{i,i-1} = P_{-1} \cdot \left(1 - \left(\frac{n-i+1}{n}\right)^2\left(\frac{i}{n}\right)^2\right)$ .

(5) When  $d \leq -2$ ,  $\hat{f}(x') \geq n - i - d - 1 \geq n - i + 1 \geq \hat{f}(x)$ . Thus, the offspring  $x'$  will always be accepted, then we have  $\forall d \leq -2 : p_{i,i+d} = P_d$ .

Using these probabilities, we have

$$\begin{aligned} \mathbb{E}[X_t - X_{t+1} \mid X_t = i] &= \sum_{d=1}^i d \cdot p_{i,i-d} - \sum_{d=1}^{n-i} d \cdot p_{i,i+d} \\ &= \left(1 - \left(\frac{n-i+1}{n}\right)^2 \left(\frac{i}{n}\right)^2\right) P_{-1} + \sum_{d=2}^i d P_{-d} - 2 \left(\frac{i+2}{n}\right)^2 \left(\frac{n-i}{n}\right)^2 P_2 \\ &\quad - \left(\left(\frac{i+1}{n}\right)^2 \left(\left(\frac{n-i}{n}\right)^2 + 2\frac{n-i}{n}\frac{i}{n}\right) + 2\frac{i+1}{n}\frac{n-i-1}{n}\left(\frac{n-i}{n}\right)^2\right) P_1 \\ &\leq \left(1 - \left(\frac{n-i+1}{n}\right)^2 \left(\frac{i}{n}\right)^2\right) \frac{i}{n} \left(1 - \frac{1}{n}\right)^{n-1} \cdot 1.14 + \frac{i}{n} \left(\left(1 + \frac{1}{n}\right)^{i-1} - 1\right) \\ &\quad - 2 \left(\frac{i+2}{n}\right)^2 \left(\frac{n-i}{n}\right)^2 \frac{(n-i)(n-i-1)}{2n^2} \left(1 - \frac{1}{n}\right)^{n-2} - \frac{n-i}{n} \left(1 - \frac{1}{n}\right)^{n-1} \\ &\quad \cdot \left(\left(\frac{i+1}{n}\right)^2 \left(\left(\frac{n-i}{n}\right)^2 + 2\frac{n-i}{n}\frac{i}{n}\right) + 2\frac{i+1}{n}\frac{n-i-1}{n}\left(\frac{n-i}{n}\right)^2\right) \\ &\quad \text{(by using the bounds of } P_d \text{ in the proof of Theorem 1)} \\ &\leq \frac{i}{n} \left(1 - \frac{1}{n}\right)^{n-1} (1.14 - 2) + O\left(\left(\frac{i}{n}\right)^2\right) \quad \text{(since } i < n^{1/4}\text{)} \\ &\leq -0.3 \cdot \frac{i}{n} + O\left(\left(\frac{i}{n}\right)^2\right). \quad \text{(by } \left(1 - \frac{1}{n}\right)^{n-1} \geq \frac{1}{e}\text{)} \end{aligned}$$

It is also easy to verify that  $P(X_{t+1} \neq i \mid X_t = i) = \Theta\left(\frac{i}{n}\right)$  for  $1 \leq i < n^{1/4}$ . Thus,  $\mathbb{E}[X_t - X_{t+1} \mid X_t = i] = -\Omega(P(X_{t+1} \neq i \mid X_t = i))$ , which implies that condition 1 of Lemma 4 holds.

Condition 2 of Lemma 4 still holds with  $\delta = 1$  and  $r(l) = \frac{32e}{7}$ . The analysis procedure is the same as that in the proof of Theorem 1, because the following inequality holds:

$$\begin{aligned} P(|X_{t+1} - X_t| \geq 1 \mid X_t = i) &\geq p_{i,i-1} = \left(1 - \left(\frac{n-i+1}{n}\right)^2 \left(\frac{i}{n}\right)^2\right) \cdot P_{-1} \\ &\geq \left(1 - \frac{n-i+1}{n} \frac{i}{n}\right) \cdot P_{-1}. \end{aligned}$$

Thus, by Lemma 4, the expected running time is exponential.  $\square$

**Proof of Theorem 10.** We use Lemma 2 to prove this theorem. The proof is very similar to that of Theorem 3 except that the probabilities  $p_{i,i+d}$  are different due to the difference on the noise and the value of  $k$ .

We use the distance function  $V(x) = |x|_0$ . Let  $i$  (where  $1 \leq i \leq n$ ) denote the number of 0-bits of the current solution  $x$ . Let  $p_{i,i+d}$  be the probability that the next solution after mutation and selection has  $i+d$  number of 0-bits (where  $-i \leq d \leq n-i$ ). Thus,

$$\mathbb{E}[V(\xi_t) - V(\xi_{t+1}) \mid \xi_t = x] = \sum_{d=1}^i d \cdot p_{i,i-d} - \sum_{d=1}^{n-i} d \cdot p_{i,i+d}. \quad (1)$$

We then analyze  $p_{i,i+d}$  ( $1 \leq i \leq n$ ). For a solution  $x$ , the fitness value output by sampling is the average of noisy fitness values by  $k$  independent evaluations, i.e.,  $\hat{f}(x) = \sum_{i=1}^k f_i^n(x)/k$ . Note that for the flipping in asymmetric one-bit noise, the probability of flipping a 0 (or 1) bit is different when  $|x|_0 = 0$ ,  $|x|_0 = n$  and  $0 < |x|_0 < n$ . In these three cases, the probabilities of flipping a 0 bit are 0, 1 and  $\frac{1}{2}$ , respectively; the probabilities of flipping a 1 bit are 1, 0 and  $\frac{1}{2}$ , respectively. Thus, the analysis of  $p_{i,i+d}$  will also be separated into several cases if necessary. Let  $P_d$  denote the probability that the offspring solution  $x'$  generated by mutation has  $i+d$  number of 0-bits.

(1) When  $d \geq 3$ ,  $\hat{f}(x') \leq n - i - d + 1 \leq n - i - 2 < \hat{f}(x)$ . Thus, the offspring  $x'$  will be discarded, then  $\forall d \geq 3 : p_{i,i+d} = 0$ .

(2) When  $d = 2$ ,  $x'$  will be accepted if and only if  $\hat{f}(x') = n - i - 1 = \hat{f}(x)$ , that is, it needs to always flip one 0-bit of  $x'$  and flip one 1-bit of  $x$  in  $k$  noisy fitness evaluations. We then consider three cases:

- $i = n$  or  $n - 1$ . It trivially holds that  $p_{i,i+2} = 0$ .
- $i = n - 2$ . Note that  $|x'|_0 = i + 2 = n$ , thus the probability of flipping a 0 bit of  $x'$  in noisy evaluation is 1. Then, we have  $p_{i,i+2} = P_2 \cdot 1^k \cdot \frac{1}{2^k}$ .
- $1 \leq i < n - 2$ . We have  $p_{i,i+2} = P_2 \cdot \frac{1}{2^k} \cdot \frac{1}{2^k}$ .

(3) When  $d = 1$ , there are two possible values for  $f^n(x')$ :  $n - i - 2$  or  $n - i$ . Similarly,  $f^n(x) = n - i - 1$  or  $n - i + 1$ . In the  $k$  independent noisy evaluations for  $x'$ , let  $k_1 \in [0, k]$  denote the number of times that  $f^n(x') = n - i$ . Similarly, let  $k_2 \in [0, k]$  denote the number of times that  $f^n(x) = n - i - 1$ . The condition for the acceptance of  $x'$  is  $\hat{f}(x') \geq \hat{f}(x)$ , which can be simplified as follows.

$$\begin{aligned} \hat{f}(x') \geq \hat{f}(x) &\Leftrightarrow \sum_{i=1}^k f_i^n(x') \geq \sum_{i=1}^k f_i^n(x) \\ &\Leftrightarrow k_1(n-i) + (k-k_1)(n-i-2) \geq k_2(n-i-1) + (k-k_2)(n-i+1) \\ &\Leftrightarrow k_1 + k_2 \geq \frac{3}{2}k. \end{aligned}$$

We then consider three cases:

- $i = n$ . It trivially holds that  $p_{i,i+1} = 0$ .
- $i = n - 1$ . Note that  $|x'|_0 = i + 1 = n$ , thus the probability of flipping a 0 bit of  $x'$  in noisy evaluation (i.e.,  $f^n(x') = n - i$ ) is 1, which implies that  $k_1 = k$ . Thus, the condition of accepting  $x'$  changes to be  $k_2 \geq \frac{k}{2}$ . Then, we have  $p_{i,i+1} = P_1 \cdot \sum_{k_2 \geq \frac{k}{2}} \binom{k}{k_2} \frac{1}{2^k}$ .
- $1 \leq i < n - 1$ . We have  $p_{i,i+1} = P_1 \cdot \sum_{k_1+k_2 \geq \frac{3}{2}k} \binom{k}{k_1} \frac{1}{2^k} \cdot \binom{k}{k_2} \frac{1}{2^k} = P_1 \cdot \sum_{k' \geq \frac{3}{2}k} \binom{2k}{k'} \frac{1}{2^{2k}}$ .

(4) When  $d = -1$ ,  $f^n(x') = n - i$  or  $n - i + 2$ ;  $f^n(x) = n - i - 1$  or  $n - i + 1$ . In the  $k$  independent noisy evaluations for  $x'$ , let  $k_1 \in [0, k]$  denote the number of times that  $f^n(x') = n - i$ . Similarly, let  $k_2 \in [0, k]$  denote the number of times that  $f^n(x) = n - i + 1$ . The condition for the rejection of  $x'$  is  $\hat{f}(x') < \hat{f}(x)$ , which can be simplified as follows.

$$\begin{aligned} \hat{f}(x') < \hat{f}(x) &\Leftrightarrow \sum_{i=1}^k f_i^n(x') < \sum_{i=1}^k f_i^n(x) \\ &\Leftrightarrow k_1(n-i) + (k-k_1)(n-i+2) < k_2(n-i+1) + (k-k_2)(n-i-1) \\ &\Leftrightarrow k_1 + k_2 > \frac{3}{2}k. \end{aligned}$$

We then consider three cases:

- $i = n$ . Note that the probability of flipping a 0 bit of  $x$  in noisy evaluation (i.e.,  $f^n(x) = n - i + 1$ ) is 1, which implies that  $k_2 = k$ . Thus, the condition of rejecting  $x'$  changes to be  $k_1 > \frac{k}{2}$ . Then, we have  $p_{i,i-1} = P_{-1} \cdot (1 - \sum_{k_1 > \frac{k}{2}} \binom{k}{k_1} \frac{1}{2^k})$ .
- $i = 1$ . Note that  $|x'|_0 = i - 1 = 0$ , thus the probability of flipping a 1 bit of  $x'$  in noise (i.e.,  $f^n(x') = n - i$ ) is 1, which implies that  $k_1 = k$ . Thus, the condition of rejecting  $x'$  changes to be  $k_2 > \frac{k}{2}$ . Then, we have  $p_{i,i-1} = P_{-1} \cdot (1 - \sum_{k_2 > \frac{k}{2}} \binom{k}{k_2} \frac{1}{2^k})$ .
- $1 < i < n$ .  $p_{i,i-1} = P_{-1} \cdot (1 - \sum_{k_1+k_2 > \frac{3}{2}k} \binom{k}{k_1} \frac{1}{2^k} \cdot \binom{k}{k_2} \frac{1}{2^k}) = P_{-1} \cdot (1 - \sum_{k' > \frac{3}{2}k} \binom{2k}{k'} \frac{1}{2^{2k}})$ .

By combining the above three cases, we can easily derive that  $p_{i,i-1} \geq P_{-1} \cdot \frac{1}{2}$ .

(5) When  $d \leq -2$ ,  $\hat{f}(x') \geq n - i - d - 1 \geq n - i + 1 \geq \hat{f}(x)$ . Thus,  $x'$  will always be accepted, then we have  $\forall d \leq -2 : p_{i,i+d} = P_d$ .

By applying these probabilities to Eq. (1), we have

$$\mathbb{E}[V(\xi_t) - V(\xi_{t+1}) \mid \xi_t = x] \geq p_{i,i-1} - p_{i,i+1} - 2 \cdot p_{i,i+2}. \quad (2)$$

We then analyze Eq. (2) in three cases.

(1) When  $i = n$ ,  $p_{i,i+2} = 0$  and  $p_{i,i+1} = 0$ . Thus, we have

$$\mathbb{E}[V(\xi_t) - V(\xi_{t+1}) \mid \xi_t = x] \geq P_{-1} \cdot \frac{1}{2} \geq \frac{i}{n} \left(1 - \frac{1}{n}\right)^{n-1} \cdot \frac{1}{2} \geq \frac{i}{2en}.$$

(2) When  $i = n - 1$ ,  $p_{i,i+2} = 0$  and  $p_{i,i+1} = P_1 \cdot \sum_{k_2 \geq \frac{k}{2}} \binom{k}{k_2} \frac{1}{2^k} < P_1 \leq \frac{n-i}{n} = \frac{1}{n}$ . Thus,

$$\mathbb{E}[V(\xi_t) - V(\xi_{t+1}) \mid \xi_t = x] \geq P_{-1} \cdot \frac{1}{2} - P_1 \geq \frac{i}{2en} - \frac{1}{n} \geq 0.01 \cdot \frac{i}{n},$$

where the last inequality holds with  $n \geq 7$ .

(3) When  $1 \leq i < n - 1$ ,  $p_{i,i+2} \leq P_2 \cdot \frac{1}{2^k}$  and  $p_{i,i+1} = P_1 \cdot \sum_{k' \geq \frac{3}{2}k} \binom{2k}{k'} \frac{1}{2^{2k}}$ . Let  $X_i$  ( $1 \leq i \leq 2k$ ) be independent random variables such that  $P(X_i = 1) = \frac{1}{2}$  and  $P(X_i = 0) = \frac{1}{2}$ . Then,  $\sum_{k' \geq \frac{3}{2}k} \binom{2k}{k'} \frac{1}{2^{2k}} = P(\sum_{i=1}^{2k} X_i \geq \frac{3}{2}k) \leq e^{-\frac{k}{12}}$ , where the “ $\leq$ ” is by Chernoff’s inequality. Thus, we have

$$\begin{aligned} \mathbb{E}[V(\xi_t) - V(\xi_{t+1}) \mid \xi_t = x] &\geq \frac{P_{-1}}{2} - P_1 \cdot e^{-\frac{k}{12}} - 2 \cdot \frac{P_2}{2^k} \\ &\geq \frac{i}{2en} - \frac{1}{n^2} - \frac{1}{n^{12}} \geq 0.01 \cdot \frac{i}{n}, \end{aligned}$$

where the second inequality is by  $k = \lceil 24 \log n \rceil$  (note that  $\log$  corresponds to the natural logarithm, i.e., the base is  $e$ ), and the last inequality holds with  $n \geq 6$ .

Thus, the condition of Lemma 2 holds with  $\mathbb{E}[V(\xi_t) - V(\xi_{t+1}) \mid \xi_t = x] \geq \frac{0.01}{n} \cdot V(x)$ . We then have

$$\mathbb{E}[\tau \mid \xi_0] \leq \frac{n}{0.01} \cdot (1 + \log V(\xi_0)) \in O(n \log n),$$

i.e., the expected iterations for finding the optimal solution is upper bounded by  $O(n \log n)$ . Because the cost of each iteration is  $2k = 2 \cdot \lceil 24 \log n \rceil$ , the expected running time is  $O(n \log^2 n)$ .  $\square$