# Analyzing Evolutionary Optimization in Noisy Environments 

Chao Qian<br>Yang Yu*<br>Zhi-Hua Zhou<br>qianc@lamda.nju.edu.cn<br>yuy@lamda.nju.edu.cn<br>National Key Laboratory for Novel Software Technology, Nanjing University, Nanjing, 210023, China


#### Abstract

Many optimization tasks must be handled in noisy environments, where the exact evaluation of a solution cannot be obtained, only a noisy one. For optimization of noisy tasks, evolutionary algorithms (EAs), a type of stochastic metaheuristic search algorithm, have been widely and successfully applied. Previous work mainly focuses on the empirical study and design of EAs for optimization under noisy conditions, while the theoretical understandings are largely insufficient. In this study, we firstly investigate how noisy fitness can affect the running time of EAs. Two kinds of noisehelpful problems are identified, on which the EAs will run faster with the presence of noise, and thus the noise should not be handled. Secondly, on a representative noiseharmful problem in which the noise has a strong negative effect, we examine two commonly employed mechanisms dealing with noise in EAs: the re-evaluation and the threshold selection strategies. The analysis discloses that using these two strategies simultaneously is effective for the one-bit noise, but ineffective for the asymmetric one-bit noise. The smooth threshold selection is then proposed, which can be proven as an effective strategy to further improve the noise tolerance ability in the problem. We then complement the theoretical analysis by experiments on both synthetic problems as well as two combinatorial problems, the minimum spanning tree and the maximum matching. The experimental results agree with the theoretical findings, and also show that the proposed smooth threshold selection can deal with the noise better.


## Keywords

Noisy optimization, evolutionary algorithms, re-evaluation, threshold selection, running time, computational complexity.

## 1 Introduction

Optimization tasks often encounter noisy environments. For example, in industrial design such as VLSI design (Guo et al. 2014), every prototype is evaluated by simulations; therefore, the result of the evaluation may not be perfect due to the simulation error. Also, with machine learning, a prediction model is evaluated only on a limited amount of data (Qian et al. 2015a); therefore, the estimated performance is shifted from the true performance. It is possible that noisy environments change the properties of an optimization problem, thus traditional optimization techniques may have low efficacy. Meanwhile, evolutionary algorithms (EAs) Bäck 1996) have been widely

[^0]and successfully adopted for noisy optimization tasks (Freitas, 2003; Ma et al. 2006; Chang and Chen, 2006).

EAs are a type of randomized metaheuristic optimization algorithm, inspired by natural phenomena including evolution of species, swarm cooperation, immune systems, and others. EAs typically involve a cycle of three stages: a reproduction stage that produces new solutions based on the currently maintained solutions; an evaluation stage that evaluates the newly generated solutions; and a selection stage that wipes out bad solutions. The concept of using EAs for noisy optimization is that the corresponding natural phenomena have been successfully processed in noisy natural environments, and hence the algorithmic simulations are also likely to be able to handle noise.

On one hand, it is believed that noise makes the optimization harder, and thus handling mechanisms have been proposed to reduce the negative effect of the noise (Fitzpatrick and Grefenstette 1988, Beyer 2000 Arnold and Beyer, 2003). Two representative strategies are the re-evaluation strategy and the threshold selection. According to the re-evaluation strategy (Jin and Branke, 2005, Goh and Tan, 2007, Doerr et al. 2012a), whenever the fitness (also called the cost or objective value) of a solution is required, EAs make an independent evaluation of the solution regardless of whether the solution has been evaluated before, such that the fitness is smoothed. According to the threshold selection strategy (Markon et al., 2001; Bartz-Beielstein and Markon, 2002, Bartz-Beielstein, 2005a), in the selection stage, EAs accept a newly generated solution only if its fitness is larger than the fitness of the old solution by at least a threshold value $\tau$, such that the risk of accepting a bad solution due to noise is reduced.

On the other hand, several empirical observations have shown cases where noise can have a positive impact on the performance of local search Selman et al., 1994; Hoos and Stützle, 2000, 2005, which indicates that noise does not always have a negative impact.

As these previous studies are mainly empirical, theoretical analysis is in need for a better understanding of evolutionary optimization in noisy environments.

### 1.1 Related Work

Despite their wide and successful application, the theoretical analysis of EAs on noisy optimization is rare. Recently, numerous theoretical results on EAs have emerged (e.g., Neumann and Witt 2010 Auger and Doerr, 2011); however, most of them focus on clean environments. In noisy environments, the optimization is more complex and more randomized, thus the theoretical analysis is difficult.

Only a few theoretical analyses for EAs on noisy optimization have been published. Gutjahr (2003, 2004) first analyzed the Ant Colony Optimization (ACO) algorithm for stochastic combinatorial optimization and proved convergence under mild conditions. Then, Droste (2004) gave a running time analysis of EAs in discrete noisy optimization for the first time. Droste analyzed the ( $1+1$ )-EA on the OneMax problem under one-bit noise and showed the maximal noise strength $\log (n) / n$ allowing a polynomial running time, where the noise strength is characterized by the noise probability in $[0,1]$ and $n$ is the problem size. Later, Sudholt and Thyssen (2012) analyzed the running time of a simple ACO for stochastic shortest path problems where edge weights are subject to noise, and showed the ability and limitation of the ACO under various noise models. For the difficulty faced by an ACO under a specific noise model, Doerr et al. (2012a) further showed that the re-evaluation strategy can overcome it,
i.e., avoid being misled by an exceptionally optimistic evaluation due to noise. Recently, Qian et al. (2014) have investigated the effectiveness of sampling, a common strategy to reduce the effect of noise. They proved a sufficient condition under which sampling is useless (i.e., sampling increases the running time), and applied it to show that sampling is useless for the $(1+1)$-EA optimizing the OneMax and the Trap problem under additive Gaussian noise.

### 1.2 Our Contribution

In this paper, we study the effect of noise to EAs, and investigate the noise handling mechanisms when noise needs to be accounted for.

Firstly, the effect of noise on the expected running time of EAs is investigated in Section 3. On deceptive and flat problems, we will prove that noise can simplify the optimization (i.e., decreases the expected running time) for EAs. The analysis results support that for some difficult problems, handling the noise is not necessary.

Secondly, in Section 4, on the OneMax problem which will be proved to be negatively affected by noise in Section 4.1, two commonly employed noise handling mechanisms are examined in Section 4.2: the re-evaluation and the threshold selection strategies. With the $(1+1)$-EA under one-bit noise, the noise handling mechanisms are evaluated by the polynomial noise tolerance (PNT), which is the range of the noise strength such that the expected running time of the algorithm is polynomial. The wider the PNT is, the better a noise handling mechanism is. For example, the one-bit noise strength (and thus the PNT) is characterized by the noise probability $p_{n} \in[0,1]$. The configurations of the ( $1+1$ )-EA that we analyzed include: without any noise handling strategy (abbreviated as single-evaluation), single-evaluation with threshold selection (abbreviated as single-evaluation with a value of $\tau$ ), and re-evaluation with threshold selection (abbreviated as re-evaluation with a value of $\tau$ ). Their PNTs are presented in Table 1 , where the PNT of the (1+1)-EA with re-evaluation (but no threshold selection) is directly derived from (Droste, 2004). The comparison shows that,

1. the re-evaluation alone makes the PNT much worse than that of the singleevaluation.
2. a threshold selection must be combined with the re-evaluation, otherwise, the EA could not tolerate any noise strength larger than 0 ; meanwhile the re-evaluation can also be better if used with the threshold selection.
3. the re-evaluation with threshold selection strategy (threshold $=1$ ) can improve upon that of the single-evaluation.
Afterwards, in Section 4.3, we disclose a weakness of the above noise handling

| Noise handling strategies | PNT |
| :---: | :---: |
| single-evaluation | $[0,1-\overline{\Theta(\text { pooly(n) })}]$ |
| single-evaluation $\& \tau>0$ | $[0,0]$ |
| re-evaluation | $\left[0, \Theta\left(\frac{\log n}{n}\right)\right]($ Droste, 2004$]$ |
| re-evaluation \& $\tau=1$ |  |
| re-evaluation \& $\tau=2$ |  |
| re-evaluation \& $\tau>2$ | $\left[\frac{1}{\Theta(\text { poly(n) })}, 1-\overline{\Theta(\text { poly }(n))}\right]$ |

Table 1: The PNT with respect to one-bit noise of the ( $1+1$ )-EA using different noise handling strategies on the OneMax problem.
mechanisms: when used with the ( $1+1$ )-EA solving the OneMax problem under asymmetric one-bit noise, all of them are ineffective (i.e., need exponential running time) when the noise probability reaches 1 . The reason of the ineffectiveness of reevaluation with threshold selection is because it has a too large probability of accepting false progresses caused by the noise when the threshold $\tau \leq 1$, it has a too small probability of accepting true progresses when $\tau \geq 2$, and setting $\tau$ between 1 and 2 is useless due to the minimum fitness gap 1 (i.e., a value of $\tau \in(1,2)$ is equivalent to $\tau=2$ ). We then introduce a modification into the threshold selection strategy to turn the original hard threshold into the smooth threshold, which allows a fractional threshold to be effective. We prove that with the smooth threshold selection strategy the PNT can be $[0,1]$, i.e., the $(1+1)$-EA is always a polynomial algorithm on the problem regardless of the noise probability.

Finally, we conducted experiments to verify and complement the theoretical results in Section 5. Firstly, we show using two problem classes, the Jump problem which is a synthetic problem, and the minimum spanning tree problem which is a common combinatorial problem, that the badness of the noise is negatively correlated with the hardness of the problem, which was previously not noticed. Therefore, when the problem is quite hard, the noise can be helpful and thus handling the noise is not necessary. Then we verify that the smooth threshold selection can better handle the noise by experiments on the maximum matching problem. Section 6 concludes the paper.

## 2 Preliminaries

### 2.1 Noisy Optimization

A general optimization problem can be represented as arg $\max _{x} f(x)$, where the objective $f$ is also called fitness in the context of evolutionary computation. In realworld optimization tasks, the fitness evaluation for a solution is usually disturbed by noise, and consequently we cannot obtain the exact fitness value but only a noisy one. Let $f^{N}(x)$ and $f(x)$ denote the noisy and true fitness of a solution $x$, respectively. In this study, we will use the following three widely investigated noise models.
[additive]: $f^{N}(x)=f(x)+\delta$, where $\delta$ is uniformly selected from $\left[\delta_{1}, \delta_{2}\right]$ at random.
[multiplicative]: $f^{N}(x)=f(x) \cdot \delta$, where $\delta$ is uniformly randomly selected from $\left[\delta_{1}, \delta_{2}\right]$. [one-bit]: $f^{N}(x)=f(x)$ with probability $\left(1-p_{n}\right)\left(p_{n} \in[0,1]\right)$; otherwise, $f^{N}(x)=f\left(x^{\prime}\right)$, where $x^{\prime}$ is generated by flipping a uniformly randomly chosen bit of $x \in\{0,1\}^{n}$. This noise is for problems where solutions are represented in binary strings.
Additive and multiplicative noise have often been used to analyze the effect of noise (Beyer 2000; Jin and Branke 2005). One-bit noise is specifically used for optimizing pseudo-Boolean problems over $\{0,1\}^{n}$, and has been investigated in the first work for analyzing the running time of EAs in noisy optimization (Droste, 2004) and used to understand the role of noise in stochastic local search (Selman et al., 1994, Hoos and Stützle, 1999: Mengshoel, 2008).

Besides the above noises, we also consider a variant of one-bit noise called asymmetric one-bit noise, in Definition 1. Inspired from the asymmetric mutation operator (Jansen and Sudholt 2010), the asymmetric one-bit noise flips a specific bit position with probability depending on the number of bit positions that take the same value. For the flipping of asymmetric one-bit noise on a solution $x \in\{0,1\}^{n}$, the probability of flipping a specific 0 bit is $\frac{1}{2} \cdot \frac{1}{|x|_{0}}$, and the probability of flipping a spe-
cific 1 bit is $\frac{1}{2} \cdot \frac{1}{n-\mid x x_{0}}$, where $|x|_{0}=n-\sum_{i=1}^{n} x_{i}$ is the number of 0 bits of $x$. Note that for one-bit noise, the probability of flipping any specific bit is $\frac{1}{n}$. For both one-bit and asymmetric one-bit noise, $p_{n}$ controls the noise strength. In this paper, we assume that the parameters of the environment (i.e., $p_{n}, \delta_{1}$ and $\delta_{2}$ ) do not change over time.
Definition 1 (Asymmetric One-bit Noise). Given a fitness function $f$ and a solution $x \in\{0,1\}^{n}$, the asymmetric one-bit noise with a parameter $p_{n} \in[0,1]$ leads to a noisy fitness value $f^{N}(x)$ as $f^{N}(x)=f(x)$ with probability $\left(1-p_{n}\right)$, otherwise $f^{N}(x)=f\left(x^{\prime}\right)$, where $x^{\prime}$ is generated by flipping the $j$-th bit of $x$, and $j$ is a uniformly randomly chosen position of

$$
\begin{cases}\text { all bits of } x, & \text { if }|x|_{0}=0 \text { or } n ; \\ \begin{cases}0 \text { bits of } x, & \text { with probability } 1 / 2 ; \\ 1 \text { bits of } x, & \text { with probability } 1 / 2 .\end{cases} & \text { otherwise. }\end{cases}
$$

It is possible that a large noise could make an optimization problem extremely hard for particular algorithms. We are thus interested in the noise strength, under which an algorithm could be "tolerant" to have a polynomial running time. The noise strength can be measured by the adjustable parameters, e.g., $\delta_{1}, \delta_{2}$ for the additive and multiplicative noise, and $p_{n}$ for the one-bit noise. We denote $g_{\boldsymbol{\theta}}(f)$ as a type of noisy fitness which disturbs the original fitness function $f$ by the noise with parameter $\boldsymbol{\theta}$ (where $\boldsymbol{\theta}$ can be a tuple, e.g., $\boldsymbol{\theta}=\left(\delta_{1}, \delta_{2}\right)$ for additive noise), and define the PNT in Definition 2, which characterizes the maximum range of the noise parameter for allowing a polynomial expected running time. Note that, the PNT is $\emptyset$ if the algorithm never has a polynomial expected running time for any noise strength. We will study the PNT of EAs in order to analyze the effectiveness of noise handling strategies.
Definition 2 (Polynomial Noise Tolerance (PNT)). For an algorithm $\mathcal{A}$ running on a problem $f$ with a type of noise $g_{\theta}$, let ERT $\left(\mathcal{A} ; g_{\theta}(f)\right)$ be the expected running time of $\mathcal{A}$ on $f$ with noise strength represented by the parameter $\theta$. Then, the polynomial noise tolerance of $\mathcal{A}$ on $f$ with the type of noise $g_{\theta}$ is the range of the noise strength in which the expected running time is polynomial to the problem size n, i.e.,

$$
\operatorname{PNT}\left(\mathcal{A} ; f, g_{\boldsymbol{\theta}}\right)=\left\{\boldsymbol{\theta} \mid E R T\left(\mathcal{A} ; g_{\boldsymbol{\theta}}(f)\right)=\operatorname{poly}(n)\right\} .
$$

### 2.2 Evolutionary Algorithms

Evolutionary algorithms (EAs) Bäck 1996) are a type of population-based metaheuristic optimization algorithm. Although many variants exist, the common procedure of EAs can be described as follows:

1. Generate an initial set of solutions (called a population);
2. Reproduce new solutions from the current population;
3. Evaluate the newly generated solutions;
4. Update the population by removing the bad solutions;
5. Repeat steps $2-5$ until a specific criterion is met.

The $(1+1)-E A$, as in Algorithm 1 is a simple EA for maximizing pseudo-Boolean problems over $\{0,1\}^{n}$, which reflects the common structure of EAs. It maintains only one solution, and repeatedly improves the current solution by using bit-wise mutation (i.e., the 3rd step of Algorithm 1 ). It has been widely used for the running time analysis of EAs, e.g., in (He and Yao 2001 Droste et al. 2002).
Algorithm 1 ((1+1)-EA). Given pseudo-Boolean function $f$ with solution length $n$, it consists of the following steps:

1. $x:=$ randomly selected from $\{0,1\}^{n}$.
2. Repeat until the termination condition is met
3. $x^{\prime}:=$ flip each bit of $x$ independently with probability $p$.
4. if $f\left(x^{\prime}\right) \geq f(x)$
5. $\quad x:=x^{\prime}$.
where $p \in(0,0.5)$ is the mutation probability.
The $(1+\lambda)-E A$, as in Algorithm2, applies an offspring population size $\lambda$. In each iteration, it first generates $\lambda$ offspring solutions by independently mutating the current solution $\lambda$ times, and then selects the best from the current solution and the offspring solutions as the next solution. It has been used to disclose the effect of offspring population size by running time analysis (Jansen et al. 2005, Neumann and Wegener, 2007). Note that the $(1+1)-E A$ is a special case of the $(1+\lambda)$-EA with $\lambda=1$.
```
Algorithm 2 ((1+\lambda)-EA). Given pseudo-Boolean function }f\mathrm{ with solution length n, it
consists of the following steps:
    x:= randomly selected from {0,1} .
    Repeat until the termination condition is met
    i:=1.
    Repeat until i> \lambda.
            x}:=\mathrm{ flip each bit of }x\mathrm{ independently with probability p.
            i:=i+1.
        if max{f(\mp@subsup{x}{1}{}),\ldots,f(\mp@subsup{x}{\lambda}{})}\geqf(x)
            x= arg max }\mp@subsup{x}{\mp@subsup{x}{}{\prime}\in{\mp@subsup{x}{1}{},\ldots,\mp@subsup{x}{\lambda}{}}}{}f(\mp@subsup{x}{}{\prime})
where p}\in(0,0.5) is the mutation probability
```

The running time of EAs is usually defined as the number of fitness evaluations (i.e., computing $f(\cdot)$ ) until an optimal solution is found for the first time, since the fitness evaluation is often the computational process with the highest cost of the algorithm (He and Yao, 2001, Yu and Zhou, 2008).

### 2.3 Markov Chain Modeling

We will analyze EAs by modeling them as Markov chains in this paper. Here, we give some preliminaries.

EAs often generate solutions only based on their currently maintained solutions, thus, they can be modeled and analyzed as Markov chains, e.g., in He and Yao, 2001; Yu and Zhou, 2008, Yu et al. 2015). A Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ modeling an EA is constructed by taking the EA's population space $\mathcal{X}$ as the chain's state space, i.e. $\xi_{t} \in \mathcal{X}$. Let $\mathcal{X}^{*} \subset \mathcal{X}$ denote the set of all optimal populations, which contain at least one optimal solution. The goal of the EA is to reach $\mathcal{X}^{*}$ from an initial population. Thus, the process of an EA seeking $\mathcal{X}^{*}$ can be analyzed by studying the corresponding Markov chain with the optimal state space $\mathcal{X}^{*}$. Note that we consider the discrete state space (i.e., $\mathcal{X}$ is discrete) in this paper.

A Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}\left(\xi_{t} \in \mathcal{X}\right)$ is a random process, where $\forall t \geq 0, \xi_{t+1}$ depends only on $\xi_{t}$. A Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ is said to be homogeneous, if

$$
\begin{equation*}
\forall t \geq 0, \forall x, y \in \mathcal{X}: P\left(\xi_{t+1}=y \mid \xi_{t}=x\right)=P\left(\xi_{1}=y \mid \xi_{0}=x\right) \tag{1}
\end{equation*}
$$

In this paper, we always denote $\mathcal{X}$ and $\mathcal{X}^{*}$ as the state space and the optimal state space of a Markov chain, respectively.

Given a Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ and $\xi_{\hat{t}}=x$, we define the first hitting time (FHT) of the chain as a random variable $\tau$ such that $\tau=\min \left\{t \mid \xi_{\hat{t}+t} \in \mathcal{X}^{*}, t \geq 0\right\}$. That is, $\tau$ is
the number of steps needed to reach the optimal state space for the first time starting from $\xi_{\hat{t}}=x$. The mathematical expectation of $\tau, \mathbb{E} \llbracket \tau \mid \xi_{\hat{t}}=x \rrbracket=\sum_{i=0}^{+\infty} i P(\tau=i)$, is called the expected first hitting time (EFHT) of this chain starting from $\xi_{\hat{t}}=x$. If $\xi_{0}$ is drawn from a distribution $\pi_{0}, \mathbb{E} \llbracket \tau\left|\xi_{0} \sim \pi_{0} \rrbracket=\sum_{x \in \mathcal{X}} \pi_{0}(x) \mathbb{E} \llbracket \tau\right| \xi_{0}=x \rrbracket$ is called the expected first hitting time of the Markov chain over the initial distribution $\pi_{0}$.

For the corresponding EA, the running time is the number of calls to the fitness function until meeting an optimal solution for the first time. Thus, the expected running time starting from $\xi_{0}$ and that starting from $\xi_{0} \sim \pi_{0}$ are respectively equal to

$$
\begin{equation*}
N_{1}+N_{2} \cdot \mathbb{E} \llbracket \tau \mid \xi_{0} \rrbracket \quad \text { and } \quad N_{1}+N_{2} \cdot \mathbb{E} \llbracket \tau \mid \xi_{0} \sim \pi_{0} \rrbracket \text {, } \tag{2}
\end{equation*}
$$

where $N_{1}$ and $N_{2}$ are the number of fitness evaluations for the initial population and each iteration, respectively. For example, for the ( $1+1$ )-EA, $N_{1}=1$ and $N_{2}=1$; for the $(1+\lambda)-\mathrm{EA}, N_{1}=1$ and $N_{2}=\lambda$. Note that, when involving the expected running time of an EA on a problem in this paper, if the initial population is not specified, it is the expected running time starting from a uniform initial distribution $\pi_{u}$, i.e., $N_{1}+N_{2} \cdot \mathbb{E} \llbracket \tau\left|\xi_{0} \sim \pi_{u} \rrbracket=N_{1}+N_{2} \cdot \sum_{x \in \mathcal{X}} \frac{1}{|\mathcal{X}|} \mathbb{E} \llbracket \tau\right| \xi_{0}=x \rrbracket$.

The following two lemmas on the EFHT of Markov chains (Freǐdlin, 1996) will be used in the paper.

## Lemma 1. Given a Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$, we have

$$
\begin{aligned}
& \forall x \in \mathcal{X}^{*}: \mathbb{E} \llbracket \tau \mid \xi_{t}=x \rrbracket=0 \\
& \forall x \notin \mathcal{X}^{*}: \mathbb{E} \llbracket \tau\left|\xi_{t}=x \rrbracket=1+\sum_{y \in \mathcal{X}} P\left(\xi_{t+1}=y \mid \xi_{t}=x\right) \mathbb{E} \llbracket \tau\right| \xi_{t+1}=y \rrbracket
\end{aligned}
$$

Lemma 2. Given a homogeneous Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$, it holds

$$
\forall t_{1}, t_{2} \geq 0, x \in \mathcal{X}: \mathbb{E} \llbracket \tau\left|\xi_{t_{1}}=x \rrbracket=\mathbb{E} \llbracket \tau\right| \xi_{t_{2}}=x \rrbracket
$$

Drift analysis is a commonly used tool for analyzing the EFHT of Markov chains. It was first introduced to the running time analysis of EAs by He and Yao (2001, 2004). Later, it has become a popular tool in this field, and advanced variants have been proposed (e.g., Doerr et al. 2012b; Doerr and Goldberg, 2013)). In this paper, we will use the additive version (i.e., Lemma 3). To use it, a function $V(x)(x \in \mathcal{X})$ has to be constructed to measure the distance of a state $x$ to the optimal state space $\mathcal{X}^{*}$. The distance function $V(x)$ satisfies that $V\left(x \in \mathcal{X}^{*}\right)=0$ and $V\left(x \notin \mathcal{X}^{*}\right)>0$. Then, we need to investigate the progress on the distance to $\mathcal{X}^{*}$ in each step, i.e., $\mathbb{E} \llbracket V\left(\xi_{t}\right)-$ $V\left(\xi_{t+1}\right) \mid \xi_{t} \rrbracket$. An upper (lower) bound of the EFHT can be derived through dividing the initial distance by a lower (upper) bound of the progress.
Lemma 3 (Additive Drift Analysis (He and Yao, 2001, 2004). Given a Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ and a distance function $V(x)$, if it satisfies that for any $t \geq 0$ and any $\xi_{t}$ with $V\left(\xi_{t}\right)>0$,

$$
0<c_{l} \leq \mathbb{E} \llbracket V\left(\xi_{t}\right)-V\left(\xi_{t+1}\right) \mid \xi_{t} \rrbracket \leq c_{u}
$$

then the EFHT of this chain satisfies that

$$
V\left(\xi_{0}\right) / c_{u} \leq \mathbb{E} \llbracket \tau \mid \xi_{0} \rrbracket \leq V\left(\xi_{0}\right) / c_{l}
$$

where $c_{l}, c_{u}$ do not depend on $\xi_{t}$ and $t$.
The simplified drift theorem (Oliveto and Witt 2011, 2012) as presented in Lemma 4 was proposed to prove exponential lower bounds on the FHT of Markov
chains, where $X_{t}$ is usually represented by a mapping of $\xi_{t}$. It requires two conditions: a constant negative drift and exponentially decaying probabilities of jumping towards or away from the goal state. To relax the requirement of a constant negative drift, advanced variants have been proposed, e.g., the simplified drift theorem with self-loops (Rowe and Sudholt, 2014) and the simplified drift theorem with scaling (Oliveto and Witt 2014, 2015). In this paper, we will use the original version (i.e., Lemma4).
Lemma 4 (Simplified Drift Theorem (Oliveto and Witt, 2011, 2012)). Let $X_{t}, t \geq 0$, be real-valued random variables describing a stochastic process over some state space. Suppose there exists an interval $[a, b] \subseteq \mathbb{R}$, two constants $\delta, \epsilon>0$ and, possibly depending on $l:=b-a$, a function $r(l)$ satisfying $1 \leq r(l)=o(l / \log (l))$ such that for all $t \geq 0$ the following two conditions hold:

$$
\begin{aligned}
& \text { 1. } \quad \mathbb{E} \llbracket X_{t}-X_{t+1} \mid a<X_{t}<b \rrbracket \leq-\epsilon, \\
& \text { 2. } \quad P\left(\left|X_{t+1}-X_{t}\right| \geq j \mid X_{t}>a\right) \leq \frac{r(l)}{(1+\delta)^{j}} \text { for } j \in \mathbb{N}_{0} .
\end{aligned}
$$

Then there is a constant $c^{*}>0$ such that for $T^{*}:=\min \left\{t \geq 0: X_{t} \leq a \mid X_{0} \geq b\right\}$ it holds $P\left(T^{*} \leq 2^{c^{*} l / r(l)}\right)=2^{-\Omega(l / r(l))}$.

### 2.4 Pseudo-Boolean Functions

The pseudo-Boolean function class in Definition3is a large function class which only requires the solution space to be $\{0,1\}^{n}$ and the objective space to be $\mathbb{R}$. Many wellknown NP-hard problems (e.g., the vertex cover problem and the 0-1 knapsack problem) belong to this class. Diverse pseudo-Boolean problems with different structures and difficulties have been used to disclose properties of EAs, e.g., in (Droste et al., 1998, 2002, He and Yao, 2001). We consider only maximization problems in this paper. In the following, let $x_{i}$ denote the $i$-th bit of a solution $x \in\{0,1\}^{n}$.
Definition 3 (Pseudo-Boolean Function). A function in the pseudo-Boolean function class has the form: $f:\{0,1\}^{n} \rightarrow \mathbb{R}$.

The Trap problem in Definition 4 is a special instance in this class, in which the aim is to maximize the number of 0 bits of a solution except for the global optimum $11 \ldots 1$ (briefly denoted as $1^{n}$ ). Its optimal function value is $2 n$, and the function value for any non-optimal solution is not larger than 0 . It has been used in the theoretical studies of EAs, and the expected running time of the (1+1)-EA with mutation probability $\frac{1}{n}$ has been proven to be $\Theta\left(n^{n}\right)$ (Droste et al. 2002). It has also been recognized as the hardest instance in the pseudo-Boolean function class with a unique global optimum for the $(1+1)$-EA Qian et al. 2012), i.e., the expected running time of the $(1+1)$-EA on the Trap problem is the largest among the class.
Definition 4 (Trap Problem). Trap Problem of size $n$ is to solve the problem

$$
\arg \max _{x \in\{0,1\}^{n}}\left(f(x)=3 n \prod_{i=1}^{n} x_{i}-\sum_{i=1}^{n} x_{i}\right)
$$

The Peak problem in Definition 5 has the same fitness for all solutions except for the global optimum $1^{n}$. It has been shown that for solving this problem, the ( $1+1$ )-EA with mutation probability $\frac{1}{n}$ needs $2^{\Omega(n)}$ running time with an overwhelming probability (Oliveto and Witt, 2011).

Definition 5 (Peak Problem). Peak Problem of sizen is to solve the problem

$$
\arg \max _{x \in\{0,1\}^{n}}\left(f(x)=\prod_{i=1}^{n} x_{i}\right) .
$$

The OneMax problem in Definition 6 aims to maximize the number of 1 bits of a solution. Its optimal solution is $1^{n}$ with the function value $n$. The running time of EAs has been well studied on the OneMax problem (He and Yao, 2001) Droste et al. 2002; Sudholt |2013); particularly, the expected running time of the ( $1+1$ )-EA with mutation probability $\frac{1}{n}$ is $\Theta(n \log n)$ (Droste et al. 2002). It has also been recognized as the easiest instance in the pseudo-Boolean function class with a unique global optimum for the (1+1)-EA (Qian et al., 2012).
Definition 6 (OneMax Problem). OneMax Problem of size $n$ is to solve the problem

$$
\arg \max _{x \in\{0,1\}^{n}}\left(f(x)=\sum_{i=1}^{n} x_{i}\right) .
$$

## 3 On the Effect of Noisy Fitness

In this section, we provide two types of problems in which the noise can make the optimization easier for EAs. By "easier", we mean that the EA with noise needs less expected running time than that without noise to find the optimal solution.

We analyze EAs by modeling them as Markov chains. Here, we first give some properties of Markov chains, which will be used in the following analysis. We define a partition of the state space of a homogeneous Markov chain based on the EFHT in Definition 7 , and then define a jumping probability of a chain from one state to one state space in Definition 8. It is easy to see that $\mathcal{X}_{0}$ in Definition 7 is just $\mathcal{X}^{*}$, since $\mathbb{E} \llbracket \tau \mid \xi_{0} \in \mathcal{X}^{*} \rrbracket=0$.
Definition 7 (EFHT-Partition). For a homogeneous Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$, the EFHTPartition is a partition of $\mathcal{X}$ into non-empty subspaces $\left\{\mathcal{X}_{0}, \mathcal{X}_{1}, \ldots, \mathcal{X}_{m}\right\}$ such that

$$
\begin{align*}
& \forall x, y \in \mathcal{X}_{i}, \mathbb{E} \llbracket \tau\left|\xi_{0}=x \rrbracket=\mathbb{E} \llbracket \tau\right| \xi_{0}=y \rrbracket  \tag{1}\\
& \mathbb{E} \llbracket \tau\left|\xi_{0} \in \mathcal{X}_{0} \rrbracket<\mathbb{E} \llbracket \tau\right| \xi_{0} \in \mathcal{X}_{1} \rrbracket<\ldots<\mathbb{E} \llbracket \tau \mid \xi_{0} \in \mathcal{X}_{m} \rrbracket .
\end{align*}
$$

Note that, the EFHT-partition is different from the fitness-partition used in the fitness-level method (Wegener, 2002; Sudholt, 2013) for EAs' running time analysis, since the solutions with the same fitness can have different EFHT, and the EFHT order can be either consistent (e.g., the ( $1+\lambda$ )-EA on the Trap problem as in Lemma 7) or inconsistent (e.g., the ( $1+\lambda$ )-EA on the OneMax problem as in Lemma 10) with the fitness order.
Definition 8. For a Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}, P_{\xi}^{t}\left(x, \mathcal{X}^{\prime}\right)=\sum_{y \in \mathcal{X}^{\prime}} P\left(\xi_{t+1}=y \mid \xi_{t}=x\right)$ is the probability of jumping from state $x$ to state space $\mathcal{X}^{\prime} \subseteq \mathcal{X}$ in one step at time $t$.

Lemma 5 compares the EFHT of two Markov chains. It intuitively means that if one chain always has a larger probability of jumping into good states (i.e., $\mathcal{X}_{j}$ with small $j$ values), it needs less time for reaching the optimal state space.
Lemma 5. Given a Markov chain $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ and a homogeneous Markov chain $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ with the same state space $\mathcal{X}$ and the same optimal space $\mathcal{X}^{*}$, let $\left\{\mathcal{X}_{0}, \mathcal{X}_{1}, \ldots, \mathcal{X}_{m}\right\}$ denote the EFHT-Partition of $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$. If for all $t \geq 0, x \in \mathcal{X}-\mathcal{X}_{0}$, and for all integers $i \in[0, m-1]$,

$$
\begin{gather*}
\sum_{j=0}^{i} P_{\xi}^{t}\left(x, \mathcal{X}_{j}\right) \geq(\leq) \sum_{j=0}^{i} P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right),  \tag{3}\\
\text { then for all } x \in \mathcal{X}, \mathbb{E} \llbracket \tau\left|\xi_{0}=x \rrbracket \leq(\geq) \mathbb{E} \llbracket \tau^{\prime}\right| \xi_{0}^{\prime}=x \rrbracket .
\end{gather*}
$$

To prove Lemma 5, we need the following lemma, which is proved by using the property of majorization and Schur-concavity.

Lemma 6. Let $m(m \geq 1)$ be an integer. If it satisfies that
(1) $0 \leq E_{0}<E_{1}<\ldots<E_{m}$;
(2) $\forall 0 \leq i \leq m, P_{i}, Q_{i} \geq 0, \sum_{i=0}^{m} P_{i}=\sum_{i=0}^{m} Q_{i}=1$;

$$
\begin{equation*}
\forall 0 \leq k \leq m-1, \sum_{i=0}^{k} P_{i} \leq \sum_{i=0}^{k} Q_{i} \tag{3}
\end{equation*}
$$

then it holds that

$$
\sum_{i=0}^{m} P_{i} \cdot E_{i} \geq \sum_{i=0}^{m} Q_{i} \cdot E_{i}
$$

Proof. Let $f\left(x_{0}, \ldots, x_{m}\right)=\sum_{i=0}^{m} E_{i} x_{i}$. Due to the condition (1) that $E_{i}$ is increasing, $f$ is Schur-concave by Theorem A. 3 in Chapter 3 of (Marshall et al., 2011). The conditions (2) and (3) imply that the vector $\left(Q_{0}, \ldots, Q_{m}\right)$ majorizes $\left(P_{0}, \ldots, P_{m}\right)$. Thus, we have $f\left(P_{0}, \ldots, P_{m}\right) \geq f\left(Q_{0}, \ldots, Q_{m}\right)$, which proves the lemma.

## Proof of Lemma5,

We prove one direction of the inequality, and the other can be proved similarly. We use Lemma 3 to derive a bound on $\mathbb{E} \llbracket \tau \mid \xi_{0} \rrbracket$, based on which this lemma holds.

To use Lemma 3 to analyze $\mathbb{E} \llbracket \tau \mid \xi_{0} \rrbracket$, we first construct a distance function $V(x)$ as

$$
\begin{equation*}
\forall x \in \mathcal{X}, V(x)=\mathbb{E} \llbracket \tau^{\prime} \mid \xi_{0}^{\prime}=x \rrbracket, \tag{4}
\end{equation*}
$$

which satisfies that $V\left(x \in \mathcal{X}^{*}\right)=0$ and $V\left(x \notin \mathcal{X}^{*}\right)>0$ by Lemma 1 .
Then, we investigate $\mathbb{E} \llbracket V\left(\xi_{t}\right)-V\left(\xi_{t+1}\right) \mid \xi_{t}=x \rrbracket$ for any $x$ with $V(x)>0$.

$$
\begin{aligned}
& \mathbb{E} \llbracket V\left(\xi_{t}\right)-V\left(\xi_{t+1}\right)\left|\xi_{t}=x \rrbracket=V(x)-\mathbb{E} \llbracket V\left(\xi_{t+1}\right)\right| \xi_{t}=x \rrbracket \\
& =V(x)-\sum_{y \in \mathcal{X}} P\left(\xi_{t+1}=y \mid \xi_{t}=x\right) V(y) \\
& =1+\sum_{y \in \mathcal{X}} P\left(\xi_{1}^{\prime}=y \mid \xi_{0}^{\prime}=x\right) \mathbb{E} \llbracket \tau^{\prime}\left|\xi_{1}^{\prime}=y \rrbracket-\sum_{y \in \mathcal{X}} P\left(\xi_{t+1}=y \mid \xi_{t}=x\right) \mathbb{E} \llbracket \tau^{\prime}\right| \xi_{0}^{\prime}=y \rrbracket
\end{aligned}
$$

(by Eq 4 and Lemma 1 )

$$
=1+\sum_{y \in \mathcal{X}} P\left(\xi_{t+1}^{\prime}=y \mid \xi_{t}^{\prime}=x\right) \mathbb{E} \llbracket \tau^{\prime}\left|\xi_{0}^{\prime}=y \rrbracket-\sum_{y \in \mathcal{X}} P\left(\xi_{t+1}=y \mid \xi_{t}=x\right) \mathbb{E} \llbracket \tau^{\prime}\right| \xi_{0}^{\prime}=y \rrbracket
$$

(by Eq 1 and Lemma 2 since $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ is homogeneous.)

$$
=1+\sum_{j=0}^{m}\left(P_{\xi^{\prime}}^{t}\left(x, \overline{\mathcal{X}_{j}}\right)-P_{\xi}^{t}\left(x, \mathcal{X}_{j}\right)\right) \mathbb{E} \llbracket \tau^{\prime} \mid \xi_{0}^{\prime} \in \mathcal{X}_{j} \rrbracket . \quad(\text { by Definitions } 7 \text { and } 8)
$$

Since $\sum_{j=0}^{m} P_{\xi}^{t}\left(x, \mathcal{X}_{j}\right)=\sum_{j=0}^{m} P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right)=1, \mathbb{E} \llbracket \tau^{\prime} \mid \xi_{0}^{\prime} \in \mathcal{X}_{j} \rrbracket$ increases with $j$ and Eq 3 holds, by Lemma 6, we have

$$
\sum_{j=0}^{m} P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right) \mathbb{E} \llbracket \tau^{\prime}\left|\xi_{0}^{\prime} \in \mathcal{X}_{j} \rrbracket \geq \sum_{j=0}^{m} P_{\xi}^{t}\left(x, \mathcal{X}_{j}\right) \mathbb{E} \llbracket \tau^{\prime}\right| \xi_{0}^{\prime} \in \mathcal{X}_{j} \rrbracket
$$

Thus, we have, for all $t \geq 0$, all $x \notin \mathcal{X}^{*}, \mathbb{E} \llbracket V\left(\xi_{t}\right)-V\left(\xi_{t+1}\right) \mid \xi_{t}=x \rrbracket \geq 1$.
By Lemma 3, we get for all $x \in \mathcal{X}, \mathbb{E} \llbracket \tau\left|\xi_{0}=x \rrbracket \leq V(x)=\mathbb{E} \llbracket \tau^{\prime}\right| \xi_{0}^{\prime}=x \rrbracket$.

### 3.1 On Deceptive Problems

Most practical EAs employ time-invariant operators, thus we can model an EA without noise by a homogeneous Markov chain. While for an EA with noise, since noise
may change over time, we can just model it by a Markov chain. In the following analysis, we always denote them respectively by $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ and $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$, and denote the EFHTPartition of $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ by $\left\{\mathcal{X}_{0}, \mathcal{X}_{1}, \ldots, \mathcal{X}_{m}\right\}$.

An evolutionary process can be characterized by the variation (i.e., producing new solutions) and the selection (i.e., weeding out bad solutions) process. Denote the state spaces before and after variation by $\mathcal{X}$ and $\mathcal{X}_{\text {var }}$ respectively, and then the variation process is a mapping $\mathcal{X} \rightarrow \mathcal{X}_{\text {var }}$ and the selection process is a mapping $\mathcal{X} \times \mathcal{X}_{\text {var }} \rightarrow \mathcal{X}$ (e.g., for the ( $1+\lambda$ )-EA on any pseudo-Boolean problem, $\mathcal{X}=\{0,1\}^{n}$ and $\left.\mathcal{X}_{\text {var }}=\left\{\left\{x_{1}, \ldots, x_{\lambda}\right\} \mid x_{i} \in\{0,1\}^{n}\right\}\right)$. Note that $\mathcal{X}$ is just the state space of the Markov chain. Let $P_{\text {var }}\left(x, x^{\prime}\right)\left(x \in \mathcal{X}, x^{\prime} \in \mathcal{X}_{\text {var }}\right)$ be the state transition probability by the variation process. Let $\mathcal{S}^{*}$ denote the optimal solution set. The considered solution set in the paper (e.g., population) may be a multiset. For two multisets $y \subseteq x$, we mean that $\forall s \in y: s \in x$.
Definition 9 (Deceptive Markov Chain). A homogeneous Markov chain $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ modeling an EA optimizing a problem without noise is deceptive, if for any $x \in \mathcal{X}_{k}(k \geq 1)$,

$$
\begin{align*}
& \forall 1 \leq j \leq k-1: P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right)=0 ;  \tag{5}\\
& \forall k+1 \leq i \leq m: \sum_{j=i}^{m} P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right) \geq \sum_{\substack{x^{\prime} \cap \\
x_{j=i}^{\prime} \mathcal{S}^{*}=\emptyset}} P_{v a r}\left(x, x^{\prime}\right) .
\end{align*}
$$

Theorem 1. For an EA $\mathcal{A}$ optimizing a problem $f$, which can be modeled by a deceptive Markov chain, if

$$
\begin{equation*}
\forall x \notin \mathcal{X}_{0}: P_{\xi}^{t}\left(x, \mathcal{X}_{0}\right)=\sum_{x^{\prime} \cap \mathcal{S}^{*} \neq \emptyset} P_{\text {var }}\left(x, x^{\prime}\right), \tag{6}
\end{equation*}
$$

then noise makes $f$ easier for $\mathcal{A}$.
The theorem intuitively means that if an evolutionary process is deceptive and the optimal solution will always be accepted once generated in the noisy evolutionary process, then noise will be helpful.

## Proof of Theorem 1 .

The two EAs with and without noise are different only on whether the fitness evaluation is disturbed by noise, thus, they must have the same values on $N_{1}$ and $N_{2}$ for their running time Eq.2. Then, comparing their expected running time is equivalent to comparing the EFHT of their corresponding Markov chains.

In one step of the evolutionary process, denote the states before and after variation by $x \in \mathcal{X}$ and $x^{\prime} \in \mathcal{X}_{\text {var }}$ respectively, and denote the state after selection by $y \in \mathcal{X}$. Because the selection process does not produce new solutions, it must satisfy that $y \subseteq x \cup x^{\prime}$. Assume that $x \in \mathcal{X}_{k}(k \geq 1)$. For $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ (i.e., without noise), we have

$$
\begin{equation*}
P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{0}\right) \leq \sum_{x^{\prime} \cap \mathcal{S}^{*} \neq \emptyset} P_{\text {var }}\left(x, x^{\prime}\right) . \tag{7}
\end{equation*}
$$

For $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ (i.e., with noise), the condition Eq $\sqrt{6}$ makes that once an optimal solution is generated, it will be always accepted. Thus, we have

By combining Eq $\sqrt[5]{ }, \mathrm{Eq}, 6, \mathrm{Eq}, 7$ and Eq $\sqrt[8]{ }$, we have

$$
\forall 1 \leq i \leq m: \sum_{j=i}^{m} P_{\xi}^{t}\left(x, \mathcal{X}_{j}\right) \leq \sum_{j=i}^{m} P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right) .
$$

Since $\sum_{j=0}^{m} P_{\xi}^{t}\left(x, \mathcal{X}_{j}\right)=\sum_{j=0}^{m} P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right)=1$, the above inequality is equivalent to

$$
\forall 0 \leq i \leq m-1: \sum_{j=0}^{i} P_{\xi}^{t}\left(x, \mathcal{X}_{j}\right) \geq \sum_{j=0}^{i} P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right)
$$

which implies that the condition Eq 3 of Lemma 5 holds. Thus, by Lemma 5 , we get $\forall x \in \mathcal{X}, \mathbb{E} \llbracket \tau\left|\xi_{0}=x \rrbracket \leq \mathbb{E} \llbracket \tau^{\prime}\right| \xi_{0}^{\prime}=x \rrbracket$, i.e., noise makes $f$ easier for $\mathcal{A}$.

Then, we give a concrete deceptive evolutionary process, i.e., the ( $1+\lambda$ )-EA optimizing the Trap problem. For the Trap problem given in Definition4 it is to maximize the number of 0 bits except for the optimal solution $1^{n}$. It is not hard to see that the EFHT $\mathbb{E} \llbracket \tau^{\prime} \mid \xi_{0}^{\prime}=x \rrbracket$ only depends on $|x|_{0}$ (i.e., the number of 0 bits). We denote $\mathbb{E}_{1}(j)$ as $\mathbb{E} \llbracket \tau^{\prime} \mid \xi_{0}^{\prime}=x \rrbracket$ with $|x|_{0}=j$. The order of $\mathbb{E}_{1}(j)$ is shown in Lemma 7 .
Lemma 7. For any mutation probability $0<p<0.5$, it holds that $\mathbb{E}_{1}(0)<\mathbb{E}_{1}(1)<$ $\mathbb{E}_{1}(2)<\ldots<\mathbb{E}_{1}(n)$.

For proving Lemma 7 , we need the following two lemmas. Lemma 8 (Witt 2013) says that it is more likely that the offspring generated by mutating a parent solution with less 0 bits has a smaller number of 0 bits. Note that we consider $|\cdot|_{0}$ instead of $|\cdot|_{1}$ in their original lemma. It obviously still holds due to the symmetry. We have also restricted $p<0.5$ instead of $p \leq 0.5$, which leads to the strict inequality in the conclusion. Lemma 9 is very similar to Lemma 6, except that the inequalities in the condition (3) and the conclusion hold strictly.
Lemma 8 ((,Witt, 2013)). Let $x, y \in\{0,1\}^{n}$ be two search points satisfying $|x|_{0}<|y|_{0}$. Denote by $x^{\prime}$ and $y^{\prime}$ the random strings obtained by flipping each bit of $x$ and $y$ independently with probability $p$, respectively. If $p<0.5$, then for any $0 \leq j \leq n-1$,

$$
P\left(\left|x^{\prime}\right|_{0} \leq j\right)>P\left(\left|y^{\prime}\right|_{0} \leq j\right)
$$

Lemma 9. Let $m(m \geq 1)$ be an integer. If it satisfies that
(1) $0 \leq E_{0}<E_{1}<\ldots<E_{m}$;
(2) $\forall 0 \leq i \leq m, P_{i}, Q_{i} \geq 0, \sum_{i=0}^{m} P_{i}=\sum_{i=0}^{m} Q_{i}=1$;
(3) $\forall 0 \leq k \leq m-1, \sum_{i=0}^{k} P_{i}<\sum_{i=0}^{k} Q_{i}$,
then it holds that $\sum_{i=0}^{m} P_{i} \cdot E_{i}>\sum_{i=0}^{m} Q_{i} \cdot E_{i}$.
Proof. Let $R_{i}=P_{i}$ for $0 \leq i \leq m-2, R_{m-1}=\sum_{i=0}^{m-1} Q_{i}-\sum_{i=0}^{m-2} P_{i}$ and $R_{m}=Q_{m}$. Then, it is easy to see that the two vectors $\left(R_{0}, \ldots, R_{m}\right)$ and $\left(Q_{0}, \ldots, Q_{m}\right)$ satisfy the conditions (2) and (3) of Lemma 6. Furthermore, the condition (1) of Lemma6 that $E_{i}$ is increasing holds. Thus, by Lemma 6, we have $\sum_{i=0}^{m} R_{i} \cdot E_{i} \geq \sum_{i=0}^{m} Q_{i} \cdot E_{i}$.

Then, we compare $\sum_{i=0}^{m} P_{i} \cdot E_{i}$ with $\sum_{i=0}^{m} R_{i} \cdot E_{i}$.

$$
\begin{aligned}
& \sum_{i=0}^{m} P_{i} \cdot E_{i}-\sum_{i=0}^{m} R_{i} \cdot E_{i}=\left(P_{m-1}-\left(\sum_{i=0}^{m-1} Q_{i}-\sum_{i=0}^{m-2} P_{i}\right)\right) E_{m-1}+\left(P_{m}-Q_{m}\right) E_{m} \\
& =\left(\sum_{i=0}^{m-1} P_{i}-\sum_{i=0}^{m-1} Q_{i}\right)\left(E_{m-1}-E_{m}\right)>0
\end{aligned}
$$

Thus, we have $\sum_{i=0}^{m} P_{i} \cdot E_{i}>\sum_{i=0}^{m} R_{i} \cdot E_{i} \geq \sum_{i=0}^{m} Q_{i} \cdot E_{i}$, i.e., the lemma holds.

## Proof of Lemma 7 ,

First, $\mathbb{E}_{1}(0)<\mathbb{E}_{1}(1)$ trivially holds, because $\mathbb{E}_{1}(0)=0$ and $\mathbb{E}_{1}(1)>0$. Then, we prove $\forall 0<j<n: \mathbb{E}_{1}(j)<\mathbb{E}_{1}(j+1)$ inductively on $j$.
(a) Initialization is to prove $\mathbb{E}_{1}(n-1)<\mathbb{E}_{1}(n)$. For $\mathbb{E}_{1}(n)$, because the next solution can be only $1^{n}$ or $0^{n}$, we have $\mathbb{E}_{1}(n)=1+\left(1-\left(1-p^{n}\right)^{\lambda}\right) \mathbb{E}_{1}(0)+\left(1-p^{n}\right)^{\lambda} \mathbb{E}_{1}(n)$, then, $\mathbb{E}_{1}(n)=1 /\left(1-\left(1-p^{n}\right)^{\lambda}\right)$. For $\mathbb{E}_{1}(n-1)$, because the next solution can be $1^{n}, 0^{n}$ or a solution with $n-1$ number of 0 bits, we have $\mathbb{E}_{1}(n-1)=1+\left(1-\left(1-p^{n-1}(1-\right.\right.$ $\left.p))^{\lambda}\right) \mathbb{E}_{1}(0)+P \cdot \mathbb{E}_{1}(n)+\left(\left(1-p^{n-1}(1-p)\right)^{\lambda}-P\right) \mathbb{E}_{1}(n-1)$, where $P$ denotes the probability that the next solution is $0^{n}$. Then, $\mathbb{E}_{1}(n-1)=\left(1+P \mathbb{E}_{1}(n)\right) /\left(1-\left(1-p^{n-1}(1-p)\right)^{\lambda}+P\right)$. Thus, we have

$$
\frac{\mathbb{E}_{1}(n-1)}{\mathbb{E}_{1}(n)}=\frac{1-\left(1-p^{n}\right)^{\lambda}+P}{1-\left(1-p^{n-1}(1-p)\right)^{\lambda}+P}<1,
$$

where the inequality is by $0<p<0.5$.
(b) Inductive Hypothesis assumes that

$$
\forall K<j \leq n-1(K \geq 1): \mathbb{E}_{1}(j)<\mathbb{E}_{1}(j+1) .
$$

Then, we consider $j=K$. Let $x$ and $y$ be a solution with $K+1$ number of 0 bits and that with $K$ number of 0 bits, respectively. Let $a$ and $b$ denote the number of 0 bits of the offspring solutions $\operatorname{mut}(x)$ and $\operatorname{mut}(y)$, respectively. That is, $a=|\operatorname{mut}(x)|_{0}$ and $b=|\operatorname{mut}(y)|_{0}$. For the $\lambda$ independent mutations on $x$ and $y$, we use $a_{1}, \ldots, a_{\lambda}$ and $b_{1}, \ldots, b_{\lambda}$, respectively. Note that, $a_{1}, \ldots, a_{\lambda}$ are independently and identically distributed (i.i.d.), and $b_{1}, \ldots, b_{\lambda}$ are also i.i.d. Let $p_{j}=P\left(a_{i} \leq j\right)$ and $q_{j}=P\left(b_{i} \leq j\right)$. Then, from Lemma 8 , we have $\forall 0 \leq j \leq n-1: p_{j}<q_{j}$.

For $\mathbb{E}_{1}(K+1)$, let $P_{0}$ and $P_{i}(1 \leq i \leq n)$ be the probability that for the $\lambda$ offspring solutions, the least number of 0 bits is 0 (i.e., $P_{0}=P\left(\min \left\{a_{1}, \ldots, a_{\lambda}\right\}=0\right)$ ), and the largest number of 0 bits is $i$ while the least number of 0 bits is larger than 0 (i.e., $\left.P_{i}=P\left(\max \left\{a_{1}, \ldots, a_{\lambda}\right\}=i \wedge \min \left\{a_{1}, \ldots, a_{\lambda}\right\}>0\right)\right)$, respectively. By considering the mutation and selection behavior of the $(1+\lambda)$-EA on the Trap problem, we have

$$
\mathbb{E}_{1}(K+1)=1+P_{0} \mathbb{E}_{1}(0)+\sum_{i=1}^{K+1} P_{i} \mathbb{E}_{1}(K+1)+\sum_{i=K+2}^{n} P_{i} \mathbb{E}_{1}(i) .
$$

For $\mathbb{E}_{1}(K)$, let $Q_{0}=P\left(\min \left\{b_{1}, \ldots, b_{\lambda}\right\}=0\right)$ and $Q_{i}=P\left(\max \left\{b_{1}, \ldots, b_{\lambda}\right\}=i \wedge\right.$ $\left.\min \left\{b_{1}, \ldots, b_{\lambda}\right\}>0\right)$. Then, we can have

$$
\mathbb{E}_{1}(K)=1+Q_{0} \mathbb{E}_{1}(0)+\sum_{i=1}^{K} Q_{i} \mathbb{E}_{1}(K)+\sum_{i=K+1}^{n} Q_{i} \mathbb{E}_{1}(i) .
$$

For comparing $\mathbb{E}_{1}(K+1)$ with $\mathbb{E}_{1}(K)$, we need to show that

$$
\begin{equation*}
\forall 0 \leq j \leq n-1: \sum_{i=0}^{j} P_{i}<\sum_{i=0}^{j} Q_{i} . \tag{9}
\end{equation*}
$$

For $\sum_{i=0}^{j} P_{i}$, we have

$$
\begin{aligned}
& \sum_{i=0}^{j} P_{i}=P\left(\min \left\{a_{1}, \ldots, a_{\lambda}\right\}=0\right)+P\left(\max \left\{a_{1}, \ldots, a_{\lambda}\right\} \leq j \wedge \min \left\{a_{1}, \ldots, a_{\lambda}\right\}>0\right) \\
& =P\left(a_{1}=0 \vee \ldots \vee a_{\lambda}=0\right)+P\left(0<a_{1} \leq j \wedge \ldots \wedge 0<a_{\lambda} \leq j\right) \\
& =1-\left(1-p_{0}\right)^{\lambda}+\left(p_{j}-p_{0}\right)^{\lambda}<1-\left(1-p_{0}\right)^{\lambda}+\left(q_{j}-p_{0}\right)^{\lambda} . \quad\left(\text { by } p_{j}<q_{j}\right)
\end{aligned}
$$

For $\sum_{i=0}^{j} Q_{i}$, we similarly have $\sum_{i=0}^{j} Q_{i}=1-\left(1-q_{0}\right)^{\lambda}+\left(q_{j}-q_{0}\right)^{\lambda}$. Thus,

$$
\begin{aligned}
& \sum_{i=0}^{j} Q_{i}-\sum_{i=0}^{j} P_{i}>\left(1-p_{0}\right)^{\lambda}-\left(1-q_{0}\right)^{\lambda}+\left(q_{j}-q_{0}\right)^{\lambda}-\left(q_{j}-p_{0}\right)^{\lambda} \\
& =\left(\left(1-q_{0}+q_{0}-p_{0}\right)^{\lambda}-\left(1-q_{0}\right)^{\lambda}\right)-\left(\left(q_{j}-q_{0}+q_{0}-p_{0}\right)^{\lambda}-\left(q_{j}-q_{0}\right)^{\lambda}\right) \\
& =f\left(1-q_{0}\right)-f\left(q_{j}-q_{0}\right)
\end{aligned}
$$

where the last equality is by letting $f(x)=\left(x+q_{0}-p_{0}\right)^{\lambda}-x^{\lambda}$.
Since $q_{0}>p_{0}$, it is easy to verify that $f(x)$ is increasing. Then, we have $f\left(1-q_{0}\right)>$ $f\left(q_{j}-q_{0}\right)$ by $q_{j}<1$. Thus, the Eq 9 holds.
By subtracting $\mathbb{E}_{1}(K)$ from $\mathbb{E}_{1}(K+1)$, we can get

$$
\begin{aligned}
& \mathbb{E}_{1}(K+1)-\mathbb{E}_{1}(K)=\left(P_{0} \mathbb{E}_{1}(0)+\sum_{i=1}^{K+1} P_{i} \mathbb{E}_{1}(K+1)+\sum_{i=K+2}^{n} P_{i} \mathbb{E}_{1}(i)-Q_{0} \mathbb{E}_{1}(0)\right. \\
& \left.-\sum_{i=1}^{K+1} Q_{i} \mathbb{E}_{1}(K+1)-\sum_{i=K+2}^{n} Q_{i} \mathbb{E}_{1}(i)\right)+\sum_{i=1}^{K} Q_{i}\left(\mathbb{E}_{1}(K+1)-\mathbb{E}_{1}(K)\right) \\
& >\sum_{i=1}^{K} Q_{i}\left(\mathbb{E}_{1}(K+1)-\mathbb{E}_{1}(K)\right),
\end{aligned}
$$

where the inequality is by applying Lemma 9 to the formula in $(\cdot)$. The three conditions of Lemma 9 can be easily verified, because $\mathbb{E}_{1}(0)=0<\mathbb{E}_{1}(K+1)<\ldots<\mathbb{E}_{1}(n)$ by inductive hypothesis; $\sum_{i=0}^{n} P_{i}=\sum_{i=0}^{n} Q_{i}=1$; and Eq 9 holds.
Because $\sum_{i=1}^{K} Q_{i}<1$, we have $\mathbb{E}_{1}(K+1)>\mathbb{E}_{1}(K)$.
(c) Conclusion: According to (a) and (b), the lemma holds.

Theorem 2. Either additive noise with $\delta_{2}-\delta_{1}<2 n$ or multiplicative noise with $\delta_{2}>$ $\delta_{1}>0$ makes the Trap problem easier for the $(1+\lambda)-E A$ with mutation probability less than 0.5.

Proof. First, we are to show that the $(1+\lambda)$-EA optimizing the Trap problem can be modeled by a deceptive Markov chain. By Lemma 7, the EFHT-Partition of $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ is $\mathcal{X}_{i}=\left\{\left.x \in\{0,1\}^{n}| | x\right|_{0}=i\right\}(0 \leq i \leq n)$ and $m$ in Definition 7 equals to $n$ here.

For any $x \in \mathcal{X}_{k}(k \geq 1)$, we denote $P(0)$ and $P(j)(1 \leq j \leq n)$ as the probability that for the $\lambda$ offspring solutions $x_{1}, \ldots, x_{\lambda}$ generated by bit-wise mutation on $x, \min \left\{\left|x_{1}\right|_{0}, \ldots,\left|x_{\lambda}\right|_{0}\right\}=0$ (i.e., the least number of 0 bits is 0 ), and $\min \left\{\left|x_{1}\right|_{0}, \ldots,\left|x_{\lambda}\right|_{0}\right\}>0 \wedge \max \left\{\left|x_{1}\right|_{0}, \ldots,\left|x_{\lambda}\right|_{0}\right\}=j$ (i.e., the largest number of 0 bits is $j$ while the least number of 0 bits is larger than 0 ), respectively. For $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$, because only the optimal solution or the solution with the largest number of 0 bit among the parent solution and $\lambda$ offspring solutions will be accepted, we have

$$
\forall 1 \leq j \leq k-1: P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right)=0 ; \quad \forall k+1 \leq j \leq n: P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right)=P(j)
$$

This implies that Eq 5 holds.
Then, we are to show that the condition of Theorem 1 (i.e., Eq 6 holds. For $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ with additive noise, since $\delta_{2}-\delta_{1}<2 n$, we have

$$
f^{N}\left(1^{n}\right) \geq f\left(1^{n}\right)+\delta_{1}>2 n+\delta_{2}-2 n=\delta_{2} ; \quad \forall y \neq 1^{n}, f^{N}(y) \leq f(y)+\delta_{2} \leq \delta_{2}
$$

For multiplicative noise, since $\delta_{2}>\delta_{1}>0$, then $f^{N}\left(1^{n}\right)>0$ and $\forall y \neq 1^{n}, f^{N}(y) \leq 0$. Thus, for these two noises, we have $\forall y \neq 1^{n}, f^{N}\left(1^{n}\right)>f^{N}(y)$, which implies that if the optimal solution $1^{n}$ is generated, it will always be accepted. Thus, we have, note that $\mathcal{X}_{0}=\left\{1^{n}\right\}, P_{\xi}^{t}\left(x, \mathcal{X}_{0}\right)=P(0)$. This implies that Eq, 6 holds.

Thus, by Theorem 1 , we get that the Trap problem becomes easier for the $(1+\lambda)-$ EA under these two types of noise.

### 3.2 On Flat Problems

Besides deceptive problems, we show that noise can also make flat problems easier for EAs. We take the Peak problem given in Definition5 as the representative problem, which has the same fitness for all solutions except for the optimal solution $1^{n}$. When using EAs to solve it, it provides no information for the search direction, thus it is hard for EAs. We analyze the ( $1+1$ )-EA* optimizing the Peak problem. The ( $1+1$ )-EA* is the same as the ( $1+1$ )-EA except that it employs the strict selection strategy. That is, the step 4 of Algorithm 1 changes to be "if $f\left(x^{\prime}\right)>f(x)$ ". The expected running time of the $(1+1)$-EA ${ }^{*}$ with mutation probability $\frac{1}{n}$ on the Peak problem has been proven to be lower bounded by $e^{n \ln (n / 2)}$ Droste et al. 2002).
Theorem 3. One-bit noise with $p_{n} \in(0,1)$ being a constant makes the Peak problem easier for the $(1+1)-E A^{*}$ with mutation probability $\frac{1}{n}$, when starting from an initial solution $x$ with $|x|_{0}>\frac{1+p_{n}}{p_{n}\left(1-\frac{1}{n}\right)}$.
Proof. Let $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ and $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ model the ( $1+1$ )-EA* with one-bit noise and without noise for maximizing the Peak problem, respectively. It is not hard to see that both the EFHT $\mathbb{E} \llbracket \tau \mid \xi_{0}=x \rrbracket$ and $\mathbb{E} \llbracket \tau^{\prime} \mid \xi_{0}^{\prime}=x \rrbracket$ only depend on $|x|_{0}$. We denote $\mathbb{E}(i)$ and $\mathbb{E}^{\prime}(i)$ as $\mathbb{E} \llbracket \tau \mid \xi_{0}=x \rrbracket$ and $\mathbb{E} \llbracket \tau^{\prime} \mid \xi_{0}^{\prime}=x \rrbracket$ with $|x|_{0}=i$, respectively.

For $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ (i.e., without noise) starting from a solution $x$ with $|x|_{0}=i>0$, in one step, any non-optimal offspring solution has the same fitness as the parent and then will be rejected due to the strict selection strategy; only the optimal solution can be accepted, which happens with probability $\frac{1}{n^{2}}\left(1-\frac{1}{n}\right)^{n-i}$. Thus, we have

$$
\mathbb{E}^{\prime}(i)=1+\frac{1}{n^{i}}\left(1-\frac{1}{n}\right)^{n-i} \mathbb{E}^{\prime}(0)+\left(1-\frac{1}{n^{i}}\left(1-\frac{1}{n}\right)^{n-i}\right) \mathbb{E}^{\prime}(i),
$$

which leads to $\mathbb{E}^{\prime}(i)=n^{i}\left(\frac{n}{n-1}\right)^{n-i}$.
For $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ (i.e., with one-bit noise), we assume using re-evaluation, which reevaluates $f(x)$ and evaluates $f\left(x^{\prime}\right)$ in each iteration of Algorithm 1. When starting from $x$ with $|x|_{0}=1$, if the generated offspring $x^{\prime}$ is the optimal solution $1^{n}$, it will be accepted with probability $\left(1-p_{n}\right)\left(1-p_{n} \frac{1}{n}\right)$ because only no bit flip for noise on $x^{\prime}$ and no 0 bit flip for noise on $x$ can make $f^{N}\left(x^{\prime}\right)>f^{N}(x)$; otherwise, $x$ will keep $|x|_{0}=1$, because $f^{N}\left(x^{\prime}\right) \leq f^{N}(x)$ for any $x^{\prime}$ with $\left|x^{\prime}\right|_{0} \geq 2$. Thus, we have

$$
\mathbb{E}(1)=1+\frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1}\left(1-p_{n}\right)\left(1-\frac{p_{n}}{n}\right) \mathbb{E}(0)+\left(1-\frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1}\left(1-p_{n}\right)\left(1-\frac{p_{n}}{n}\right)\right) \mathbb{E}(1)
$$

which leads to $\mathbb{E}(1)=n\left(\frac{n}{n-1}\right)^{n-1} \frac{1}{\left(1-p_{n}\right)\left(1-\frac{p_{n}}{n}\right)}$.
When starting from $x$ with $|x|_{0}=i \geq 2$, if the offspring $x^{\prime}$ is $1^{n}$, it will be accepted with probability $\left(1-p_{n}\right)$ because only no bit flip for noise on $x^{\prime}$ can make $f^{N}\left(x^{\prime}\right)>f^{N}(x)$; if $\left|x^{\prime}\right|_{0}=1$, it will be accepted with probability $p_{n} \frac{1}{n}$ because only flipping the unique 0 bit for noise on $x^{\prime}$ can make $f^{N}\left(x^{\prime}\right)>f^{N}(x)$; otherwise, $x$ keeps $|x|_{0}=i$, because $f^{N}\left(x^{\prime}\right)=f^{N}(x)$ for any $x^{\prime}$ with $\left|x^{\prime}\right|_{0} \geq 2$. Let $m u t_{i \rightarrow 1}$ be the probability of mutating $|x|_{0}=i$ to $\left|x^{\prime}\right|_{0}=1$ by bit-wise mutation with $p=\frac{1}{n}$. Then, we have, for $i \geq 2$,

$$
\begin{aligned}
\mathbb{E}(i)= & 1+\frac{(n-1)^{n-i}}{n^{n}}\left(1-p_{n}\right) \mathbb{E}(0)+\text { mut }_{i \rightarrow 1} \cdot \frac{p_{n}}{n} \mathbb{E}(1) \\
& +\left(1-\frac{(n-1)^{n-i}}{n^{n}}\left(1-p_{n}\right)-m u t_{i \rightarrow 1} \frac{p_{n}}{n}\right) \mathbb{E}(i),
\end{aligned}
$$

which leads to $\mathbb{E}(i)=\frac{1+m u t_{i \rightarrow 1} \frac{p_{n}}{n} \mathbb{E}(1)}{\frac{1}{n^{i}}\left(1-\frac{1}{n}\right)^{n-i}\left(1-p_{n}\right)+m u t_{i \rightarrow 1} \frac{p_{n}}{n}}$.
From Eq 2, we know that the expected running time for without noise and with one-bit noise is $1+\mathbb{E}^{\prime}(i)$ and $1+2 \mathbb{E}(i)$, respectively. To prove that one-bit noise can be helpful, we need to show that there exists $i \geq 1$ such that $2 \mathbb{E}(i)<\mathbb{E}^{\prime}(i)$. Obviously, $i=1$ is impossible because $\mathbb{E}(1)>\mathbb{E}^{\prime}(1)$. Then, for larger $i$ with $i>\frac{1+p_{n}}{p_{n}\left(1-\frac{1}{n}\right)}$,

$$
\begin{aligned}
& \left(2 \mathbb{E}(i)-\mathbb{E}^{\prime}(i)\right) \cdot\left(\frac{1}{n^{i}}\left(1-\frac{1}{n}\right)^{n-i}\left(1-p_{n}\right)+m u t_{i \rightarrow 1} \frac{p_{n}}{n}\right) \\
& =1+p_{n}-m u t_{i \rightarrow 1} \frac{p_{n}}{n}\left(n^{i}\left(\frac{n}{n-1}\right)^{n-i}-2 \mathbb{E}(1)\right) \\
& \leq 1+p_{n}-\frac{i}{n^{i-1}}\left(1-\frac{1}{n}\right)^{n-i+1} \frac{p_{n}}{n} n^{i}\left(\frac{n}{n-1}\right)^{n-i}=1+p_{n}-i p_{n}\left(1-\frac{1}{n}\right)<0
\end{aligned}
$$

where the 1 st inequality is because mut $_{i \rightarrow 1} \geq \frac{i}{n^{i-1}}\left(1-\frac{1}{n}\right)^{n-i+1}$ and $\mathbb{E}(1) \ll n^{i}\left(\frac{n}{n-1}\right)^{n-i}$ for large enough $n$ and $p_{n}$ being constant, and the last inequality is by $i>\frac{1+p_{n}}{p_{n}\left(1-\frac{1}{n}\right)}$. This is equivalent to $2 \mathbb{E}(i)-\mathbb{E}^{\prime}(i)<0$, which implies that noise is helpful when starting from an initial solution $x$ with $|x|_{0}>\frac{1+p_{n}}{p_{n}\left(1-\frac{1}{n}\right)}$.

This theorem implies that the Peak problem becomes easier under noise when starting from an initial solution $x$ with a large number of 0 bits. From the analysis, we can see that the reason of requiring a large $|x|_{0}$ is to make $m u t_{i \rightarrow 1}$ much larger than $m u t_{i \rightarrow 0}$, which makes that the negative effect of rejecting the optimal solution by noise can be compensated by the positive effect of accepting the solution $x$ with $|x|_{0}=1$.

For the $(1+1)$-EA solving the Peak problem, any offspring solution will be accepted because its fitness is always not less than the fitness of the parent solution; thus the solution $x$ in the evolutionary process almost performs a random walk over $\{0,1\}^{n}$. In this case, we can intuitively find a similar effect of one-bit noise as that found in the $(1+1)-E A^{*}$ solving the Peak problem. Here, we assume that the singleevaluation strategy is used. Under one-bit noise, for any non-optimal parent solution $x$, if $|x|_{0} \geq 2$, then $f^{N}(x)=0$ and any offspring will be accepted; if $|x|_{0}=1$ and $f^{N}(x)=0$, then any offspring will be accepted; if $|x|_{0}=1$ and $f^{N}(x)=1$, any offspring $x^{\prime}$ with $\left|x^{\prime}\right|_{0} \geq 2$ will be rejected because $f^{N}\left(x^{\prime}\right)=0<f^{N}(x)$, and the optimal solution with $\left|x^{\prime}\right|_{0}=0$ will be rejected with probability $p_{n}$. Compared with the transition behavior without noise, noise only has an effect when $|x|_{0}=1$ and $f^{N}(x)=1$ : the negative effect of rejecting the optimal solution, which has the probability $\frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1} p_{n}$, and the positive effect of rejecting $\left|x^{\prime}\right|_{0} \geq 2$, which has the probability at least $\frac{n-1}{n}\left(1-\frac{1}{n}\right)^{n-1}$. Obviously, the negative effect can be compensated by the positive effect, which implies that one-bit noise is helpful. Thus, we have the following conjecture. The rigorous analysis is not easy, and we leave it as a future work. We will instead verify it in the experiment section.
Conjecture 1. One-bit noise makes the Peak problem easier for the (1+1)-EA with mutation probability $\frac{1}{n}$.

## 4 On the Effect of Noise Handling Strategies

In the previous section, we have studied the effect of the noise, and found that noise can also make optimization easier for EAs, when the problem presents some decep-
tiveness and flatness. Meanwhile, on some other problems noisy fitness evaluation can make an optimization harder for EAs. For example, Droste (2004) proved that the running time of the $(1+1)$-EA on the OneMax problem can increase from polynomial to exponential due to the presence of noise. Thus, in this section, we will investigate how well different noise handling strategies can perform when the noise is indeed harmful.

### 4.1 A Noise-Harmful Case

We consider the case that the $(1+\lambda)$-EA is used for optimizing the OneMax problem. Let $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$ and $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ model the ( $1+\lambda$ )-EA with and without noise for maximizing OneMax, respectively. It is not hard to see that the EFHT $\mathbb{E} \llbracket \tau^{\prime} \mid \xi_{0}^{\prime}=x \rrbracket$ only depends on $|x|_{0}$. We denote $\mathbb{E}_{2}(j)$ as $\mathbb{E} \llbracket \tau^{\prime} \mid \xi_{0}^{\prime}=x \rrbracket$ with $|x|_{0}=j$. The order of $\mathbb{E}_{2}(j)$ is shown in Lemma 10
Lemma 10. For any mutation probability $0<p<0.5$, it holds that $\mathbb{E}_{2}(0)<\mathbb{E}_{2}(1)<$ $\mathbb{E}_{2}(2)<\ldots<\mathbb{E}_{2}(n)$.
Proof. We prove $\forall 0 \leq j<n: \mathbb{E}_{2}(j)<\mathbb{E}_{2}(j+1)$ inductively on $j$.
(a) Initialization is to prove $\mathbb{E}_{2}(0)<\mathbb{E}_{2}(1)$, which holds since $\mathbb{E}_{2}(1)>0=\mathbb{E}_{2}(0)$.
(b) Inductive Hypothesis assumes that

$$
\forall 0 \leq j<K(K \leq n-1): \mathbb{E}_{2}(j)<\mathbb{E}_{2}(j+1) .
$$

Then, we consider $j=K$. We use the similar analysis method as that in the proof of Lemma 7 to compare $\mathbb{E}_{2}(K+1)$ with $\mathbb{E}_{2}(K)$.

For $\mathbb{E}_{2}(K+1)$, let $P_{i}(0 \leq i \leq n)$ be the probability that the least number of 0 bits for the $\lambda$ offspring solutions is $i$ (i.e., $P_{i}=P\left(\min \left\{a_{1}, \ldots, a_{\lambda}\right\}=i\right)$ ). By considering the mutation and selection behavior of the $(1+\lambda)$-EA on the OneMax problem, we have

$$
\mathbb{E}_{2}(K+1)=\sum_{i=0}^{K} P_{i} \mathbb{E}_{2}(i)+\sum_{i=K+1}^{n} P_{i} \mathbb{E}_{2}(K+1) .
$$

For $\mathbb{E}_{2}(K)$, let $Q_{i}=P\left(\min \left\{b_{1}, \ldots, b_{\lambda}\right\}=i\right)$. We have

$$
\mathbb{E}_{2}(K)=\sum_{i=0}^{K-1} Q_{i} \mathbb{E}_{2}(i)+\sum_{i=K}^{n} Q_{i} \mathbb{E}_{2}(K) .
$$

By subtracting $\mathbb{E}_{2}(K)$ from $\mathbb{E}_{2}(K+1)$, we can get

$$
\begin{aligned}
& \mathbb{E}_{2}(K+1)-\mathbb{E}_{2}(K)=\sum_{i=K+1}^{n} P_{i}\left(\mathbb{E}_{2}(K+1)-\mathbb{E}_{2}(K)\right)+ \\
& \left(\sum_{i=0}^{K-1} P_{i} \mathbb{E}_{2}(i)+\sum_{i=K}^{n} P_{i} \mathbb{E}_{2}(K)-\sum_{i=0}^{K-1} Q_{i} \mathbb{E}_{2}(i)-\sum_{i=K}^{n} Q_{i} \mathbb{E}_{2}(K)\right) \\
& >\sum_{i=K+1}^{n} P_{i}\left(\mathbb{E}_{2}(K+1)-\mathbb{E}_{2}(K)\right),
\end{aligned}
$$

where the inequality is by applying Lemma 9 to the formula in $(\cdot)$. The three conditions of Lemma 9 can be easily verified, because $\mathbb{E}_{2}(0)<\mathbb{E}_{2}(1)<\ldots<\mathbb{E}_{2}(K)$ by inductive hypothesis; $\sum_{i=0}^{n} P_{i}=\sum_{i=0}^{n} Q_{i}=1$; and the following inequality holds.

$$
\begin{aligned}
& \sum_{i=0}^{j} Q_{i}-\sum_{i=0}^{j} P_{i}=P\left(\min \left\{b_{1}, \ldots, b_{\lambda}\right\} \leq j\right)-P\left(\min \left\{a_{1}, \ldots, a_{\lambda}\right\} \leq j\right) \\
& =P\left(b_{1} \leq j \vee \ldots \vee b_{\lambda} \leq j\right)-P\left(a_{1} \leq j \vee \ldots \vee a_{\lambda} \leq j\right)
\end{aligned}
$$

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$$
=1-\left(1-q_{j}\right)^{\lambda}-\left(1-\left(1-p_{j}\right)^{\lambda}\right)>0 . \quad\left(\text { by } p_{j}<q_{j}\right)
$$

Because $\sum_{i=K+1}^{n} P_{i}<1$, we have $\mathbb{E}_{2}(K+1)>\mathbb{E}_{2}(K)$.
(c) Conclusion: According to (a) and (b), the lemma holds.

Theorem 4. Any noise makes the OneMax problem harder for the (1+ 1 )-EA with mutation probability less than 0.5.

Proof. We use Lemma 5 to prove it. By Lemma 10 , the EFHT-Partition of $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$ is $\mathcal{X}_{i}=\left\{\left.x \in\{0,1\}^{n}| | x\right|_{0}=i\right\}(0 \leq i \leq n)$ 。

For any non-optimal solution $x \in \mathcal{X}_{k}(k>0)$, we denote $P(j)(0 \leq j \leq n)$ as the probability that the least number of 0 bits for the $\lambda$ offspring solutions generated by bit-wise mutation on $x$ is $j$. For $\left\{\xi_{t}^{\prime}\right\}_{t=0}^{+\infty}$, because the solution with the least number of 0 bits among the parent solution and $\lambda$ offspring solutions will be accepted, we have

$$
\forall 0 \leq j \leq k-1: P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right)=P(j) ; \quad \quad P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{k}\right)=\sum_{j=k}^{n} P(j)
$$

For $\left\{\xi_{t}\right\}_{t=0}^{+\infty}$, due to the fitness evaluation disturbed by noise, the solution with the least number of 0 bits among the parent and $\lambda$ offspring solutions may be rejected. Thus, we have

$$
0 \leq i \leq k-1: \sum_{j=0}^{i} P_{\xi}^{t}\left(x, \mathcal{X}_{j}\right) \leq \sum_{j=0}^{i} P(j)
$$

Then, we get

$$
\forall 0 \leq i \leq n-1: \sum_{j=0}^{i} P_{\xi}^{t}\left(x, \mathcal{X}_{j}\right) \leq \sum_{j=0}^{i} P_{\xi^{\prime}}^{t}\left(x, \mathcal{X}_{j}\right)
$$

which implies that the condition Eq 3 of Lemma 5 holds. Thus, we can get $\forall x \in \mathcal{X}$, $\mathbb{E} \llbracket \tau\left|\xi_{0}=x \rrbracket \geq \mathbb{E} \llbracket \tau^{\prime}\right| \xi_{0}^{\prime}=x \rrbracket$, i.e., noise makes the OneMax problem harder for the ( $1+\lambda$ )-EA.

In the following subsections, we will analyze the effect of different noise handling strategies for the $(1+1)$-EA (a specific case of the $(1+\lambda)$-EA) optimizing the OneMax problem to investigate their usefulness.

### 4.2 On Re-evaluation and Threshold Selection Strategies

### 4.2.1 Re-evaluation

There are naturally two fitness evaluation options for EAs Arnold and Beyer, 2002, Jin and Branke, 2005, Goh and Tan, 2007):

- single-evaluation: we evaluate a solution once, and use the evaluated fitness for this solution in the future.
- re-evaluation: every time we access the fitness of a solution by evaluation.

For example, for the (1+1)-EA in Algorithm 1 , if using re-evaluation, both $f\left(x^{\prime}\right)$ and $f(x)$ will be calculated and recalculated in each iteration; if using single-evaluation, only $f\left(x^{\prime}\right)$ will be calculated and the previous obtained fitness $f(x)$ will be reused. Note that the analysis in the previous section without explicitly indicating the employed evaluation strategy assumes single-evaluation.

In Sudholt and Thyssen, 2012), for an ACO with single-evaluation solving stochastic shortest path problems, an example graph was constructed to show that
exponential running time is required for approximating real shortest paths. The difficulty is because once a path is luckily evaluated to have a relatively small length due to noise, it will always be preferred and make the ACO get stuck in an inferior solution. By using re-evaluation instead of single-evaluation when evaluating the best-so-far path, the ACO can easily solve the example graph (Doerr et al., 2012a). Re-evaluation has also been employed for EAs solving noisy multi-objective optimization problems, e.g., in (Buche et al., 2002 Park and Ryu, 2011, Fieldsend and Everson, 2015).

Intuitively, re-evaluation can smooth noise and thus could be better for noisy optimizations, but it also increases the fitness evaluation cost and thus increases the running time. Its usefulness was not yet clear.

In this subsection, we compare these two options for the ( $1+1$ )-EA solving the OneMax problem under one-bit noise to show whether re-evaluation is useful. Note that for one-bit noise, $p_{n}$ controls the noise strength, that is, noise becomes stronger as $p_{n}$ gets larger, and it is also the parameter of the PNT. In the following analysis, let $\operatorname{poly}(n)$ indicate any polynomial of $n$.
Theorem 5. For the (1+1)-EA with mutation probability $\frac{1}{n}$ solving the OneMax problem under one-bit noise, if using single-evaluation, the PNT is $[0,1-1 / \Theta($ poly $(n))]$.

The theorem is straightforwardly derived from the following two lemmas.
(1) Lemma 11 tells us the expected running time upper bound $O\left(n^{2}+n /\left(1-p_{n}\right)\right)$, which implies that the expected running time is polynomial if $\frac{1}{1-p_{n}} \in O(\operatorname{poly}(n))$, i.e., $p_{n} \in 1-\frac{1}{O(\text { poly }(n))}$.
(2) Lemma 12 tells us the lower bound $\Omega\left(n \log n+p_{n} /\left(1-p_{n}\right)\right)$, which implies that the running time is super-polynomial if $\frac{1}{1-p_{n}} \in w($ poly $(n))$, i.e., $p_{n} \in 1-\frac{1}{w(p o l y(n))}$.
By combining (1) with (2), we can get that the maximum noise strength allowing polynomial expected running time is $1-\frac{1}{\Theta(p o l y(n))}$, i.e., the PNT is $[0,1-1 / \Theta($ poly $(n))]$.
Lemma 11. For the ( $1+1$ )-EA using single-evaluation with mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, the expected running time is upper bounded by $O\left(n^{2}+n /\left(1-p_{n}\right)\right)$.

Proof. Let $L$ denote the noisy fitness value $f^{N}(x)$ of the current solution $x$. Because the $(1+1)$-EA does not accept a solution with a smaller fitness (i.e., the 4th step of Algorithm (1) and it doesn't re-evaluate the fitness of the current solution $x, L$ ( $0 \leq$ $L \leq n$ ) will never decrease. By applying the fitness-level technique (Wegener, 2002, Sudholt, 2013), we first analyze the expected steps until $L$ increases when starting from $L=i$ (denoted by $\mathbb{E} \llbracket i \rrbracket$ ), and then sum them up to get an upper bound $\sum_{i=0}^{n-1} \mathbb{E} \llbracket i \rrbracket$ for the expected steps until $L$ reaches the maximum value $n$. For $\mathbb{E} \llbracket i \rrbracket$, we analyze the probability $P$ that $L$ increases in two steps when $L=i$, then $\mathbb{E} \llbracket i \rrbracket=2 \cdot \frac{1}{P}$. Note that, one-bit noise can make $L$ be $|x|_{1}-1,|x|_{1}$ or $|x|_{1}+1$, where $|x|_{1}=\sum_{i=1}^{n} x_{i}$ is the number of 1 bits. When analyzing the noisy fitness $f^{N}\left(x^{\prime}\right)$ of the offspring $x^{\prime}$ in each step, we need to first consider bit-wise mutation on $x$ and then one random bit flip for noise.

When $0<L<n-1,|x|_{1}=L-1, L$ or $L+1$.
(1) For $|x|_{1}=L-1, P \geq \frac{n-L+1}{n}\left(1-\frac{1}{n}\right)^{n-1} p_{n} \frac{n-L}{n}+\frac{n-L+1}{n}\left(1-\frac{1}{n}\right)^{n-1}\left(1-p_{n}\right)^{n-L} n(1-$ $\left.\frac{1}{n}\right)^{n-1}\left(1-p_{n}\right)$, since it is sufficient to flip one 0 bit for mutation and one 0 bit for noise in the first step, or flip one 0 bit for mutation and no bit for noise in the first step and flip one 0 bit for mutation and no bit for noise in the second step.
(2) For $|x|_{1}=L, P \geq\left(1-\frac{1}{n}\right)^{n} p_{n} \frac{n-L}{n}+\frac{n-L}{n}\left(1-\frac{1}{n}\right)^{n-1}\left(1-p_{n}\right)$, since it is sufficient to flip no bit for mutation and one 0 bit for noise, or flip one 0 bit for mutation and no bit for noise in the first step.
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(3) For $|x|_{1}=L+1, P \geq\left(1-\frac{1}{n}\right)^{n}\left(1-p_{n}+p_{n} \frac{n-L-1}{n}\right)$, since it is sufficient to flip no bit for mutation and no bit or one 0 bit for noise in the first step.
Thus, for these three cases, we have

$$
\begin{aligned}
P & \geq p_{n}\left(1-\frac{1}{n}\right)^{n-1} \frac{n-L}{n} \frac{n-L-1}{n}+\left(1-\frac{1}{n}\right)^{2(n-1)}\left(1-p_{n}\right)^{2} \frac{n-L}{n} \frac{n-L-1}{n} \\
& \geq \frac{3(n-L)(n-L-1)}{4 e^{2} n^{2}} . \quad\left(\text { by }\left(1-\frac{1}{n}\right)^{n-1} \geq \frac{1}{e} \text { and } 0 \leq p_{n} \leq 1\right)
\end{aligned}
$$

When $L=0,|x|_{1}=0$ or 1 . By considering case (2) and (3), we can get the same lower bound for $P$.

When $L=n-1$ and the optimal solution $1^{n}$ has not been found, $|x|_{1}=n-2$ or $n-1$. By considering case (1) and (2), we can get $P \geq 3 /\left(2 e^{2} n^{2}\right)$.

Based on the above analysis, we get that the expected steps until $L=n$ is at most

$$
\sum_{i=0}^{n-1} \mathbb{E} \llbracket i \rrbracket \leq 2 \cdot\left(\sum_{L=0}^{n-2} \frac{4 e^{2} n^{2}}{3(n-L)(n-L-1)}+\frac{2 e^{2} n^{2}}{3}\right) \in O\left(n^{2}\right)
$$

When $L=n,|x|_{1}=n-1$ or $n .|x|_{1}=n$ means that the optimal solution has been found. Because we are to get an upper bound for the expected running time of finding $1^{n}$, we can pessimistically assume that $|x|_{1}=n-1$. Starting from $|x|_{1}=n-1$ and $L=n$ (i.e., the current solution has $n-1$ one bits and the fitness is $n$ ), it will always keep in such a situation before finding $1^{n}$, and the optimal solution $1^{n}$ can be generated and accepted in one step only through flipping the unique 0 bit for mutation and no bit for noise, which happens with probability $\frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1}\left(1-p_{n}\right) \geq \frac{\left(1-p_{n}\right)}{e n}$. This implies that the expected steps for finding the optimal solution is at most $\frac{e n}{\left(1-p_{n}\right)}$.

Thus, the total expected running time is upper bounded by $O\left(n^{2}+\frac{n}{1-p_{n}}\right)$.
Lemma 12. For the (1+1)-EA using single-evaluation with mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, the expected running time is lower bounded by $\Omega\left(n \log n+p_{n} /\left(1-p_{n}\right)\right)$.
Proof. Assume that the number of 1 bits of the initial solution $x$ is less than $n-1$, i.e., $|x|_{1}<n-1$. Let $T$ denote the running time of finding the optimal solution $1^{n}$ when starting from $x$. Denote $A$ as the event that in the evolutionary process, any solution $x^{\prime}$ with $\left|x^{\prime}\right|_{1}=n-1$ is never found. By the law of total expectation, we have

$$
\mathbb{E} \llbracket T \rrbracket=\mathbb{E} \llbracket T|A \rrbracket \cdot P(A)+\mathbb{E} \llbracket T| \bar{A} \rrbracket \cdot P(\bar{A})
$$

We are first to show that $P(\bar{A}) \geq P(A)$. Let $l: x=x_{1} \rightarrow x_{2} \rightarrow \ldots \rightarrow x_{m-1} \rightarrow x_{m}=$ $1^{n}$ denote an evolutionary path from $x$ to the optimal solution $1^{n}$, which satisfies that $\forall i<m,\left|x_{i}\right|_{1} \leq n-2$. Then, $P(A)$ is the sum of the probabilities of all possible such $l$. For any such $l$, there must exist a corresponding set of paths $S(l)=\left\{x_{1} \rightarrow x_{2} \rightarrow \ldots \rightarrow\right.$ $\left.\left.x_{m-1} \rightarrow y_{m} \rightarrow \ldots \rightarrow 1^{n}| | y_{m}\right|_{1}=n-1\right\}$, in which the first $m-1$ solutions of any path are the same as that of $l$ and the $m$-th solution has $n-1$ number of one bits. Let $q$ denote the probability of the sub-path $x_{1} \rightarrow \cdots \rightarrow x_{m-1}$, and let $\left|x_{m-1}\right|_{1}=n-j \leq$ $n-2$. Then, $P(l)=q \cdot \frac{1}{n^{j}}\left(1-\frac{1}{n}\right)^{n-j}$. The probability of mutating from $x_{m-1}$ to $y_{m}$ is at least $\frac{j}{n^{j-1}}\left(1-\frac{1}{n}\right)^{n-j+1}$, and the acceptance probability of $y_{m}$ is at least $1-p_{n}+p_{n} \frac{1}{n}$, which is reached when $\left|x_{m-1}\right|_{1}=n-2$ and $f^{N}\left(x_{m-1}\right)=n-1$. Thus, we have
$P(S(l)) \geq q \cdot \frac{j}{n^{j-1}}\left(1-\frac{1}{n}\right)^{n-j+1} \cdot\left(1-p_{n}+\frac{p_{n}}{n}\right) \geq q \cdot \frac{j}{n^{j-1}}\left(1-\frac{1}{n}\right)^{n-j+1} \cdot \frac{1}{n} \geq P(l)$.

Moreover, for any two different paths $l_{1}, l_{2}$, it must hold that $S\left(l_{1}\right) \cap S\left(l_{2}\right)=\emptyset$. Thus, $P(\bar{A}) \geq P(A)$. Because $P(\bar{A})+P(A)=1$, we can get $P(\bar{A}) \geq 1 / 2$. Then,

$$
\mathbb{E} \llbracket T \rrbracket \geq \mathbb{E} \llbracket T\left|\bar{A} \rrbracket \cdot P(\bar{A}) \geq \frac{1}{2} \mathbb{E} \llbracket T\right| \bar{A} \rrbracket .
$$

We are then to derive a lower bound on $\mathbb{E} \llbracket T \mid \bar{A} \rrbracket$. We further divide the running time $T$ into two parts: the running time until finding a solution $x^{\prime}$ with $\left|x^{\prime}\right|_{1}=n-1$ for the first time (denoted by $T_{1}$ ), and the remaining running time for finding the optimal solution (denoted by $T_{2}$ ). Thus, we have

$$
\mathbb{E} \llbracket T\left|\bar{A} \rrbracket=\mathbb{E} \llbracket T_{1}\right| \bar{A} \rrbracket+\mathbb{E} \llbracket T_{2} \mid \bar{A} \rrbracket .
$$

For $\mathbb{E}\left[T_{2} \mid \bar{A} \rrbracket\right.$, when finding a solution $x^{\prime}$ with $\left|x^{\prime}\right|_{1}=n-1$ for the first time, we consider the case that the fitness is evaluated as $n$, which happens with probability $p_{n} \frac{1}{n}$. If it happens, due to the single-evaluation strategy, the solution will always have $n-$ 1 number of 1 bits and its fitness will always be $n$. From the upper bound analysis in Lemma 11 we know that the probability of generating and accepting the optimal solution in one step in such a situation is $\frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1}\left(1-p_{n}\right) \leq \frac{\left(1-p_{n}\right)}{n}$. Thus,

$$
\mathbb{E} \llbracket T_{2} \left\lvert\, \bar{A} \rrbracket \geq p_{n} \frac{1}{n} \cdot \frac{n}{1-p_{n}}=\frac{p_{n}}{1-p_{n}}\right.,
$$

which implies that $\mathbb{E} \llbracket T \left\lvert\, \bar{A} \rrbracket \geq \frac{p_{n}}{\left(1-p_{n}\right)}\right.$, and thus $\mathbb{E} \llbracket T \rrbracket \geq \frac{p_{n}}{2\left(1-p_{n}\right)}$.
Because the initial solution is uniformly distributed over $\{0,1\}^{n}$, we have $P\left(|x|_{1}<n-1\right)=1-\frac{n+1}{2^{n}}$. Thus, the expected running time of the whole process is lower bounded by $\left(1-\frac{n+1}{2^{n}}\right) \cdot \frac{p_{n}}{2\left(1-p_{n}\right)}$, i.e., $\Omega\left(\frac{p_{n}}{1-p_{n}}\right)$.

Note that when $1-p_{n} \in \Omega(1)$, the derived lower bound $\Omega\left(\frac{p_{n}}{1-p_{n}}\right)$ would be quite loose. Thus, for filling up this gap, we are to derive another lower bound which does not depend on $p_{n}$. From Lemma 10 in (Droste et al. 2002), we know that the expected running time of the ( $1+1$ )-EA to optimize linear functions with positive weights is $\Omega(n \log n)$. Their proof idea is to analyze the expected running time until all the 0 bits of the initial solution have been flipped at least once, which is obviously a lower bound on the expected running time of finding the optimal solution $1^{n}$. Because noise will not affect this analysis process, we can directly apply their result to our setting, and then get the lower bound $\Omega(n \log n)$.

By combining the derived two lower bounds, we can get that the expected running time of the whole process is lower bounded by $\Omega\left(n \log n+p_{n} /\left(1-p_{n}\right)\right)$.

We then show the PNT using re-evaluation in the following theorem, which can be straightforwardly derived from Lemma 13 (Droste 2004).
Theorem 6. For the ( $1+1$ )-EA with mutation probability $\frac{1}{n}$ solving the OneMax problem under one-bit noise, if using re-evaluation, the PNT is $\left[0, \Theta\left(\frac{\log (n)}{n}\right)\right]$.
Lemma 13 ((Droste, 2004)). For the (1+1)-EA using re-evaluation with mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, the expected running time is polynomial when $p_{n} \in O(\log (n) / n)$, and super-polynomial when $p_{n} \in \omega(\log (n) / n)$.

### 4.2.2 Threshold Selection

During the process of evolutionary optimization, most of the improvements in one generation are small. When using re-evaluation, due to noisy fitness evaluation, a
considerable portion of these improvements are not real, where a worse solution appears to have a "better" fitness and then survives to replace the true better solution which has a "worse" fitness. This may mislead the search direction of EAs, and then slow down the efficiency of EAs or make EAs get trapped in the local optimal solution, as observed in Section 4.2.1. To deal with this problem, a selection strategy for EAs handling noise was proposed (Markon et al. 2001 Bartz-Beielstein, 2005a).

- threshold selection: an offspring solution will be accepted only if its fitness is larger than the parent solution by at least a predefined threshold $\tau \geq 0$.
For example, for the ( $1+1$ )-EA with threshold selection as in Algorithm 3 its 4th step changes to be "if $f\left(x^{\prime}\right) \geq f(x)+\tau^{\prime}$ " rather than "if $f\left(x^{\prime}\right) \geq f(x)$ " in Algorithm 1 . Such a strategy can reduce the risk of accepting a bad solution due to noise. Although the good local performance (i.e., the progress of one step) of EAs with threshold selection has been shown on some problems (Markon et al. 2001 Bartz-Beielstein and Markon. 2002: Bartz-Beielstein 2005b), its usefulness for the global performance (i.e., the running time until finding the optimal solution) of EAs under noise is not yet clear.
Algorithm 3 (( $1+1$ )-EA with threshold selection). Given pseudo-Boolean function $f$ with solution length $n$, and a threshold $\tau \geq 0$, it consists of the following steps:

1. $x:=$ randomly selected from $\{0,1\}^{n}$.
2. Repeat until the termination condition is met
3. $\quad x^{\prime}:=$ flip each bit of $x$ independently with probability $p$.
4. if $f\left(x^{\prime}\right) \geq f(x)+\tau$
5. $\quad x:=x^{\prime}$.
where $p \in(0,0.5)$ is the mutation probability.
In this subsection, we analyze the running time of the $(1+1)$-EA with threshold selection solving OneMax under one-bit noise to see whether threshold selection is useful. Note that the analysis here assumes re-evaluation. This is because using singleevaluation and threshold selection simultaneously will lead to infinite expected running time for any noise strength $p_{n}>0$, as shown in the following theorem.
Theorem 7. For the ( $1+1$ )-EA with mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, if using single-evaluation with threshold selection $\tau>0$, the PNT is $\{0\}$.

Proof. For the noise strength $p_{n}>0$, it is easy to see that in the evolutionary process, there exists some positive probability that a solution $x$ with $|x|_{1}=n-1$ is found and its fitness is evaluated as $n$. Once it happens, it will always keep in such a situation, because the fitness of the parent solution will never be re-evaluated; the fitness of the offspring solution is at most $n$; and then any offspring solution will be rejected by the threshold selection strategy. This implies that the optimal solution $1^{n}$ will never be found. Thus, the expected running time is infinite for $p_{n}>0$.

For $p_{n}=0$ (i.e., without noise), it is easy to see that in the evolution process, the number of 1 bits $i$ of the solution will never decrease. When using threshold selection $\tau=1, i$ can increase in one step with probability at least $\frac{n-i}{n}\left(1-\frac{1}{n}\right)^{n-1} \geq \frac{n-i}{e n}$, since it is sufficient to flip one 0 bit and keep other bits unchanged. Thus, by using the fitness-level method (Wegener, 2002; Sudholt, 2013), the expected running time is at most $\sum_{i=0}^{n-1} \frac{e n}{n-i}$, i.e., $O(n \log n)$.

Then, we are to analyze the PNT of the ( $1+1$ )-EA using re-evaluation and threshold selection on the OneMax problem for different threshold values $\tau$. Note that the minimal fitness gap for the OneMax problem is 1 . Thus, we first analyze $\tau=1$.

Theorem 8. For the (1+1)-EA with mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, if using re-evaluation with threshold selection $\tau=1$, the PNT is $[0,1]$.

The theorem can be directly derived from the following lemma, which implies the expected running time upper bound $O(n \log n)$ for $p_{n} \leq 1 /(\sqrt{2} e)$ and $O\left(n^{2} \log n\right)$ for $p_{n}>1 /(\sqrt{2} e)$.
Lemma 14. For the (1+1)-EA using re-evaluation with threshold selection $\tau=1$ and mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, the expected running time is upper bounded by $O\left(n^{2} \log n / p_{n}^{2}\right)$ when $p_{n} \in[0,1]$, and specifically $O(n \log n)$ when $p_{n} \leq \frac{1}{\sqrt{2} e}$.

Proof. We use additive drift analysis (i.e., Lemma 3 to prove it. Let $H_{i}=\sum_{j=1}^{i} \frac{1}{j}$ denote the $i$-th harmonic number, and $H_{0}=0$. We first construct a distance function $V(x)$ as $\forall x \in \mathcal{X}=\{0,1\}^{n}, V(x)=H_{|x|_{0}}$, where $|x|_{0}=n-\sum_{i=1}^{n} x_{i}$ is the number of 0 bits of the solution $x$. It is easy to verify that $V\left(x \in \mathcal{X}^{*}=\left\{1^{n}\right\}\right)=0$ and $V\left(x \notin \mathcal{X}^{*}\right)>0$.

Then, we investigate $\mathbb{E} \llbracket V\left(\xi_{t}\right)-V\left(\xi_{t+1}\right) \mid \xi_{t}=x \rrbracket$ for any $x$ with $V(x)>0$ (i.e., $\left.x \notin \mathcal{X}^{*}\right)$. We denote the number of 0 bits of the current solution $x$ by $i(1 \leq i \leq n)$. Let $p_{i, i+d}$ be the probability that the next solution after bit-wise mutation and selection has $i+d(-i \leq d \leq n-i)$ number of 0 bits. We then have

$$
\begin{equation*}
\mathbb{E} \llbracket V\left(\xi_{t}\right)-V\left(\xi_{t+1}\right) \mid \xi_{t}=x \rrbracket=H_{i}-\sum_{d=-i}^{n-i} p_{i, i+d} \cdot H_{i+d} \tag{10}
\end{equation*}
$$

Then, we analyze $p_{i, i+d}$ for $1 \leq i \leq n$. Let $P_{d}$ denote the probability that the offspring solution $x^{\prime}$ by bit-wise mutation on $x$ has $i+d(-i \leq d \leq n-i)$ number of 0 bits. Note that one-bit noise can change the true fitness of a solution by at most 1 , i.e., $\left|f^{N}(x)-f(x)\right| \leq 1$.
(1) When $d \geq 2$, $f^{N}\left(x^{\prime}\right) \leq n-i-d+1 \leq n-i-1 \leq f^{N}(x)$. Because an offspring solution will be accepted only if $f^{N}\left(x^{\prime}\right) \geq f^{N}(x)+1$, the offspring $x^{\prime}$ will be discarded in this case, which implies that $\forall d \geq 2: p_{i, i+d}=0$.
(2) When $d=1$, the offspring solution $x^{\prime}$ will be accepted only if $f^{N}\left(x^{\prime}\right)=n-i \wedge$ $f^{N}(x)=n-i-1$, the probability of which is $p_{n} \frac{i+1}{n} \cdot p_{n} \frac{n-i}{n}$, since it needs to flip one 0 bit of $x^{\prime}$ and flip one 1 bit of $x$. Thus, $p_{i, i+1}=P_{1} \cdot\left(p_{n} \frac{i+1}{n} p_{n} \frac{n-i}{n}\right)$.
(3) When $d=-1$, if $f^{N}(x)=n-i-1$, the probability of which is $p_{n} \frac{n-i}{n^{\prime}}$, the offspring solution $x^{\prime}$ will be accepted, since $f^{N}\left(x^{\prime}\right) \geq n-i+1-1=n-i>f^{N}(x)$; if $f^{N}(x)=$ $n-i \wedge f^{N}\left(x^{\prime}\right) \geq n-i+1$, the probability of which is $\left(1-p_{n}\right) \cdot\left(1-p_{n}+p_{n} \frac{i-1}{n}\right), x^{\prime}$ will be accepted; if $f^{N}(x)=n-i+1 \wedge f^{N}\left(x^{\prime}\right)=n-i+2$, the probability of which is $p_{n} \frac{i}{n} \cdot p_{n} \frac{i-1}{n}, x^{\prime}$ will be accepted; otherwise, $x^{\prime}$ will be discarded. Thus, $p_{i, i-1}=$ $P_{-1} \cdot\left(p_{n} \frac{n-i}{n}+\left(1-p_{n}\right)\left(1-p_{n}+p_{n} \frac{i-1}{n}\right)+p_{n} \frac{i}{n} p_{n} \frac{i-1}{n}\right)$.
(4) When $d \leq-2$, it is easy to see that $p_{i, i+d}>0$.

By applying these probabilities to Eq. 10 , we have

$$
\begin{align*}
& \mathbb{E} \llbracket V\left(\xi_{t}\right)-V\left(\xi_{t+1}\right) \mid \xi_{t}=x \rrbracket  \tag{11}\\
& \geq H_{i}-p_{i, i-1} H_{i-1}-p_{i, i+1} H_{i+1}-\left(1-p_{i, i-1}-p_{i, i+1}\right) H_{i} \\
& =p_{i, i-1} \cdot \frac{1}{i}-p_{i, i+1} \cdot \frac{1}{i+1} \\
& \geq P_{-1}\left(p_{n} \frac{n-i}{n}+p_{n}^{2} \frac{i(i-1)}{n^{2}}\right) \frac{1}{i}-P_{1} p_{n}^{2} \frac{(i+1)(n-i)}{n^{2}} \frac{1}{i+1} .
\end{align*}
$$

We then bound the two mutation probabilities $P_{-1}$ and $P_{1}$. For decreasing the number of 0 bits by 1 in mutation, it is sufficient to flip one 0 bit and keep other bits unchanged, thus we have $P_{-1} \geq \frac{i}{n}\left(1-\frac{1}{n}\right)^{n-1}$. For increasing the number of 0 bits by 1 , it needs to flip one more 1 bit than the number of 0 bits it flips, thus we have

$$
\begin{aligned}
& P_{1}=\sum_{k=1}^{\min \{n-i, i+1\}}\binom{n-i}{k}\binom{i}{k-1} \frac{1}{n^{2 k-1}}\left(1-\frac{1}{n}\right)^{n-2 k+1} \\
& \leq \frac{n-i}{n}\left(1-\frac{1}{n}\right)^{n-1}+\sum_{k=2}^{\min \{n-i, i+1\}} \frac{1}{k!(k-1)!} \frac{(n-i)^{k}}{n^{k}} \frac{i^{k-1}}{n^{k-1}}\left(1-\frac{1}{n}\right)^{n-2 k+1} \\
& \leq \frac{n-i}{n}\left(1-\frac{1}{n}\right)^{n-1}+\frac{i}{n} \cdot \sum_{k=2}^{\min \{n-i, i+1\}} \frac{1}{k!(k-1)!}\left(1-\frac{1}{n}\right)^{n-1} \\
& \leq \frac{n-i}{n}\left(1-\frac{1}{n}\right)^{n-1}+\frac{i}{n} \cdot \sum_{k=2}^{+\infty} \frac{1}{k!}\left(1-\frac{1}{n}\right)^{n-1} \\
& =\frac{n-i}{n}\left(1-\frac{1}{n}\right)^{n-1}+(e-2) \frac{i}{n}\left(1-\frac{1}{n}\right)^{n-1} .
\end{aligned}
$$

By applying these two bounds of $P_{-1}$ and $P_{1}$ to Eq, 11 we can have

$$
\begin{aligned}
& \mathbb{E} \llbracket V\left(\xi_{t}\right)-V\left(\xi_{t+1}\right) \left\lvert\, \xi_{t}=x \rrbracket \geq \frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1} p_{n}^{2}\left(\frac{n-i}{n}+\frac{i(i-1)}{n^{2}}\right)\right. \\
& \quad-\left(\frac{n-i}{n}+(e-2) \frac{i}{n}\right)\left(1-\frac{1}{n}\right)^{n-1} p_{n}^{2} \frac{n-i}{n^{2}} \\
& \geq(3-e) \frac{i}{n^{2}}\left(1-\frac{1}{n}\right)^{n} p_{n}^{2} \\
& \geq \frac{3-e}{2 e} \frac{p_{n}^{2}}{n^{2}} . \quad\left(\text { by } i \geq 1 \text { and }\left(1-\frac{1}{n}\right)^{n} \geq \frac{1}{2 e}\right)
\end{aligned}
$$

Thus, by Lemma3, we get, noting that $V(x) \leq H_{n}<1+\log n$,

$$
\mathbb{E} \llbracket \tau \left\lvert\, \xi_{0} \rrbracket \leq \frac{2 e}{3-e} \frac{n^{2}}{p_{n}^{2}} V\left(\xi_{0}\right) \in O\left(\frac{n^{2} \log n}{p_{n}^{2}}\right)\right.
$$

i.e., the expected running time of the $(1+1)$-EA with $\tau=1$ on the OneMax problem is upper bounded by $O\left(n^{2} \log n / p_{n}^{2}\right)$.

For $p_{n} \leq \frac{1}{\sqrt{2} e}$, we can derive a tighter upper bound $O(n \log n)$ by applying proper bounds of the two probabilities $p_{i, i-1}$ and $p_{i, i+1}$ to Eq.11. From cases (3) and (2) in the analysis of $p_{i, i+d}$, we have

$$
\begin{aligned}
& p_{i, i-1} \geq P_{-1}\left(1-p_{n}\right)^{2} \geq \frac{i}{n}\left(1-\frac{1}{n}\right)^{n-1}\left(1-p_{n}\right)^{2} \\
& p_{i, i+1}=P_{1} p_{n}^{2} \frac{(i+1)(n-i)}{n^{2}} \leq \frac{(i+1)(n-i)^{2}}{n^{3}} p_{n}^{2}
\end{aligned}
$$

where the last inequality is by $P_{1} \leq \frac{n-i}{n}$ since it is necessary to flip at least one 1 bit.

Then, Eq. 11 becomes

$$
\begin{aligned}
& \mathbb{E} \llbracket V\left(\xi_{t}\right)-V\left(\xi_{t+1}\right) \left\lvert\, \xi_{t}=x \rrbracket \geq \frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1}\left(1-p_{n}\right)^{2}-\frac{(n-i)^{2}}{n^{3}} p_{n}^{2}\right. \\
& \geq \frac{1}{n}\left(\frac{1}{e}\left(1-p_{n}\right)^{2}-p_{n}^{2}\right)>0.13 \cdot \frac{1}{n},
\end{aligned}
$$

where the last inequality is because $\frac{1}{e}\left(1-p_{n}\right)^{2}-p_{n}^{2}$ decreases with $p_{n}$ for $p_{n} \leq \frac{1}{\sqrt{2} e}$. Thus, by Lemma 3. we have

$$
\mathbb{E} \llbracket \tau \left\lvert\, \xi_{0} \rrbracket \leq \frac{n}{0.13} V\left(\xi_{0}\right) \in O(n \log n)\right.
$$

i.e., the expected running time of the ( $1+1$ )-EA with $\tau=1$ on the OneMax problem is upper bounded by $O(n \log n)$ for $p_{n} \leq \frac{1}{\sqrt{2}}$.

Then, we analyze the effect of a relatively large threshold value $\tau=2$ on the PNT.
Theorem 9. For the ( $1+1$ )-EA with mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, if using re-evaluation with threshold selection $\tau=2$, the PNT is $[1 / \Theta(\operatorname{poly}(n)), 1-1 / \Theta(\operatorname{poly}(n))]$.

The theorem can be directly derived from the following two lemmas.
Lemma 15. For the ( $1+1$ )-EA using re-evaluation with threshold selection $\tau=2$ and mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, the expected running time is upper bounded by $O\left(n \log n /\left(p_{n}\left(1-p_{n}\right)\right)\right)$.
Proof. Let $i(0 \leq i \leq n)$ denote the number of 0 bits of the current solution $x$. Here, an offspring $x^{\prime}$ will be accepted only if $f^{N}\left(x^{\prime}\right)-f^{N}(x) \geq 2$. As in the proof of Lemma 14 . we can derive $\quad \forall d \geq 1: p_{i, i+d}=0 ; \quad \forall d \geq 2: p_{i, i-d}>0$;

$$
p_{i, i-1}=P_{-1}\left(p_{n} \frac{n-i}{n}\left(1-p_{n}+p_{n} \frac{i-1}{n}\right)+\left(1-p_{n}\right)\left(p_{n} \frac{i-1}{n}\right)\right) .
$$

Thus, $i$ never increases, and it decreases in one step with probability at least

$$
p_{i, i-1} \geq \frac{i}{n}\left(1-\frac{1}{n}\right)^{n-1}\left(\left(1-p_{n}\right) p_{n}\left(1-\frac{1}{n}\right)+p_{n}^{2} \frac{(n-i)(i-1)}{n^{2}}\right) \geq \frac{1}{2 e}\left(1-p_{n}\right) p_{n} \frac{i}{n} .
$$

Then, the expected steps until $i=0$ (i.e., the optimal solution is found) is at most

$$
\sum_{i=1}^{n} \frac{2 e n}{i\left(1-p_{n}\right) p_{n}} \in O\left(\frac{n \log n}{p_{n}\left(1-p_{n}\right)}\right)
$$

Lemma 16. For the ( $1+1$ )-EA using re-evaluation with threshold selection $\tau=2$ and mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, the expected running time is lower bounded by $\Omega\left(n \log n+n /\left(p_{n}\left(1-p_{n}\right)\right)\right)$.
Proof. The proof is very similar to that of Lemma 12 except the calculation of $\mathbb{E} \llbracket T_{2} \mid \bar{A} \rrbracket$. Here we first use a different but simple idea to show that $P(\bar{A}) \geq 1 / 2$. For any evolutionary path with the event $A$ happening, it has to flip at least two bits in the last step
for finding the optimal solution, because any solution $x^{\prime}$ with $\left|x^{\prime}\right|_{1}=n-1$ is never found. Thus, $P(A) \leq\binom{ n}{2} \frac{1}{n^{2}} \leq \frac{1}{2}$. Since $P(\bar{A})+P(A)=1$, we have $P(\bar{A}) \geq 1 / 2$.

Then, we analyze $\mathbb{E} \llbracket T_{2} \mid \bar{A} \rrbracket$, which is the expected running time for finding the optimal solution when starting from a solution $x$ with $|x|_{1}=n-1$ (i.e., $|x|_{0}=1$ ). From the upper bound analysis in the proof of Lemma 15, we know that once a solution $x$ with $|x|_{0}=1$ is found, it will always satisfy $|x|_{0}=1$ before finding the optimal solution, because $\forall d \geq 1: p_{i, i+d}=0$. Meanwhile, the optimal solution (i.e., $|x|_{0}=0$ ) will be found in one step with probability $p_{1,0}=\frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1} p_{n}\left(1-p_{n}\right)\left(1-\frac{1}{n}\right) \leq \frac{p_{n}\left(1-p_{n}\right)}{e n}$. We then have $\mathbb{E} \llbracket T_{2} \left\lvert\, \bar{A} \rrbracket \geq \frac{e n}{p_{n}\left(1-p_{n}\right)}\right.$. Thus, the lemma holds.

For larger threshold values $\tau>2$, we have:
Theorem 10. For the (1+1)-EA with mutation probability $\frac{1}{n}$ on the OneMax problem under one-bit noise, if using re-evaluation with threshold selection $\tau>2$, the PNT is $\emptyset$.

Proof. Let $i=|x|_{0}$ for the current solution $x$. An offspring solution $x^{\prime}$ will be accepted only if $f^{N}\left(x^{\prime}\right)-f^{N}(x) \geq \tau>2$. Then, we can have

$$
\forall d \geq 1: p_{i, i+d}=0 ; \quad p_{i, i-1}= \begin{cases}P_{-1} \cdot\left(p_{n} \frac{n-i}{n} p_{n} \frac{i-1}{n}\right), & \text { if } \tau=3 \\ 0, & \text { otherwise }\end{cases}
$$

In the evolutionary process, it is easy to see that there exists some positive probability that a solution $x$ with $|x|_{0}=1$ is found (i.e., $i=1$ ). Once it happens, $i=1$ will always hold because $p_{1,0}=0$ and $p_{1,1+d}=0$ for any $d \geq 1$. In such a situation, the optimal solution $1^{n}$ will never be found. Thus, the expected running time is infinite for any $p_{n} \in[0,1]$.

### 4.3 Smooth Threshold Selection

We have shown that for the (1+1)-EA solving the OneMax problem under one-bit noise, the re-evaluation with threshold selection $\tau=1$ can improve the PNT to $[0,1]$, which means that the expected running time of the $(1+1)$-EA is always polynomial regardless of the noise strength. Under asymmetric one-bit noise, we will prove that all the above strategies are however not effective when the noise probability $p_{n}$ equals 1 , as shown in Theorem 11 .
Theorem 11. For the (1+1)-EA with mutation probability $\frac{1}{n}$ on the OneMax problem under asymmetric one-bit noise, if using threshold selection $\tau \geq 0$ with either singleevaluation or re-evaluation, the expected running time is at least exponential for $p_{n}=1$.

Proof. We analyze the expected running time for each strategy, respectively.
For single-evaluation with threshold selection $\tau \geq 0$, where single-evaluation with threshold selection $\tau=0$ is equivalent to single-evaluation alone, it is easy to see that there exists some positive probability that a solution $x$ with $|x|_{0}=1$ and $f^{N}(x)=n$ is found. Because the fitness is not re-evaluated; $f^{N}\left(1^{n}\right)=n-1$ due to $p_{n}=1$; and $f^{N}(x) \leq n-1$ for $x$ with $|x|_{0} \geq 2$, it will always keep in such a state. Thus, the expected running time for finding the optimal solution $1^{n}$ is infinite.

For re-evaluation with threshold selection $\tau=0$ (i.e., re-evaluation alone), we use the simplified drift theorem (i.e., Lemma 4) to prove an exponential running time lower bound. Let $X_{t}$ be the number of 0 bits of the solution after $t$ iterations of the $(1+1)$-EA. We consider the interval $\left[0, n^{1 / 4}\right]$, i.e., the parameters $a=0$ (i.e., the global optimum) and $b=n^{1 / 4}$ in Lemma 4 . Then, we analyze the drift $\mathbb{E} \llbracket X_{t}-X_{t+1} \mid X_{t}=i \rrbracket$ for $1 \leq i<n^{1 / 4}$. Let $p_{i, i+d}$ denote the probability that the next solution after bit-wise
mutation and selection has $i+d(-i \leq d \leq n-i)$ number of 0 bits (i.e., $X_{t+1}=i+d$ ). We thus have

$$
\begin{equation*}
\mathbb{E} \llbracket X_{t}-X_{t+1} \mid X_{t}=i \rrbracket=\sum_{d=1}^{i} d \cdot p_{i, i-d}-\sum_{d=1}^{n-i} d \cdot p_{i, i+d} . \tag{12}
\end{equation*}
$$

Let $P_{d}$ denote the probability that the offspring solution generated by bit-wise mutation has $i+d$ number of 0 bits. Using the same analysis procedure for $p_{i, i+d}$ as in the proof of Lemma 14 , we have, noting that $p_{n}=1$ and $\tau=0$ here,

$$
\begin{array}{rrr}
\forall d \geq 3, p_{i, i+d}=0 ; & p_{i, i+2}=P_{2} / 4 ; & p_{i, i+1}=P_{1} / 4 ; \\
p_{i, i-1} \leq 3 P_{-1} / 4 ; & \forall 2 \leq d \leq i, p_{i, i-d}=P_{-d} . &
\end{array}
$$

Furthermore, $P_{1} \geq \frac{n-i}{n}\left(1-\frac{1}{n}\right)^{n-1} \geq \frac{n-i}{e n}$, since it is sufficient to flip one 1 bit and keep other bits unchanged, and $\left.P_{-d} \leq \begin{array}{c}i \\ d\end{array}\right) \frac{1}{n^{d}}$, since it is necessary to flip at least $d$ number of 0 bits. By applying these probabilities to Eq 12, we can have

$$
\begin{aligned}
& \mathbb{E} \llbracket X_{t}-X_{t+1} \left\lvert\, X_{t}=i \rrbracket \leq \frac{3 P_{-1}}{4}+\sum_{d=2}^{i} d \cdot P_{-d}-\frac{P_{1}}{4}\right. \\
& \leq \frac{3 i}{4 n}+\sum_{d=2}^{i} d \cdot\binom{i}{d} \frac{1}{n^{d}}-\frac{n-i}{4 e n}=\frac{i}{n}\left(\left(1+\frac{1}{n}\right)^{i-1}+\frac{1}{4 e}-\frac{1}{4}\right)-\frac{1}{4 e} \\
& =-\frac{1}{4 e}+O\left(\frac{n^{1 / 4}}{n}\right) \cdot\left(\text { since } i<n^{1 / 4}\right)
\end{aligned}
$$

Thus, $\mathbb{E} \llbracket X_{t}-X_{t+1} \mid X_{t}=i \rrbracket=-\Omega(1)$, which implies that the condition 1 of Lemma 4 holds. For its condition 2, we need to investigate $P\left(\left|X_{t+1}-X_{t}\right| \geq j \mid X_{t} \geq 1\right)$. Because it is necessary to flip at least $j$ bits, we have

$$
P\left(\left|X_{t+1}-X_{t}\right| \geq j \mid X_{t} \geq 1\right) \leq\binom{ n}{j} \frac{1}{n^{j}} \leq \frac{1}{j!} \leq 2 \cdot \frac{1}{2^{j}},
$$

which implies that the condition 2 of Lemma 4 holds with $\delta=1$ and $r(l)=2$. Note that $l=b-a=n^{1 / 4}$. Thus, by Lemma 4 , the probability that the running time is $2^{O\left(n^{1 / 4}\right)}$ when starting from a solution $x$ with $|x|_{0} \geq n^{1 / 4}$ is exponentially small. Due to the uniform initial distribution, the probability that the initial solution $x$ has $|x|_{0}<n^{1 / 4}$ is exponentially small by Chernoff's inequality. Thus, the expected running time is exponential.

For re-evaluation with threshold selection $\tau=1$, we use the same analysis procedure as $\tau=0$. The only difference is the calculation of $p_{i, i+d}$ :

$$
\begin{array}{rr}
\forall d \geq 2, p_{i, i+d}=0 ; & p_{i, i+1}=P_{1} / 4 ; \\
p_{i, i-1} \leq 3 P_{-1} / 4 ; & p_{i, i-2} \leq 3 P_{-2} / 4 ;
\end{array} \quad \forall 3 \leq d \leq i, p_{i, i-d}=P_{-d} .
$$

It is easy to verify that the analysis of the two conditions of Lemma 4 will not be affected. Thus, we derive the same result as $\tau=0$ : the expected running time is exponential.

For re-evaluation with threshold selection $\tau \geq 2$, we have

$$
\forall d \geq 1, p_{i, i+d}=0 ; \quad \forall 2 \leq d \leq i, p_{i, i-d}>0
$$

For $p_{i, i-1}$, we need to consider two cases:

$$
\text { (1) for } i \geq 2, p_{i, i-1}=P_{-1} / 4 \text {; }
$$

(2) for $i=1, p_{1,0}=0$.

It is easy to see that there exists some positive probability that a solution $x$ with $|x|_{0}=$ 1 is found. For $\tau \geq 2$, we have $\forall d \geq 1, p_{1,1+d}=0$ and $p_{1,0}=0$. Thus, it will always keep in such a state (i.e., $|x|_{0}=1$ ), which implies that the expected running time for finding the optimal solution $1^{n}$ is infinite.

Therefore, the re-evaluation with threshold selection is ineffective in this case. When the threshold $\tau \leq 1$, it has a too large probability of accepting false progresses, which leads to a negative drift and thus the exponential running time. When $\tau \geq 2$, although the probability of accepting false progresses is 0 (i.e., $\forall d \geq 1, p_{i, i+d}=0$ ), it has a too small probability of accepting true progresses (i.e., $p_{1,0}=0$ ), which leads to the infinite running time. However, setting $\tau$ between 1 and 2 is useless, because the minimum fitness gap is 1 , which makes a value of $\tau \in(1,2)$ equivalent to $\tau=2$.

We propose the smooth threshold selection as in Definition 10, which modifies the original threshold selection by changing the hard threshold value to a smooth one. The "smooth" means that the offspring solution will be accepted with some probability when the fitness gap between the offspring and the parent is just the threshold. For example, (1) $+(0.1)$-smooth threshold selection accepts the offspring solution with probability 0.9 when the fitness gap is 1 ; this makes a fractional threshold 1.1 effective. Such a strategy of accepting new solutions probabilistically based on the fitness is similar to the acceptance strategy of simulated annealing (Kirkpatrick, 1984). We are to show that using the smooth threshold selection with proper threshold values can improve the PNT to $[0,1]$ in this case.
Definition 10 (Smooth Threshold Selection). Let $\delta$ be the gap between the fitness of the offspring solution $x^{\prime}$ and the parent solution x, i.e., $\delta=f\left(x^{\prime}\right)-f(x)$. Given a threshold
$(A)+(B)$ with $B \in[0,1]$, the selection process will behave as follows:
(1) if $\delta<A, x^{\prime}$ will be rejected;
(2) if $\delta=A, x^{\prime}$ will be accepted with probability $1-B$;
(3) if $\delta>A, x^{\prime}$ will be accepted.

In the following analysis, we will view the evolutionary process as a random walk on a graph (i.e., Algorithm (4), which has often been used for analyzing randomized search heuristics, e.g., in (Giel and Wegener 2003 Neumann and Witt 2010). Lemma 17 gives an upper bound on the expected steps for a random walk to visit each vertex of a graph at least once.
Algorithm 4 (Random Walk). Given an undirected connected graph $G=(V, E)$ with vertex set $V$ and edge set $E$, it consists of the following steps:

1. start at a vertex $v \in V$.
2. Repeat until the termination condition is met
3. choose a neighbor $u$ ofv uniformly at random.
4. setv $:=u$.

Lemma 17 ((Aleliunas et al., 1979)). Given an undirected connected graph $G=(V, E)$, the expected number of steps until each vertex $v \in V$ has been visited at least once for a random walk on $G$ is upper bounded by $2|E|(|V|-1)$.

We first analyze a smooth threshold depending on the current search point.

Theorem 12. For the ( $1+1$ )-EA with mutation probability $\frac{1}{n}$ on the OneMax problem under asymmetric one-bit noise, if using re-evaluation with $(1)+\left(1-\frac{|x|_{0}}{2 e n}\right)$-smooth threshold selection, the PNT is $[0,1]$.

Proof. Let $i(0 \leq i \leq n)$ denote the number of 0 bits of the current solution $x$. We first analyze $p_{i, i+d}$ as that analyzed in the proof of Lemma 14 Note that there are two differences in the analyses, which are caused by different threshold and noise settings, respectively. Due to the threshold difference, the acceptance probability is different when the fitness gap between the offspring $x^{\prime}$ and the parent solution $x$ is $1: x^{\prime}$ will be accepted with probability $\frac{|x|_{0}}{2 e n}$ here, while it will be always accepted in the proof of Lemma 14. Due to the noise difference, the probability of flipping a 0 or 1 bit is different: a random 0 or 1 bit will be flipped with an equal probability of $\frac{1}{2}$ here, while a uniformly randomly chosen bit will be flipped in the proof of Lemma 14 .

Thus, we can similarly derive the value of $p_{i, i+d}$ for $1 \leq i \leq n-1$. It is easy to see that $\forall d \geq 2: p_{i, i+d}=0$, and $p_{i, i-d}>0$. For $p_{i, i+1}$, the offspring solution $x^{\prime}$ will be accepted only if $f^{N}\left(x^{\prime}\right)=n-i \wedge f^{N}(x)=n-i-1$, and the acceptance probability is $\frac{i}{2 e n}$. The probability of $f^{N}\left(x^{\prime}\right)=n-i$ is at most $p_{n}$, since it needs to flip one 0 bit of $x^{\prime}$ in noise; the probability of $f^{N}(x)=n-i-1$ is $p_{n} \frac{1}{2}$, since it needs to flip one 1 bit of $x$. Thus, $p_{i, i+1} \leq P_{1}\left(p_{n} \frac{1}{2} \cdot p_{n}\right) \cdot \frac{i}{2 e n}$. For $p_{i, i-1}$, we need to consider two cases:
(1) $2 \leq i \leq n-1$. If $f^{N}(x)=n-i-1$, the probability of which is $p_{n} \frac{1}{2}$, there are three cases for the offspring solution $x^{\prime}$ : if $f^{N}\left(x^{\prime}\right)=n-i$ (the probability is $p_{n} \frac{1}{2}$ ), the acceptance probability $\frac{i}{2 e n}$, since $f^{N}\left(x^{\prime}\right)=f^{N}(x)+1$; if $f^{N}\left(x^{\prime}\right)=n-i+1$ or $f^{N}\left(x^{\prime}\right)=n-i+2$ (the probability is $\left(1-p_{n}\right)+p_{n} \frac{1}{2}$ ), the acceptance probability 1 , since $f^{N}\left(x^{\prime}\right)>f^{N}(x)+1$. If $f^{N}(x)=n-i$, the probability of which is $1-p_{n}$, there are two cases for the acceptance of $x^{\prime}$ : if $f^{N}\left(x^{\prime}\right)=n-i+1$ (the probability is $\left(1-p_{n}\right)$ ), the acceptance probability is $\frac{i}{2 e n}$; if $f^{N}\left(x^{\prime}\right)=n-i+2$ (the probability is $p_{n} \frac{1}{2}$ ), the acceptance probability is 1 . If $f^{N}(x)=n-i+1$, the probability of which is $p_{n} \frac{1}{2}, x^{\prime}$ will be accepted only if $f^{N}\left(x^{\prime}\right)=n-i+2$ (the probability is $p_{n} \frac{1}{2}$ ), and the acceptance probability is $\frac{i}{2 e n}$. Thus, we have

$$
\begin{aligned}
p_{i, i-1}=P_{-1} & \left(p_{n} \frac{1}{2}\left(p_{n} \frac{1}{2} \cdot \frac{i}{2 e n}+\left(1-p_{n}\right)+p_{n} \frac{1}{2}\right)\right. \\
& \left.+\left(1-p_{n}\right)\left(\left(1-p_{n}\right) \cdot \frac{i}{2 e n}+p_{n} \frac{1}{2}\right)+p_{n} \frac{1}{2} p_{n} \frac{1}{2} \cdot \frac{i}{2 e n}\right) .
\end{aligned}
$$

(2) $i=1$. The analysis is similar to case (1). The only difference is that when the noise happens, $f^{N}\left(x^{\prime}\right)=n-i$ with probability 1 here since $\left|x^{\prime}\right|_{0}=i-1=0$ reaches the extreme case, while in case (1), $f^{N}\left(x^{\prime}\right)=n-i$ or $n-i+2$ with an equal probability of $\frac{1}{2}$. Note that $P_{-1}=\frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1}$ for $i=1$. Thus, we have

$$
p_{1,0}=\frac{1}{n}\left(1-\frac{1}{n}\right)^{n-1} \cdot\left(p_{n} \frac{1}{2}\left(p_{n} \frac{1}{2 e n}+\left(1-p_{n}\right)\right)+\left(1-p_{n}\right)\left(1-p_{n}\right) \frac{1}{2 e n}\right) .
$$

Our goal is to reach $i=0$ (i.e., the global optimum). Starting from $i=1, i$ will reach 0 in one step with probability

$$
p_{1,0} \geq \frac{1}{e n} \cdot \frac{1}{2 e n} \cdot\left(\frac{p_{n}^{2}}{2}+\left(1-p_{n}\right)^{2}\right) \geq \frac{1}{6 e^{2} n^{2}} . \quad\left(b y 0 \leq p_{n} \leq 1\right)
$$

Thus, for reaching $i=0$, we need to reach $i=1$ for $O\left(n^{2}\right)$ times in expectation.

Then, we analyze the expected running time until $i=1$. In this process, we can pessimistically assume that $i=0$ will never be reached, because our final goal is to get the running time upper bound for reaching $i=0$. For $2 \leq i \leq n-1$, we have

$$
\frac{p_{i, i-1}}{p_{i, i+1}} \geq \frac{P_{-1} \cdot\left(p_{n} \frac{1}{2} p_{n} \frac{1}{2}\right)}{P_{1} \cdot\left(p_{n} \frac{1}{2} p_{n}\right) \cdot \frac{i}{2 e n}} \geq \frac{\frac{i}{n}\left(1-\frac{1}{n}\right)^{n-1} \cdot\left(p_{n} \frac{1}{2} p_{n} \frac{1}{2}\right)}{\frac{n-i}{n} \cdot\left(p_{n} \frac{1}{2} p_{n}\right) \cdot \frac{i}{2 e n}} \geq \frac{n}{n-i}>1 .
$$

Again, we can pessimistically assume that $p_{i, i-1}=p_{i, i+1}$ and $\forall d \geq 2, p_{i, i-d}=0$, because we are to get the upper bound on the expected running time until $i=1$. Then, we can view the evolutionary process for reaching $i=1$ as a random walk on the path $\{1,2, \ldots, n-1, n\}$. We call a step that jumps to the neighbor state a relevant step. Thus, by Lemma 17 it needs at most $2(n-1)^{2}$ expected relevant steps to reach $i=1$. Because the probability of a relevant step is at least

$$
p_{i, i-1} \geq \frac{i}{e n} \cdot \frac{i}{2 e n}\left(\left(1-p_{n}\right)^{2}+p_{n}^{2} \frac{1}{2}\right) \geq \frac{4}{2 e^{2} n^{2}} \cdot \frac{1}{3}
$$

the expected running time for a relevant step is $O\left(n^{2}\right)$. Then, the expected running time for reaching $i=1$ is $O\left(n^{4}\right)$.

Thus, the expected running time of the whole optimization process is $O\left(n^{6}\right)$ for any $p_{n} \in[0,1]$, and then the PNT is $[0,1]$.

Although we have shown that the $(1)+\left(1-\frac{|x|_{0}}{2 e n}\right)$-smooth threshold selection can improve the PNT to be $[0,1]$, the threshold value depends on the current search point; this implies that designing proper thresholds may require problem knowledge, which might be unrealistic. Thus, we are then to show that a smooth threshold without dependence on the current search point can also be effective. The proof of Theorem 13 is the same as that of Theorem 12 except that the acceptance probability for the fitness gap 1 is $\frac{1}{2 e n}$ instead of $\frac{|x|_{0}}{2 e n}$.
Theorem 13. For the (1+1)-EA with mutation probability $\frac{1}{n}$ on the OneMax problem under asymmetric one-bit noise, if using re-evaluation with $(1)+\left(1-\frac{1}{2 e n}\right)$-smooth threshold selection, the PNT is $[0,1]$.

We draw an intuitive understanding from the proof of Theorem 12 of why the smooth threshold selection can be better than the original ones. By changing the hard threshold to be smooth, it can not only make the probability of accepting a false better solution in one step small enough, i.e., $p_{i, i-1} \geq p_{i, i+1}$, but also make the probability of producing real progress in one step large enough, i.e., $p_{i, i-1}$ is not small.

## 5 Experiments

In this section we will employ experiments to complement the theoretical analyses. For any given configuration, we will run the EA 1000 times independently. In each run, we record the number of fitness evaluations until an optimal solution is found. Then the running time values of the 1000 runs are averaged as the estimation of the expected running time, called as the estimated ERT.

### 5.1 On Noise Helpful Cases

We first present the experiment results on the Trap and the Peak problem to verify Theorems 2 and 3 .

- As for Theorem 2, we conduct experiments using the $(1+n)-E A$, a specific case of the $(1+\lambda)$-EA with $n$ being the dimensionality of the problem, on the Trap problem.

We estimate the expected running time of the $(1+n)$-EA starting from the solution $x$ with $|x|_{0}=i$ for each $i(0 \leq i \leq n)$. Following Theorem 2 we compare the estimated ERT of the $(1+n)$-EA without noise, with additive noise and with multiplicative noise, respectively. For the mutation probability of the $(1+n)$-EA, we use the common setting $p=\frac{1}{n}$. For additive noise, $\delta_{1}=-n$ and $\delta_{2}=n-1$, and for multiplicative noise, $\delta_{1}=0.1$ and $\delta_{2}=10$. The results for $n=5,6,7$ are plotted in Figure 1. We can observe that the curves by these two kinds of noise are always under the curve without noise, which is consistent with our theoretical result. Note that, the three curves meet at the first point, since the initial solution with $|x|_{0}=0$ is the optimal solution and then ERT $=1$.


Figure 1: Estimated ERT comparison for the $(1+n)$-EA solving the Trap problem with or without noise.

- As for Theorem 3. we conduct experiments using the $(1+1)-E A^{*}$ on the Peak problem. The one-bit noise is set with $p_{n}=0.5$. The results for $n=6,7,8$ are plotted in Figure 2. We can observe that the curve with one-bit noise is always under the curve without noise when $|x|_{0}$ is large enough, which is consistent with the analysis result. We also run the $(1+1)$-EA on the Peak problem, the results of which are shown in Figure 3. The observation that the curve with noise is always under that without noise agrees with Conjecture 1 .


Figure 2: Estimated ERT comparison for the (1+1)-EA* solving the Peak problem with or without noise.

We have shown that noise can make deceptive and flat problems easier for EAs. For deceptive problems, it is intuitively because the EA searches along the deceptive direction while noise can add some randomness to make the EA have some possibility to run along the right direction; for flat problems, the EA has no guided information for search while under some situations noise can make the EA have a larger probability to run along the right direction than the wrong direction.

Note that, though the Trap and Peak problems are respectively extremely deceptive and flat, in real applications we often encounter optimization problems with some degree of deceptiveness and flatness. We then test whether the finding on the extreme cases also holds on other problems. We employ the ( $1+1$ )-EA with mutation


Figure 3: Estimated ERT comparison for the (1+1)-EA solving the Peak problem with or without noise.
probability $\frac{1}{n}$ on the minimum spanning tree (MST) problem. Given an undirected connected graph $G=(V, E)$ on $n$ vertices and $m$ edges, the MST problem is to find a connected graph $G^{\prime}=\left(V, E^{\prime} \subseteq E\right)$ with the minimal sum of edge weights. As in (Neumann and Wegener 2007), a solution $x$ is represented by a Boolean string of length $m$, i.e., $x \in\{0,1\}^{m}$, where $x_{i}=1$ means that the edge $i$ is selected by $x$; and the following fitness function is used for minimization,

$$
f(x)=(c(x)-1) w_{u b}^{2}+\left(\sum_{i=1}^{m} x_{i}-n+1\right) w_{u b}+\sum_{i=1}^{m} x_{i} w_{i},
$$

where $c(x)$ is the number of connected components of the subgraph represented by $x$, and $w_{u b}=n^{2} \cdot \max \left\{w_{i} \mid 1 \leq i \leq m\right\}$. It is easy to see that each non-minimum spanning tree is local optimal in the Hamming space, thus the MST problem is a multimodal problem with local deceptiveness.

We conduct experiments to compare the $(1+1)$-EA without noise and with onebit noise ( $p_{n}=0.5$ ) on the graphs with the number of edges $m \in \Theta(n), \Theta(n \sqrt{n})$ and $\Theta\left(n^{2}\right)$ respectively. Let $v_{1}, v_{2}, \ldots, v_{n}$ denote the $n$ nodes.
sparse graph: we use cyclic graph where $v_{1}$ is connected with $v_{n}$ and $v_{2}, v_{i}(1<i<n)$ is connected with $v_{i-1}$ and $v_{i+1}$, and $v_{n}$ is connected with $v_{n-1}$ and $v_{1}$. Thus, $m=n$.
moderate graph: we use the graph where $v_{i}$ is connected with $v_{i+1}, v_{i+2}, \ldots, v_{i+\lfloor\sqrt{n}\rfloor}$ for $1 \leq i \leq n-\lfloor\sqrt{n}\rfloor$. Thus, $m=(n-\lfloor\sqrt{n}\rfloor)\lfloor\sqrt{n}\rfloor$.
dense graph: we use complete graph where each node is connected with all the other nodes. Thus, $m=n(n-1) / 2$.

For each type of graph, and each independent run, the graph is constructed by setting the weight of each edge be an integer randomly selected from $[1, n]$. The experiment results are plotted in Figure 4. We can observe that the curves with one-bit noise can appear under the curves without noise, which supports our finding that noise can make problems easier for EAs.

We have derived that in the deceptive and the flat cases noise can make the problem easier for EAs. Noticing that the deceptiveness and the flatness are two factors that can block the search of EAs, we hypothesize that the negative effect by noise decreases as the problem hardness increases, and noise will bring a positive effect when the problem is quite hard. The effect of noise can be measured by the estimated ERT gap,

$$
\text { gap }=\left(\mathbb{E} \llbracket \tau \rrbracket-\mathbb{E} \llbracket \tau^{\prime} \rrbracket\right) / \mathbb{E} \llbracket \tau^{\prime} \rrbracket,
$$

where $\mathbb{E} \llbracket \tau \rrbracket$ and $\mathbb{E} \llbracket \tau^{\prime} \rrbracket$ denote the expected running time of the EA optimizing the problem with and without noise, respectively. Note that the noise is harmful if the gap is positive, and is helpful if the gap is negative.


Figure 4: Estimated ERT comparison for the $(1+1)$-EA solving the MST problem with or without noise.

Definition 11 (Jump ${ }_{m, n}$ Problem). Jump ${ }_{m, n}$ Problem of size $n$ with $1 \leq m \leq n$ is to find an $n$ bits binary string $x^{*}$ such that

$$
x^{*}=\arg \max _{x \in\{0,1\}^{n}}\left(\operatorname{Jump}_{m, n}(x)=\left\{\begin{array}{ll}
m+\sum_{i=1}^{n} x_{i} & \text { if } \sum_{i=1}^{n} x_{i} \leq n-m \text { or }=n \\
n-\sum_{i=1}^{n} x_{i} & \text { otherwise }
\end{array}\right)\right.
$$

To verify our hypothesis, we test the (1+1)-EA with mutation probability $\frac{1}{n}$ on the $J^{\prime} \mathrm{Jmp}_{m, n}$ problem as in Definition 11, as well as the MST problem.

- The Jump ${ }_{m, n}$ problem has an adjustable difficulty and can be configured as the OneMax problem when $m=1$ and the Trap problem when $m=n$. It is known that the expected running time of the (1+1)-EA on the Jump ${ }_{m, n}$ problem is $\Theta\left(n^{m}+\right.$ $n \log n$ ) (Droste et al. 2002), which implies that the Jump ${ }_{m, n}$ problem with larger value of $m$ is harder. In the experiment, we set $n=10$, and for noise, we use the additive noise with $\delta_{1}=-0.5 n \wedge \delta_{2}=0.5 n$, the multiplicative noise with $\delta_{1}=1 \wedge \delta_{2}=2$, and the one-bit noise with $p_{n}=0.5$, respectively. The experiment results on gap values are plotted in Figure 5 . We can have a clear observation that the gap values for larger $m$ are lower (i.e., the negative effect by noise decreases as the problem hardness increases).


Figure 5: Estimated ERT gap for the (1+1)-EA solving the Jump ${ }_{m, 10}$ problem with or without noise.

- The MST problem with sparse, moderate and dense graphs are tested. The expected running time of the $(1+1)$-EA on the MST problem has been proven to be $O\left(m^{2}\left(\log n+\log w_{\max }\right)\right)$ (Neumann and Wegener, 2007, Doerr et al. 2012b). With the assumption that this theoretical upper bound is tight, the hardness order of the three types of graphs is "sparse" < "moderate" < "dense". The results are plotted in Figure6, We can observe that the height order of the curves is "sparse" > "moderate" > "dense", which is consistent with our hypothesis.

Experiments on both the artificial Jump ${ }_{m, n}$ problem and the combinatorial MST


Figure 6: Estimated ERT gap for the (1+1)-EA solving the MST problem with or without one-bit noise.
problem reveal the same trend that the effect of the noise can be related to the hardness of the problem to the EA. When the problem is hard, the noise can be helpful to the EA, and thus the noise handling is not necessary.

### 5.2 On Noise Harmful Cases

We first verify the theoretical result that any noise will do harm to the OneMax problem. The experiment setting is the same as that for the $(1+\lambda)$-EA on the Trap problem in Section5.1. The results for $n=10,20,30$ are plotted in Figure 7. We can observe that the curve by any noise is always above the curve without noise, which is consistent with our theoretical result.


Figure 7: Estimated ERT comparison for the $(1+n)$-EA on the OneMax problem with or without noise.

The smooth threshold selection allows us to choose a fractional threshold, and through the running time analysis on the OneMax problem, we have shown that the fractional threshold is essential for the EA to keep efficient with the noise. We then run the (1+1)-EA with mutation probability $\frac{1}{n}$ on the OneMax problem under asymmetric one-bit noise. For the noise strength, we set $p_{n}$ to be the maximum value 1 . We test the smooth threshold values $(A)+(B)$ with $A=0,1$ and $B=0,0.1, \ldots, 0.9,1$, which correspond to the threshold value set $\{0,0.1,0.2, \ldots, 2\}$ on the $x$-axis. The results are plotted in Figure 8. Note that, $x=0$ corresponds to the re-evaluation strategy, and $x=1,2$ corresponds to the original threshold selection with $\tau=x$. We can observe that the curves reach the lowest point when $x \approx 1.9$, which corresponds to a smooth threshold. We have also tested the single-evaluation strategy, and the running time is empirically shown to be infinite. Thus, these empirical observations agree with our theoretical analyses.

To verify whether the fractional threshold is also useful in practice, we then carry out experiments to test the $(1+1)$-EA with mutation probability $\frac{1}{n}$ solving the maxi-


Figure 8: Estimated ERT for the (1+1)-EA with different threshold values solving the OneMax problem under asymmetric one-bit noise $p_{n}=1$.
mum matching problem under one-bit noise. For the noise strength, we set $p_{n}=1$. Given an undirected graph $G=(V, E)$ on $n$ vertices and $m$ edges, a matching is a subset $E^{\prime}$ of the edge set $E$, such that no two edges in $E^{\prime}$ share a common vertex. The maximum matching problem is to find a matching with the largest number of edges. As in (Giel and Wegener, 2003, 2006), a solution is represented as a Boolean string $x \in\{0,1\}^{m}$, where $x_{i}=1$ means that the edge $i$ is selected by $x$, and the following fitness function is used for maximization,

$$
f(x)=\sum_{i=1}^{m} x_{i}-c \cdot \sum_{v \in V} p(v, x)
$$

where $p(v, x)=\max \{0, d(v, x)-1\}, d(v, x)$ is the degree of the vertex $v$ on the subgraph represented by $x$, and $c \geq m+1$ is a penalty coefficient which makes any matching have a larger fitness than any non-matching. Note that (Qian et al., 2015b) recently discloses that a variable solution representation can be better than Boolean string.

We test on the dense (complete) graph with the number of nodes $n=7,8,9$. We test the smooth threshold values $(A)+(B)$ with $A=0,1,2,3$ and $B=0,0.1, \ldots, 0.9,1$, which correspond to the threshold value set $\{0,0.1,0.2, \ldots, 4\}$ on the $x$-axis. The results are plotted in Figure 9 . We can observe that the curves reach the lowest point when $x$ is fractional between 1 and 2, which corresponds to a smooth threshold. The running time using the single-evaluation strategy is empirically shown to be infinite. These empirical observations suggest that smooth threshold selection can lead to better performance in noisy environments.


Figure 9: Estimated ERT for the (1+1)-EA with different threshold values solving the maximum matching problem under one-bit noise $p_{n}=1$.

## 6 Discussions and Conclusions

This work studies some theoretical issues of noisy optimization using EAs.

First, we have proven that on deceptive and flat problems, the noise can make the optimization easier for EAs. Experiments on the minimum spanning tree problem (a multimodal problem with local deceptiveness) support our theoretical findings. As deceptive and flat problems are EA-hard, while the noise can also be shown harmful on the EA-easy problem OneMax, we hypothesize that the negative effect by noise decreases as the problem hardness increases, and noise can even bring a positive effect when the problem is quite hard. This hypothesis is supported by experiments on the $\mathrm{Jump}_{m, n}$ problem and the minimum spanning tree problem, both of which have an adjustable difficulty parameter.

In problems where the noise has a negative effect, we studied the usefulness of two commonly employed noise-handling strategies: re-evaluation and threshold selection. We took the OneMax problem as the representative problem, where the noise significantly harms the expected running time of the $(1+1)$-EA. We used the PNT as the performance measure, and analyzed the PNT of each EA under one-bit noise, as shown in Table 1

The re-evaluation strategy seems to be a reasonable method for reducing random noise. However, we derived that the ( $1+1$ )-EA with single-evaluation (i.e., the ( $1+1$ )-EA without any noise handling method) has the PNT $\left[0,1-\frac{1}{\Theta(\text { poly(n) })}\right]$ from Theorem 5 . while the $(1+1)$-EA with re-evaluation has the PNT $\left[0, \Theta\left(\frac{\log n}{n}\right)\right]$. It is surprising to see that the re-evaluation strategy leads to a much worse noise tolerance than that without any noise handling method.

The re-evaluation with threshold selection strategy has a better PNT comparing with the re-evaluation alone. When the threshold is 1 , we derived the PNT $[0,1]$ from Theorem 8 . and when the threshold is 2 , we obtained $\left[\frac{1}{\Theta(p o l y(n))}, 1-\frac{1}{\Theta(p o l y(n))}\right]$ from Theorem 9 , The improvement from re-evaluation alone could be explained by the fact that the threshold selection filters out false progresses that are caused by the noise. Furthermore, it shows an improvement from the ( $1+1$ )-EA without any noise handling method when selecting the proper threshold $\tau=1$.

We have also studied the single-evaluation with threshold selection. The PNT is $[0,0]$, which implies that threshold selection alone cannot help single-evaluation.

Finally, we analyzed the usefulness of these noise handling strategies under a variant of one-bit noise. All of them are shown to be ineffective when the noise probability reaches the maximum value 1 . We then proposed the smooth threshold selection, which allows a fractional threshold to be effective. We proved that the ( $1+1$ )-EA with (1) $+\left(1-\frac{|x|_{0}}{2 e n}\right.$ or $\left.1-\frac{1}{2 e n}\right)$-smooth threshold selection has the PNT $[0,1]$ from Theorems 12 and 13, and found that the fractional threshold is essential to the proof. Our explanation is that, like the original threshold selection, the proposed one filters out false progresses, while also retaining some chances of accepting true progresses. We further carry out experiments to verify whether the smooth threshold could be helpful in practical problems. The experiments on the maximum matching problem show that the best performance can be achieved at fractional thresholds.

For analyzing the usefulness of noise handling strategies, we have studied a simplified noise model called one-bit noise. A direct generalization that will be studied in the future is to analyze the bit-wise noise, which flips each bit independently with some probability. The bit-wise noise can change the solution greatly in evaluation and thus may make the analysis much more difficult. We shall also improve some currently derived running time bounds, for example, the current running time upper and lower bounds of the single-evaluation still have a gap of $n$. To theoretically analyze the relationship between the effect of noise and the hardness of optimization
problems is also an interesting future work.

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[^0]:    * Corresponding author

