



# Lecture 12. Stochastic Bandits

Advanced Optimization (Fall 2022)

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# Outline

- Multi-Armed Bandits
  - Explore-Then-Exploit
  - Upper Confidence Bound
- Linear Bandits
  - LinUCB Algorithm
  - Generalized Linear Bandits
- Advanced Topics

# Stochastic Multi-Armed Bandit (MAB)

- MAB: A player is facing  $K$  arms. At each time  $t$ , the player pulls one arm  $a \in [K]$  and then receives a reward  $r_t(a) \in [0, 1]$ :

Arm 1	$r_1(1)$	$r_2(1)$	0.6	$r_4(1)$	$r_5(1)$
Arm 2	1	$r_2(2)$	$r_3(2)$	0.2	$r_5(2)$
Arm 3	$r_1(3)$	0.7	$r_3(3)$	$r_4(3)$	0.3

- Stochastic:

Each arm  $a \in [K]$  has an unknown distribution  $\mathcal{D}_a$  with mean  $\mu(a)$ , such that rewards  $r_1(a), r_2(a), \dots, r_T(a)$  are i.i.d samples from  $\mathcal{D}_a$ .

# Stochastic MAB: Formulation

At each round  $t = 1, 2, \dots$

- (1) the player first chooses an arm  $a_t \in [K]$ ;
- (2) and then environment reveals a reward  $r_t(a_t) \in [0, 1]$ ;
- (3) the player updates the model by the pair  $(a_t, r_t(a_t))$ .

- The goal is to minimize the *(pseudo)-regret*:

$$\mathbb{E}[\text{Regret}_T] = \max_{a \in [K]} \mathbb{E} \left[ \sum_{t=1}^T r_t(a) - \sum_{t=1}^T r_t(a_t) \right] = T\mu(a^*) - \sum_{t=1}^T \mu(a_t)$$

where  $a^* = \arg \max_{a \in [K]} \mu(a)$  is the best arm in the sense of expectation.

# Deploying Exp3 to Stochastic MAB

- Stochastic MAB is a special case of Adversarial MAB

⇒ Directly deploying Exp3 for stochastic MAB achieves

**Theorem 1.** Suppose that  $\forall t \in [T]$  and  $i \in [K], 0 \leq \ell_t(i) \leq 1$ , then Exp3 with learning rate  $\eta = \sqrt{(\ln K)/(TK)}$  guarantees

$$\mathbb{E}[\text{Regret}_T] = \mathbb{E} \left[ \sum_{t=1}^T \ell_t(a_t) \right] - \min_{a \in [K]} \sum_{t=1}^T \ell_t(a) \leq \mathcal{O} \left( \sqrt{TK \ln K} \right),$$

where the expectation is taken on the randomness of the algorithm.

⇒ Not yet exploit benign *stochastic* assumption.... *instance-dependent analysis*

# Regret Decomposition

- For stochastic MAB, a natural characterization of the arms:

(i) Suboptimality gap:  $\Delta_a = \mu(a^*) - \mu(a)$ ;

(ii) Number of times arm  $a$  is pulled in  $t$  rounds:  $n_t(a) = \sum_{\tau=1}^t \mathbf{1}\{a_\tau = a\}$ .

- Regret can be reformulated as

$$\begin{aligned}\mathbb{E}[\text{Regret}_T] &= \max_{a \in [K]} \mathbb{E} \left[ \sum_{t=1}^T r_t(a) - \sum_{t=1}^T r_t(a_t) \right] = T\mu(a^*) - \sum_{t=1}^T \mu(a_t) \\ &= \sum_{a \in [K]} (\mu(a^*) - \mu(a)) n_T(a) = \sum_{a \in [K]} \Delta_a n_T(a)\end{aligned}$$

# A Natural Solution

- **Explore-then-Exploit (ETE):**

- (1) Do explore for the first  $T_0$  round by pulling each arm for  $T_0/K$  times;
- (2) Do exploit for the rest  $T - T_0$  round by always pulling  $\hat{a} = \arg \max_{a \in [K]} \hat{\mu}_{T_0}(a)$ .

**Theorem 1.** Suppose that  $\forall t \in [T]$  and  $a \in [K], 0 \leq r_t(a) \leq 1$ , then ETE with exploration period  $T_0$  guarantees

$$\mathbb{E}[\text{Regret}_T] \leq \sum_{a \in [K]} \left( \frac{T_0}{K} + 2T \exp \left( -\frac{2T_0 \Delta_a^2}{K} \right) \right) \Delta_a.$$

# Proof of ETE Regret Bound

**Proof.**  $\mathbb{E}[\text{Regret}_T] = \sum_{a \in [K]} \Delta_a n_T(a)$

Exploration      Exploitation

$$n_T(a) = T_0/K + (T - T_0) \Pr \{ \hat{a} = a \}$$

$$\leq T_0/K + (T - T_0) \Pr \{ \hat{\mu}_{T_0}(a) \geq \hat{\mu}_{T_0}(a^*) \}$$

$$\text{or } \hat{\mu}_{T_0}(a) \leq \frac{\mu(a) + \mu(a^*)}{2} \leq \hat{\mu}_{T_0}(a^*)$$

$$\leq T_0/K + (T - T_0) \Pr \left\{ \hat{\mu}_{T_0}(a) \geq \frac{\mu(a) + \mu(a^*)}{2} \cup \hat{\mu}_{T_0}(a^*) \leq \frac{\mu(a) + \mu(a^*)}{2} \right\}$$

$$\leq T_0/K + (T - T_0) \left( \Pr \left\{ \hat{\mu}_{T_0}(a) \geq \frac{\mu(a) + \mu(a^*)}{2} \right\} + \Pr \left\{ \hat{\mu}_{T_0}(a^*) \leq \frac{\mu(a) + \mu(a^*)}{2} \right\} \right)$$

Union bound  $\Pr\{X \cup Y\} \leq \Pr\{X\} + \Pr\{Y\}$



# Proof of ETE Regret Bound

**Proof.**  $n_T(a) \leq T_0/K + (T - T_0) \left( \Pr \left\{ \hat{\mu}_{T_0}(a) \geq \frac{\mu(a) + \mu(a^*)}{2} \right\} + \Pr \left\{ \hat{\mu}_{T_0}(a^*) \leq \frac{\mu(a) + \mu(a^*)}{2} \right\} \right)$

**Hoeffding's inequality.** for  $X_i \in [0, 1], i \in [m], \bar{X} = \frac{1}{m} \sum_{i=1}^m X_i$ , we have

$$\Pr \{ \bar{X} - \mathbb{E}[\bar{X}] \geq \epsilon \} \leq \exp(-2m\epsilon^2);$$
$$\Pr \{ \bar{X} - \mathbb{E}[\bar{X}] \leq -\epsilon \} \leq \exp(-2m\epsilon^2).$$

$$\Rightarrow \Pr \left\{ \hat{\mu}_{T_0}(a) \geq \frac{\mu(a) + \mu(a^*)}{2} \right\} = \Pr \{ \hat{\mu}_{T_0}(a) \geq \mu(a) + \Delta_a \} \leq \exp \left( -\frac{2T_0\Delta_a^2}{K} \right) \quad \Delta_a = \mu(a^*) - \mu(a)$$

$$\Rightarrow \Pr \left\{ \hat{\mu}_{T_0}(a^*) \leq \frac{\mu(a) + \mu(a^*)}{2} \right\} = \Pr \{ \hat{\mu}_{T_0}(a^*) \leq \mu(a^*) + \Delta_a \} \leq \exp \left( -\frac{2T_0\Delta_a^2}{K} \right)$$

$$\Rightarrow \text{Regret}_T = \sum_{a \in [K]} \Delta_a n_T(a) \leq \sum_{a \in [K]} \left( \frac{T_0}{K} + 2T \exp \left( -\frac{2T_0\Delta_a^2}{K} \right) \right) \Delta_a \quad \square$$

# Issue of ETE

**Theorem 1.** Suppose that  $\forall t \in [T]$  and  $a \in [K], 0 \leq r_t(a) \leq 1$ , then ETE with explore period  $T_0$  guarantees

$$\mathbb{E}[\text{Regret}_T] \leq \sum_{a \in [K]} \left( \frac{T_0}{K} + 2T \exp \left( -\frac{2T_0 \Delta_a^2}{K} \right) \right) \Delta_a.$$

- Need to tune  $T_0$

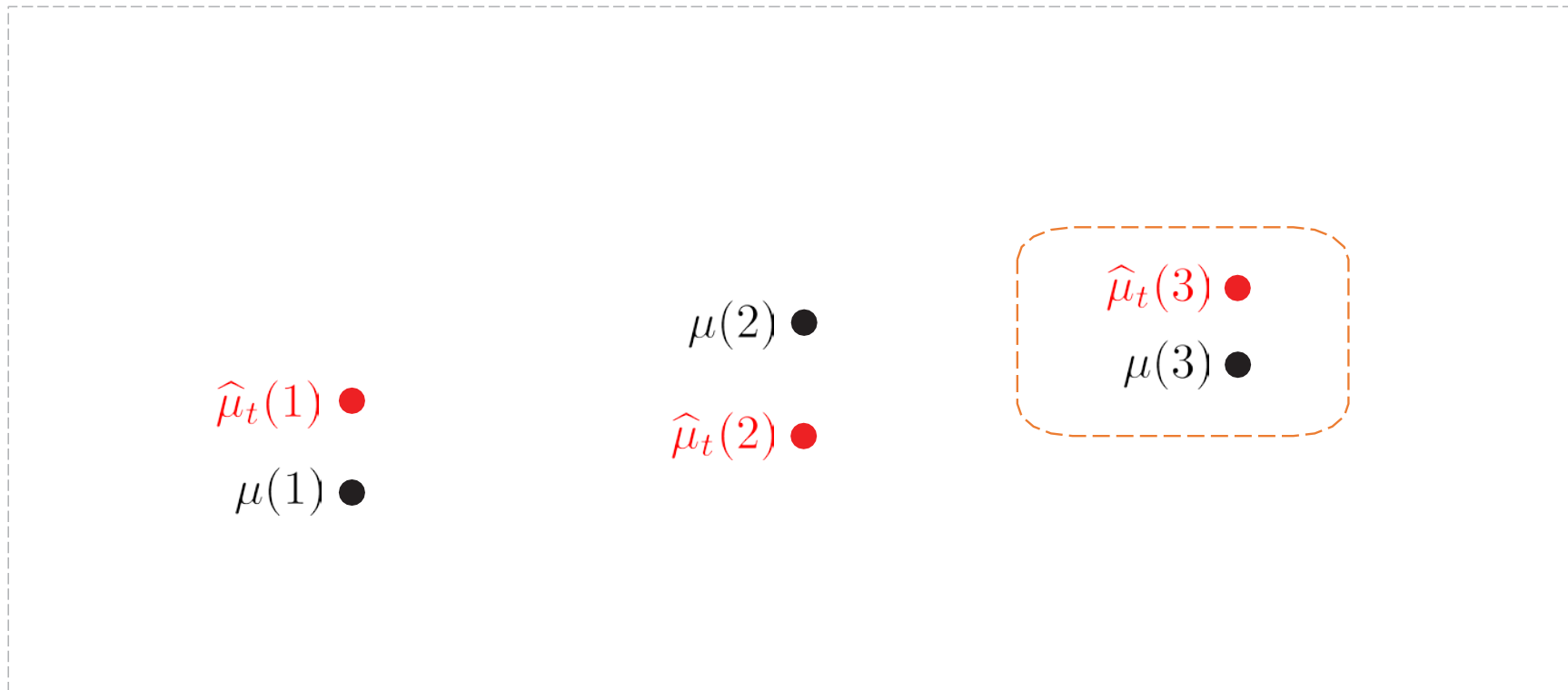
Tune  $T_0$  with prior of suboptimality gap  $\Delta_a$ :  $\mathbb{E}[\text{Regret}_T] = \tilde{\mathcal{O}}(\sqrt{T})$

Tune  $T_0$  without prior of suboptimality gap  $\Delta_a$ :  $\mathbb{E}[\text{Regret}_T] = \tilde{\mathcal{O}}(T^{2/3})$

$\Rightarrow$  Solution: do explore and exploit adaptively.

# Upper Confidence Bound

- ETE

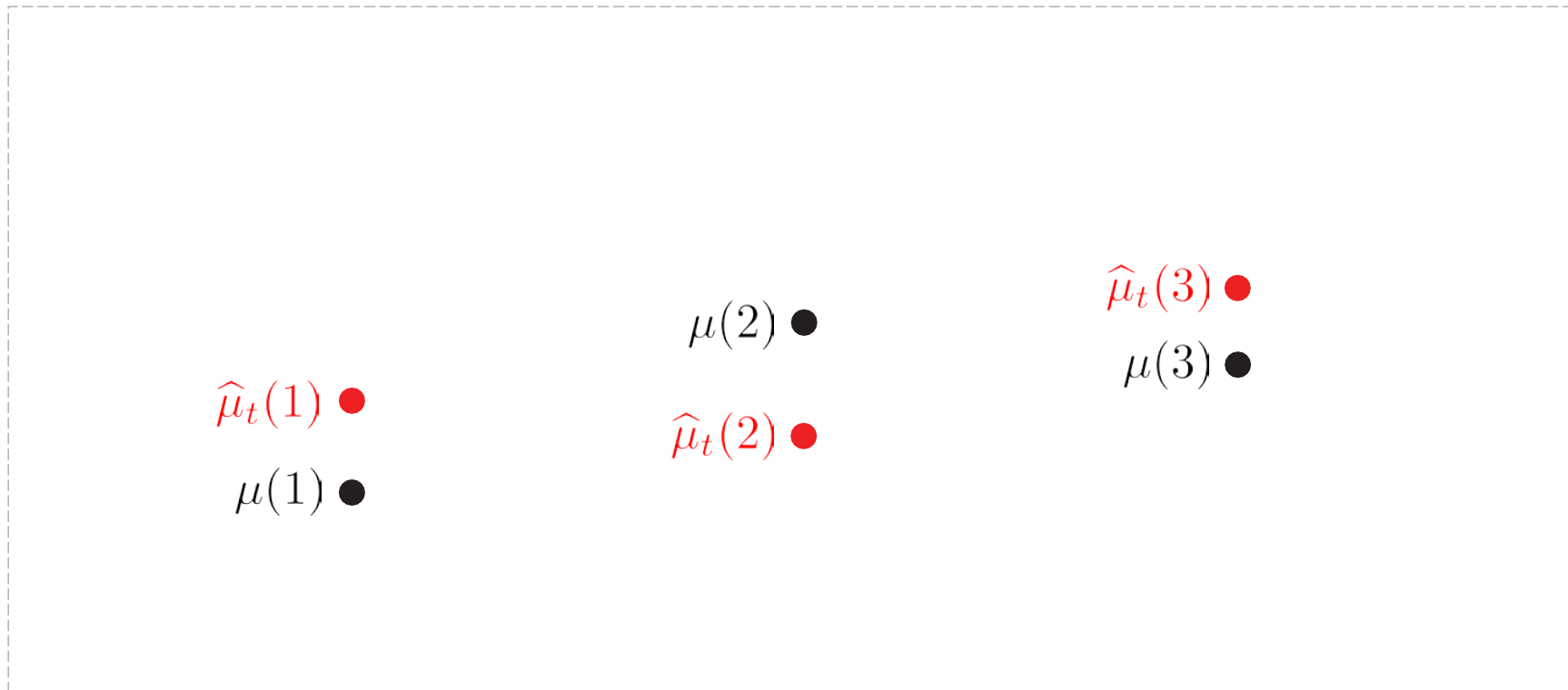


Relying on the estimate of the previous  $T_0$  rounds.

**There is no way to revise the estimate!**

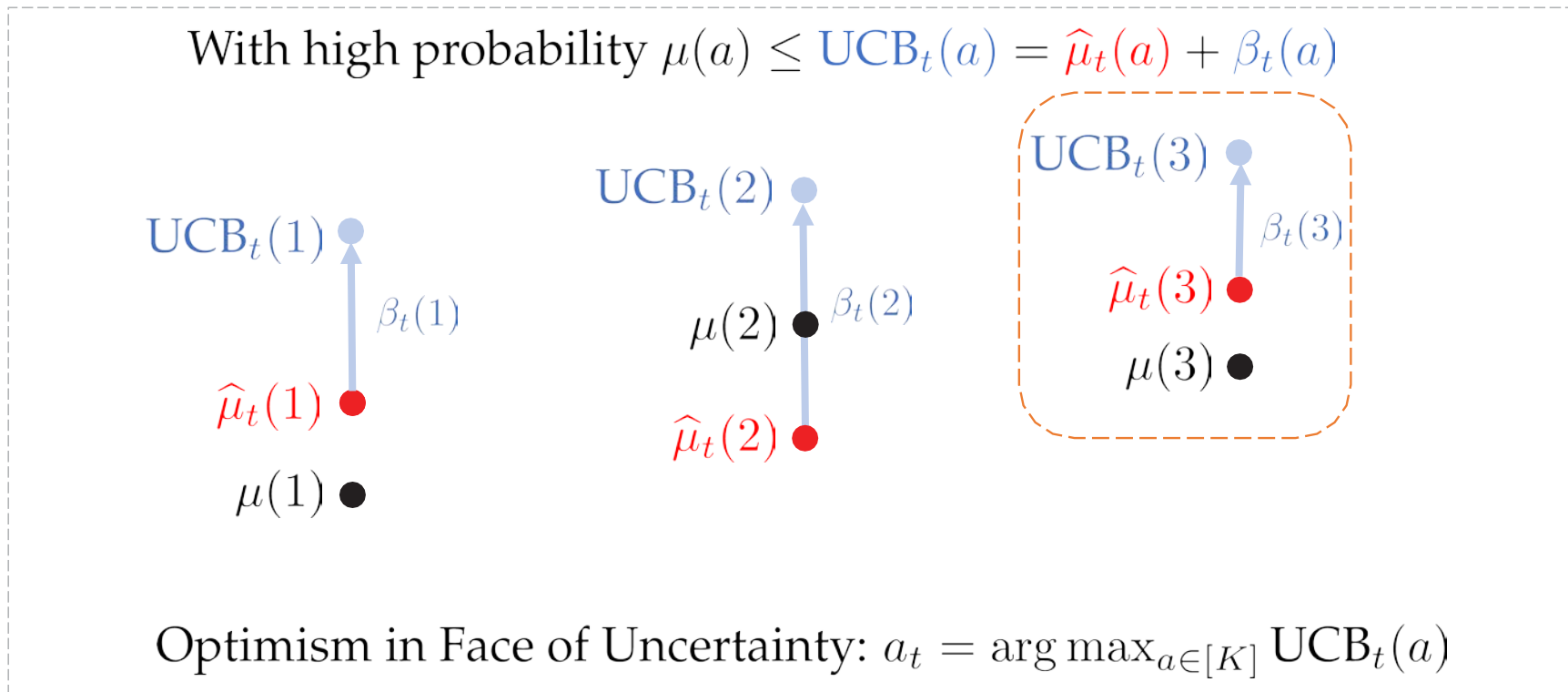
# Upper Confidence Bound

- UCB



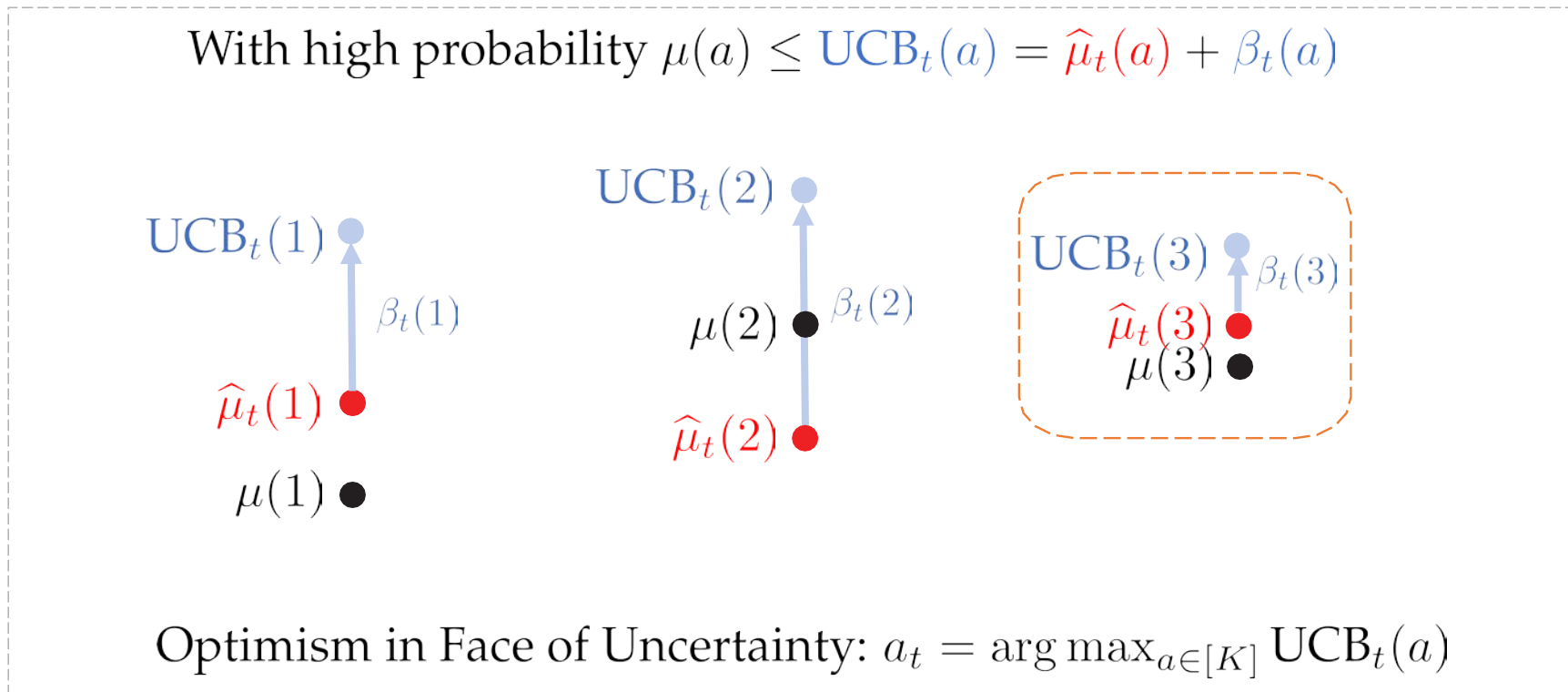
# Upper Confidence Bound

- UCB



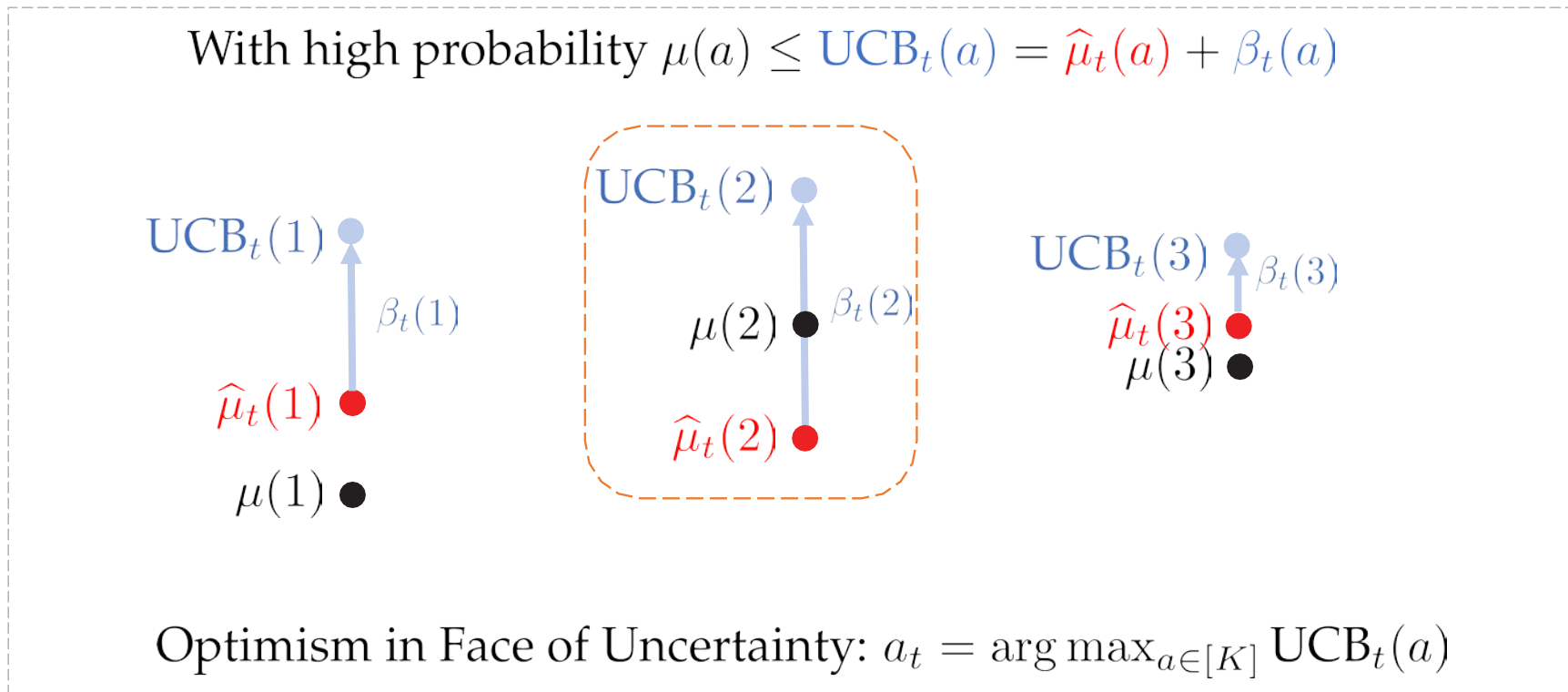
# Upper Confidence Bound

- UCB



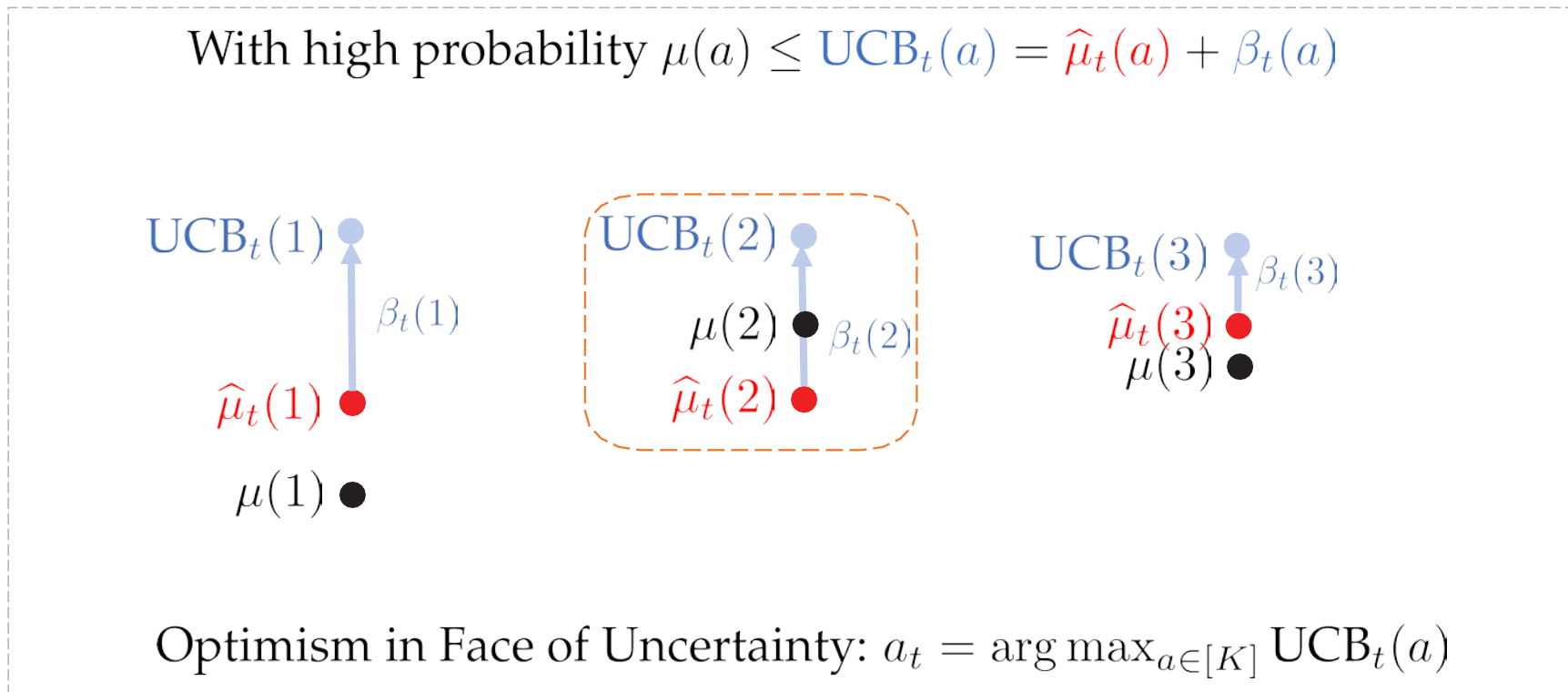
# Upper Confidence Bound

- UCB



# Upper Confidence Bound

- UCB

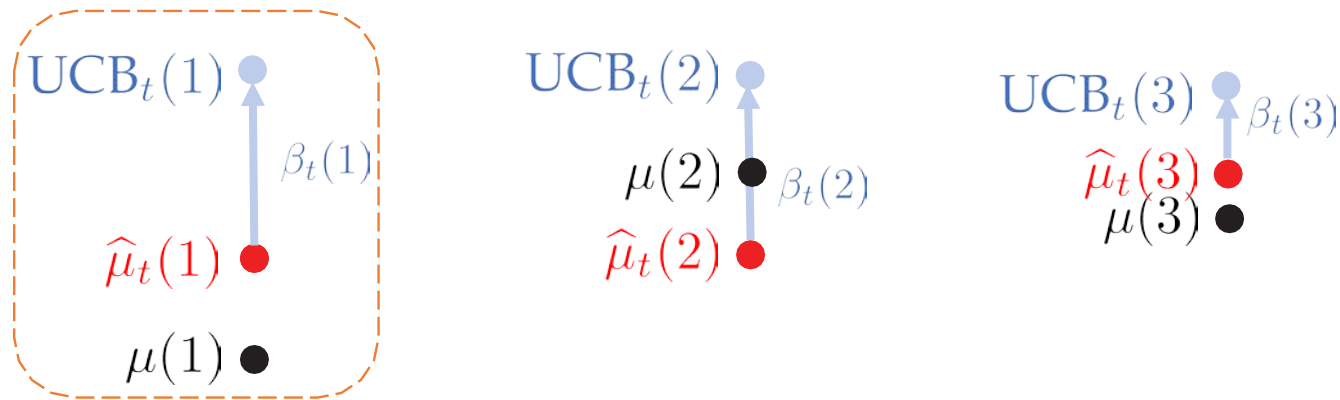




# Upper Confidence Bound

- UCB

With high probability  $\mu(a) \leq \text{UCB}_t(a) = \hat{\mu}_t(a) + \beta_t(a)$

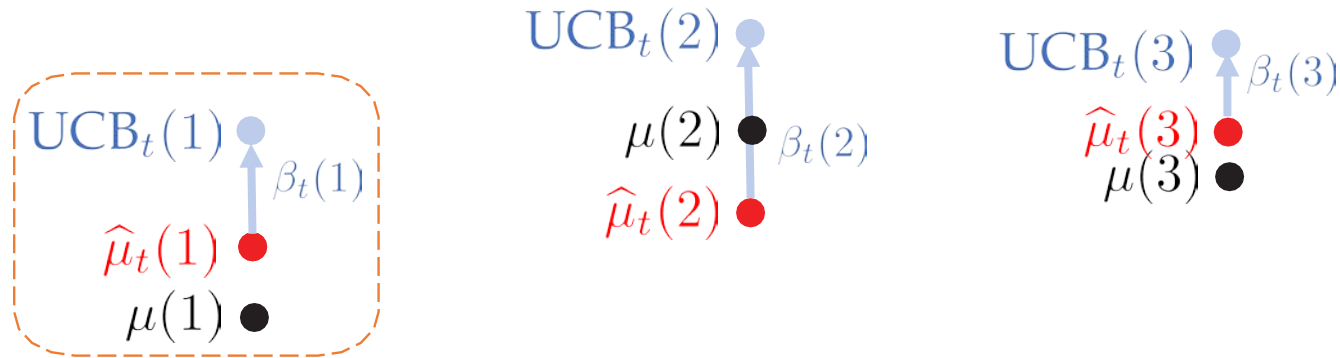


Optimism in Face of Uncertainty:  $a_t = \arg \max_{a \in [K]} \text{UCB}_t(a)$

# Upper Confidence Bound

- UCB

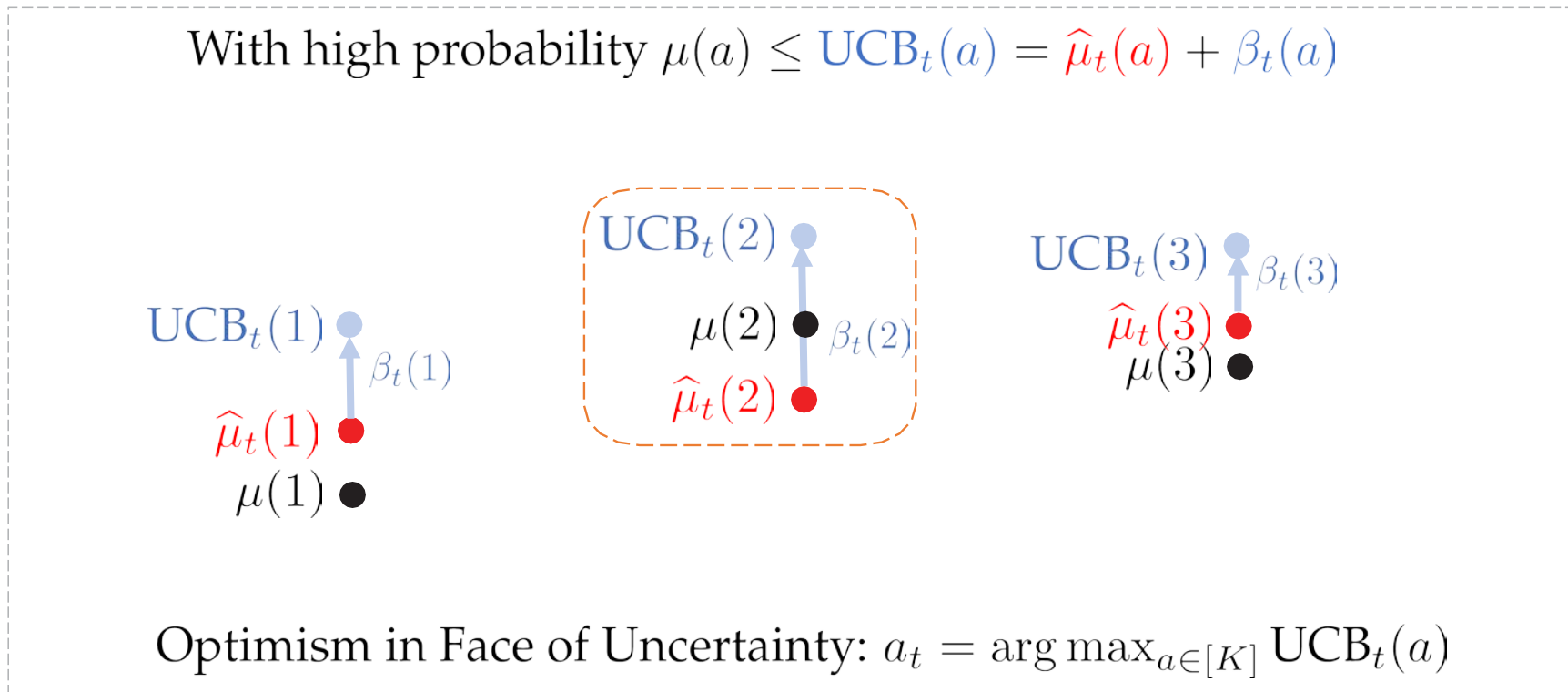
With high probability  $\mu(a) \leq \text{UCB}_t(a) = \hat{\mu}_t(a) + \beta_t(a)$



Optimism in Face of Uncertainty:  $a_t = \arg \max_{a \in [K]} \text{UCB}_t(a)$

# Upper Confidence Bound

- UCB



A large UCB means **uncertainty** or **good arm**.

Choosing the largest UCB means either **exploring** or **exploiting**.

# UCB Algorithm: Formulation

## UCB Algorithm

At each round  $t = 1, 2, \dots$

- (1) Choose arm  $a_t = \arg \max_{a \in [K]} \text{UCB}_{t-1}(a)$
- (2) Observe reward  $r_t$  and update the estimation  $\hat{\mu}_t$
- (3) Update upper confidence bounds  $\text{UCB}_t$  by new estimation

- Estimation: empirical average

$$\hat{\mu}_t(a) = \frac{1}{n_t(a)} \sum_{\tau=1}^t \mathbf{1}\{a_\tau = a\} r_\tau(a)$$

- UCB construction: Hoeffding's inequality

# Construct UCB

**Lemma 1** (Estimation error). *With probability at least  $1 - 2K/T$ , we have,*

$$\forall a \in [K], t \in [T], |\mu(a) - \hat{\mu}_t(a)| \leq \sqrt{\frac{\ln 1/\delta}{n_t(a)}}.$$

*So we have  $\mu(a) \leq \text{UCB}_t(a) \triangleq \hat{\mu}_t(a) + \sqrt{\frac{\ln T}{n_t(a)}}$*

**Proof.** For each arm  $a$ , by Hoeffding inequality and union bound, we have

$$\Pr \left\{ |\mu(a) - \hat{\mu}_t(a)| \leq \sqrt{\frac{\ln 1/\delta}{2n_t(a)}} \right\} \geq 1 - 2\delta \quad \begin{array}{l} \Pr \{ \bar{X} - \mathbb{E}[\bar{X}] \geq \epsilon \} \leq \exp(-2m\epsilon^2) \\ \Pr \{ \bar{X} - \mathbb{E}[\bar{X}] \leq -\epsilon \} \leq \exp(-2m\epsilon^2) \end{array}$$

Further more, by union bound again and let  $\delta = 1/T^2$ ,

$$\Pr \left\{ \forall a \in [K], t \in [T], |\mu(a) - \hat{\mu}_t(a)| \leq \sqrt{\frac{\ln T}{n_t(a)}} \right\} \geq 1 - 2\frac{K}{T} \quad \square$$

# UCB: Distribution-Dependent Bound

**Theorem 2** (Distribution-dependent). Suppose that  $\forall t \in [T]$  and  $a \in [K]$ ,  $0 \leq r_t(a) \leq 1$ , then with probability at least  $1 - 2K/T$ , UCB satisfies

$$\mathbb{E}[\text{Regret}_T] \leq \sum_{a: \Delta_a > 0} \frac{4 \ln T}{\Delta_a} + \Delta_a.$$

**Proof.**  $\mathbb{E}[\text{Regret}_T] = \sum_{a \in [K]} \Delta_a n_T(a)$

With probability at least  $1 - 2K/T$

$$\begin{aligned} \Delta_{a_t} &= \mu(a^*) - \mu(a_t) \leq \text{UCB}_{t-1}(a^*) - \mu(a_t) \quad \forall a \in [K], \mu(a) \leq \text{UCB}_t(a) \\ &\leq \text{UCB}_{t-1}(a_t) - \mu(a_t) \quad a_t = \arg \max_{a \in [K]} \text{UCB}_{t-1}(a) \\ &\leq 2 \sqrt{\frac{\ln T}{n_{t-1}(a_t)}} \quad \mu(a) \leq \text{UCB}_t(a) \triangleq \hat{\mu}_t(a) + \sqrt{\frac{\ln T}{n_t(a)}} \end{aligned}$$

# Proof of UCB Regret Bound

**Proof.**  $\Delta_{a_t} \leq 2\sqrt{\frac{\ln T}{n_{t-1}(a_t)}}$

Let  $t$  be the last time  $a$  is selected, then with probability at least  $1 - 2K/T$ ,

$$\Delta_a \leq 2\sqrt{\frac{\ln T}{n_{t-1}(a)}} = 2\sqrt{\frac{\ln T}{n_T(a) - 1}}$$

$$\Rightarrow n_T(a) \leq 4\frac{\ln T}{\Delta_a^2} + 1$$

$$\mathbb{E}[\text{Regret}_T] = \sum_{a \in [K]} \Delta_a n_T(a) \leq \sum_{a: \Delta_a > 0} \Delta_a \left( 4\frac{\ln T}{\Delta_a^2} + 1 \right) = \sum_{a: \Delta_a > 0} 4\frac{\ln T}{\Delta_a} + \Delta_a$$

□

# UCB: Distribution-Dependent Bound

**Theorem 2** (Distribution-dependent). *Suppose that for all  $t \in [T]$  and  $a \in [K]$ ,  $0 \leq r_t(a) \leq 1$ , then with probability at least  $1 - 2K/T$ , UCB satisfies*

$$\mathbb{E}[\text{Regret}_T] \leq \sum_{a: \Delta_a > 0} \frac{4 \ln T}{\Delta_a} + \Delta_a.$$

- Smaller the  $\Delta_a$ , larger the regret. Its harder to distinguish the optimal arm from the suboptimal one.
- However, tiny  $\Delta_a$  should not lead to larger regret. Always pick arm  $a$  should just lead to  $\mathbb{E}[\text{Regret}_T] = \Delta_a T$ .

$$\implies \mathbb{E}[\text{Regret}_T] \leq \min \left\{ \max_{a \in [K]} \Delta_a T, \sum_{a: \Delta_a > 0} \frac{4 \ln T}{\Delta_a} + \Delta_a \right\}$$



# UCB: Distribution-Free Bound

**Theorem 3** (Distribution-free). Suppose that for all  $t \in [T]$  and  $a \in [K]$ ,  $0 \leq r_t(a) \leq 1$ , then UCB satisfies

$$\mathbb{E}[\text{Regret}_T] \leq 2\sqrt{TK \ln T} + \sum_{a \in [K]} \Delta_a = \mathcal{O}\left(\sqrt{TK \ln T}\right)$$

**Proof.**

$$\begin{aligned} \mathbb{E}[\text{Regret}_T] &= \sum_{a \in [K]} \Delta_a n_T(a) = \sum_{a: \Delta_a < \Delta} \Delta_a n_T(a) + \sum_{a: \Delta_a \geq \Delta} \Delta_a n_T(a) \\ &\leq T\Delta + \sum_{a: \Delta_a \geq \Delta} \Delta_a \left(4 \frac{\ln T}{\Delta_a^2} + 1\right) \leq T\Delta + 4 \frac{K \ln T}{\Delta} + \sum_{a \in [K]} \Delta_a \quad n_T(a) \leq 4 \frac{\ln T}{\Delta_a^2} + 1 \\ &\leq 2\sqrt{TK \ln T} + \sum_{a \in [K]} \Delta_a \quad \text{Choosing } \Delta = 2\sqrt{K \ln T / T} \quad \square \end{aligned}$$

# Upper Bound and Lower Bound

**Theorem 3** (Distribution-free). *Suppose that for all  $t \in [T]$  and  $a \in [K]$ ,  $0 \leq r_t(a) \leq 1$ , then UCB satisfies*

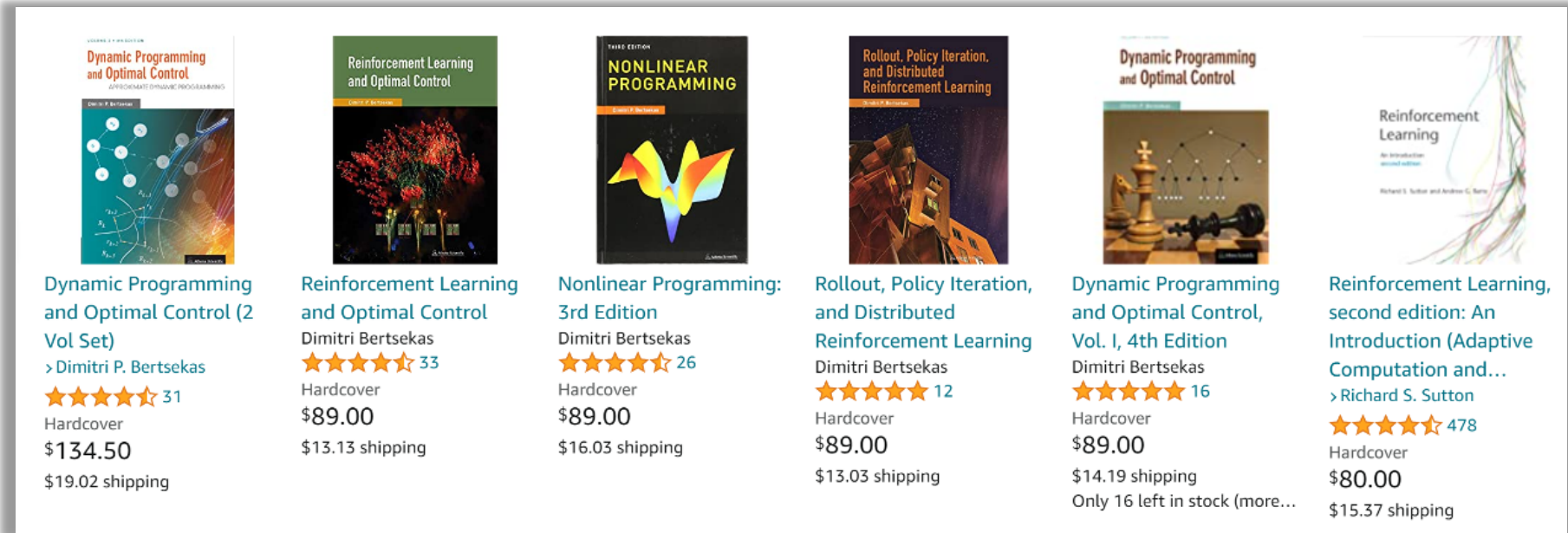
$$\mathbb{E}[\text{Regret}_T] \leq 2\sqrt{TK \ln T} + \sum_{a \in [K]} \Delta_a = \mathcal{O}\left(\sqrt{TK \ln T}\right)$$

**Theorem 2** (Lower Bound for MAB). *For any bandit algorithm  $\mathcal{A}$ , there exists a sequence of loss vectors such that*

$$\inf_{\mathcal{A}} \sup_{\ell_1, \dots, \ell_T} \mathbb{E}[\text{Regret}_T] = \Omega(\sqrt{TK})$$

# Stochastic Linear Bandits

- A ubiquitous problem in real life:



- Each arm represent a book and has side information;
- Arm set could be very large or even infinite.

# Stochastic LB: Formulation

Each arm is represented as a **feature vector**  $\mathbf{x} \in \mathcal{X} \subseteq \mathbb{R}^d$

At each round  $t = 1, 2, \dots$

- (1) the player first chooses an arm  $X_t$  from arm set  $\mathcal{X}$ ;
- (2) and then environment reveals a reward  $r_t \in \mathbb{R}$ .

	Multi-Armed Bandits	Linear Bandits
Arm set	finite arm set $[K]$	infinite arm set $\mathcal{X} = \{\ \mathbf{x}\ _2 \leq L\}$
Model	$\mathbb{E}[r(a)] = \mu(a)$ $\forall t \in [T], a \in [K], r_t(a) \in [0, 1]$	$r_t = X_t^\top \theta_* + \eta_t \quad \mu(\mathbf{x}) = \mathbf{x}^\top \theta_*$ $\eta_t$ : sub-Gaussian noise
Regret	$\mathbb{E}[\text{Regret}_T] = T \max_{a \in [K]} \mu(a) - \sum_{t=1}^T \mu(a_t)$	$\mathbb{E}[\text{Regret}_T] = T \max_{\mathbf{x} \in \mathcal{X}} \mathbf{x}^\top \theta_* - \sum_{t=1}^T X_t^\top \theta_*$

# Deploying UCB to Linear Bandits

- Linear Bandits is a special case of MAB with infinite arm:

⇒ Why not directly deploy UCB to address Linear Bandits?

**Theorem 3** (Distribution-free). *Suppose that  $\forall t \in [T]$  and  $a \in [K]$ ,  $0 \leq r_t(a) \leq 1$ , then UCB satisfies*

$$\mathbb{E}[\text{Regret}_T] \leq 2\sqrt{TK \ln T} + \sum_{a \in [K]} \Delta_a = \mathcal{O}\left(\sqrt{TK \ln T}\right)$$

Infinite arm set ( $K \rightarrow \infty$ ) leads to meaningless regret guarantee!

⇒ Not yet exploit the addition *contextual feature information*...

# LinUCB Algorithm: Formulation

## LinUCB Algorithm

At each round  $t = 1, 2, \dots$

- (1) Select  $X_t = \arg \max_{\mathbf{x} \in \mathcal{X}} \text{UCB}_{t-1}(\mathbf{x})$
- (2) Observe reward  $r_t$  and update the estimation  $\hat{\mu}_t$
- (3) update upper confidence bounds  $\text{UCB}_t$  by new estimation

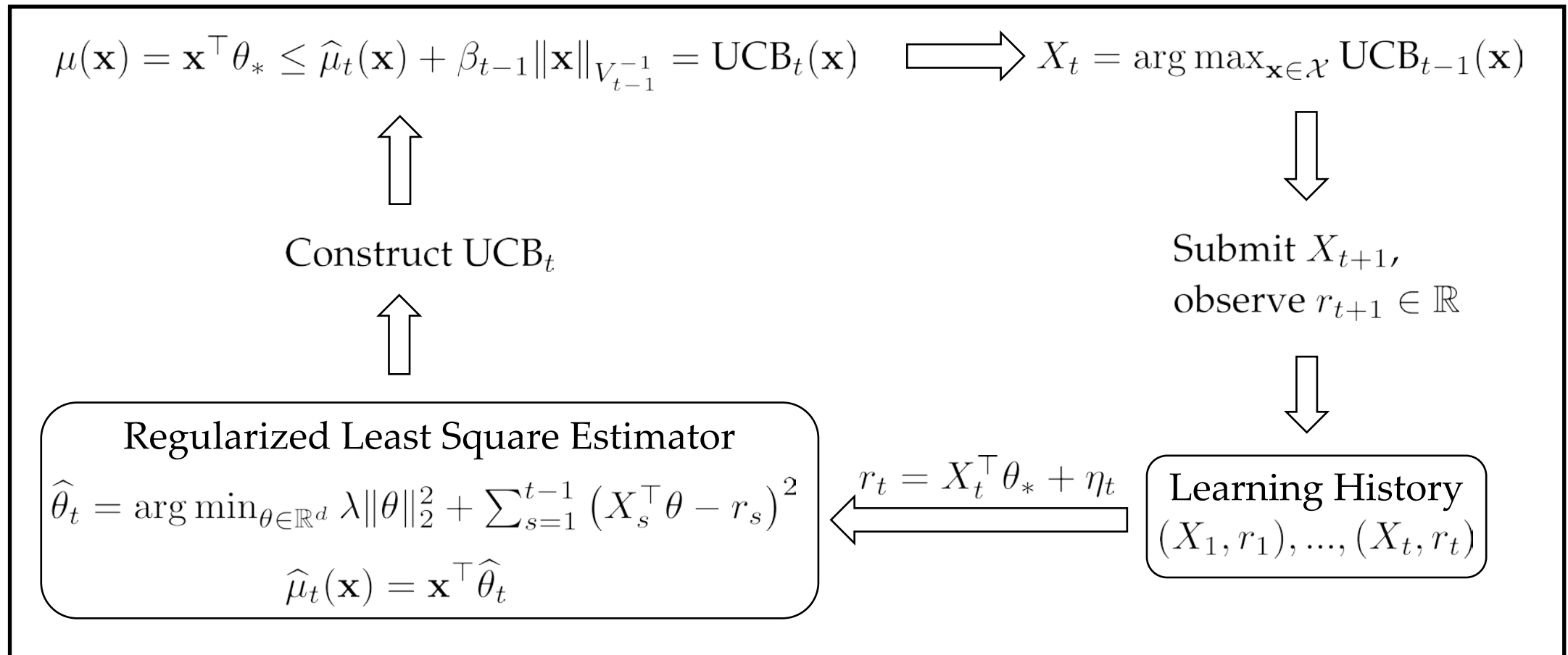
- Estimation: regularized least square (linear regression)

$$\hat{\theta}_t = \arg \min_{\theta \in \mathbb{R}^d} \lambda \|\theta\|_2^2 + \sum_{s=1}^{t-1} (X_s^\top \theta - r_s)^2$$

$$\text{Closed form: } \hat{\theta}_t = V_{t-1}^{-1} \left( \sum_{s=1}^{t-1} r_s X_s \right), V_{t-1} = \lambda I + \sum_{s=1}^{t-1} X_s X_s^\top$$

# LinUCB Algorithm

*Optimism in Face of Uncertainty*



# LinUCB Algorithm

- UCB for stochastic MAB
  - (1) estimate  $\mu(a)$  by average estimation;
  - (2) construct upper confidence bound for  $\mu(a)$  by concentration inequalities.
- UCB for stochastic LB (LinUCB)
  - More information can be used to estimate expected reward.

## UCB estimation

$$\hat{\mu}_t(a) = \frac{1}{n_t(a)} \sum_{\tau=1}^t \mathbf{1}\{a_\tau = a\} r_\tau(a)$$

## LinUCB estimation

$$\hat{\theta}_t = \arg \min_{\theta \in \mathbb{R}^d} \lambda \|\theta\|_2^2 + \sum_{s=1}^{t-1} (X_s^\top \theta - r_s)^2$$
$$\hat{\mu}_t(\mathbf{x}) = \mathbf{x}^\top \hat{\theta}_t$$



# Construct UCB

**Lemma 2** (Estimation error). For any  $\mathbf{x} \in \mathcal{X}$ ,  $\delta \in (0, 1)$ , with probability at least  $1 - \delta$ , the following holds for all  $t \in [T]$

$$\left| \mathbf{x}^\top (\hat{\theta}_t - \theta) \right| \leq \beta_{t-1} \|\mathbf{x}\|_{V_{t-1}^{-1}}, \quad \text{where } \beta_{t-1} = R \sqrt{2 \log \left( \frac{1}{\delta} \right) + d \log \left( 1 + \frac{(t-1)L^2}{\lambda d} \right)} + \sqrt{\lambda} S$$

So we have  $\mu(\mathbf{x}) \leq \text{UCB}_t(\mathbf{x}) \triangleq \hat{\mu}_t(\mathbf{x}) + \beta_{t-1} \|\mathbf{x}\|_{V_{t-1}^{-1}}$

**Proof.**

$$\begin{aligned} \hat{\theta}_t - \theta_* &= V_{t-1}^{-1} \left( \sum_{s=1}^{t-1} r_s X_s \right) - \theta_* & \hat{\theta}_t &= V_{t-1}^{-1} \left( \sum_{s=1}^{t-1} r_s X_s \right) \\ &= V_{t-1}^{-1} \left( \sum_{s=1}^{t-1} (X_s^\top \theta_* + \eta_s) X_s \right) - V_{t-1}^{-1} \left( \lambda I_d + \sum_{s=1}^{t-1} X_s X_s^\top \right) \theta_* \\ &= V_{t-1}^{-1} \left( \sum_{s=1}^{t-1} \eta_s X_s - \lambda \theta_* \right) & V_{t-1} &= \lambda I + \sum_{s=1}^{t-1} X_s X_s^\top \end{aligned}$$

# Proof of Estimation Error Bound

*Proof.*  $\hat{\theta}_t - \theta_* = V_{t-1}^{-1} \left( \sum_{s=1}^{t-1} \eta_s X_s - \lambda \theta_* \right) \quad V_{t-1} = \lambda I + \sum_{s=1}^{t-1} X_s X_s^\top$

$$\left| \mathbf{x}^\top (\hat{\theta}_t - \theta_*) \right| \leq \|\mathbf{x}\|_{V_{t-1}^{-1}} \left\| \hat{\theta}_t - \theta_* \right\|_{V_{t-1}} \quad \text{Cauchy-Schwarz inequality: } |a^\top b| \leq \|a\| \|b\|_*$$

$$\leq \|\mathbf{x}\|_{V_{t-1}^{-1}} \left( \left\| \sum_{s=1}^{t-1} \eta_s X_s \right\|_{V_{t-1}^{-1}} + \|\lambda \theta_*\|_{V_{t-1}^{-1}} \right)$$

**Core difficulty:** The actions  $\{X_s\}_{s=1,\dots,t}$  are neither fixed nor independent but are intricately correlated via the rewards  $\{r_s\}_{s=1,\dots,t}$

# Self-Normalized Concentration

**Theorem 4** (Self-normalized concentration for Vector-Valued Martingales). *Let  $\{F_t\}_{t=0}^\infty$  be a filtration. Let  $\{\eta_t\}_{t=0}^\infty$  be a real-valued stochastic process such that  $\eta_t$  is  $F_t$ -measurable and  $\eta_t$  is conditionally  $R$ -sub-Gaussian for some  $R \geq 0$  i.e.*

$$\forall \lambda \in \mathbb{R}, \mathbb{E} [\exp(\lambda \eta_t) | X_{1:t}, \eta_{1:t-1}] \leq \exp \left( \frac{\lambda^2 R^2}{2} \right).$$

*Let  $\{X_t\}_{t=1}^\infty$  be an  $\mathbb{R}^d$ -valued stochastic process such that  $X_t$  is  $F_{t-1}$ -measurable. Assume that  $V$  is a  $d \times d$  positive definite matrix. For any  $t \geq 0$ , define*

$$V_t = V_0 + \sum_{s=1}^t X_s X_s^\top, \quad S_t = \sum_{s=1}^t \eta_s X_s.$$

*Then, for any  $\delta > 0$ , with probability at least  $1 - \delta$ , for all  $t \geq 0$ ,*

$$\|S_t\|_{V_t^{-1}}^2 \leq 2R^2 \log \left( \frac{\det(V_t)^{\frac{1}{2}} \det(V_0)^{-\frac{1}{2}}}{\delta} \right).$$

# Proof of Self-Normalized Concentration

**Lemma 3** (Method of Mixture).

$$\exp \left( \frac{1}{2R^2} \|S_t\|_{V_t^{-1}}^2 \right) = \frac{\det(V_t)}{\det(V_0)} \int_{\mathbb{R}^d} \exp \left( \frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t - V_0}^2 \right) dh(\lambda), \text{ where } h(\lambda) = \mathcal{N}(0, V_0^{-1})$$

**Proof.**

$$\begin{aligned} & \int_{\mathbb{R}^d} \exp \left( \frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t}^2 \right) d\lambda \\ &= \int_{\mathbb{R}^d} \exp \left( \frac{\langle \lambda_*, S_t \rangle}{R} - \frac{1}{2} \|\lambda_*\|_{V_t}^2 - \frac{1}{2} \|\lambda - \lambda_*\|_{V_t}^2 \right) d\lambda \quad \text{2nd order Taylor expansion} \\ &= \sqrt{\frac{(2\pi)^d}{\det(V_t)}} \exp \left( \frac{\langle \lambda_*, S_t \rangle}{R} - \frac{1}{2} \|\lambda_*\|_{V_t}^2 \right) \int_{\mathbb{R}^d} \sqrt{\frac{\det(V_t)}{(2\pi)^d}} \exp \left( -\frac{1}{2} \|\lambda - \lambda_*\|_{V_t}^2 \right) d\lambda \\ &= \sqrt{(2\pi)^d / \det(V_t)} \exp \left( \frac{\langle \lambda_*, S_t \rangle}{R} - \frac{1}{2} \|\lambda_*\|_{V_t}^2 \right) \quad \mathcal{N}(\lambda_*, V_t^{-1}) \\ &= \sqrt{(2\pi)^d / \det(V_t)} \exp \left( \frac{1}{2R^2} \|S_t\|_{V_t^{-1}}^2 \right) \end{aligned}$$

**Definition of conjugate function**

$$\frac{1}{2R^2} \|S_t\|_{V_t^{-1}}^2 = \frac{\langle \lambda_*, S_t \rangle}{R} - \frac{1}{2} \|\lambda_*\|_{V_t}^2$$

# Proof of Self-Normalized Concentration

**Lemma 3** (Method of Mixture).

$$\exp\left(\frac{1}{2R^2} \|S_t\|_{V_t^{-1}}^2\right) = \frac{\det(V_t)}{\det(V_0)} \int_{\mathbb{R}^d} \exp\left(\frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t - V_0}^2\right) dh(\lambda), \text{ where } h(\lambda) = \mathcal{N}(0, V_0^{-1})$$

**Proof.**  $\int_{\mathbb{R}^d} \exp\left(\frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t}^2\right) d\lambda = \sqrt{(2\pi)^d / \det(V_t)} \exp\left(\frac{1}{2R^2} \|S_t\|_{V_t^{-1}}^2\right)$

$$\begin{aligned} \exp\left(\frac{1}{2R^2} \|S_t\|_{V_t^{-1}}^2\right) &= \sqrt{\frac{\det(V_t)}{(2\pi)^d}} \int_{\mathbb{R}^d} \exp\left(\frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t - V_0}^2\right) \exp\left(-\frac{1}{2} \|\lambda\|_{V_0}^2\right) d\lambda \\ &= \sqrt{\frac{\det(V_t)}{(2\pi)^d}} \sqrt{\frac{(2\pi)^d}{\det(V_0)}} \int_{\mathbb{R}^d} \exp\left(\frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t - V_0}^2\right) \underbrace{\sqrt{\frac{\det(V_0)}{(2\pi)^d}} \exp\left(-\frac{1}{2} \|\lambda\|_{V_0}^2\right)}_{h(\lambda)} d\lambda \\ &= \sqrt{\frac{\det(V_t)}{\det(V_0)}} \int_{\mathbb{R}^d} \exp\left(\frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t - V_0}^2\right) dh(\lambda) \end{aligned}$$

□

# Proof of Self-Normalized Concentration

**Lemma 3** (Method of Mixture).

$$\exp \left( \frac{1}{2R^2} \|S_t\|_{V_t^{-1}}^2 \right) = \frac{\det(V_t)}{\det(V_0)} \int_{\mathbb{R}^d} \exp \left( \frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t - V_0}^2 \right) dh(\lambda),$$

where  $h(\lambda) = \mathcal{N}(0, V_0^{-1})$

**Lemma 4.** Let  $\lambda \in \mathbb{R}^d$  be arbitrary and consider for any  $t > 0$

$$M_t^\lambda = \exp \left( \frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t - V_0}^2 \right)$$

Let  $\tau$  be a stopping time with respect to the filtration  $\{F_t\}_{t=0}^\infty$ . Then  $M_\tau^\lambda$  is almost surely well-defined and

$$\mathbb{E} [M_\tau^\lambda] \leq 1$$

# Proof of Self-Normalized Concentration

## *Proof of Lemma 4.*

$$\begin{aligned}
 \mathbb{E} [M_t^\lambda \mid F_{t-1}] &= \mathbb{E} \left[ \exp \left( \frac{\langle \lambda, S_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_t - V_0}^2 \right) \mid F_{t-1} \right] \\
 &= \mathbb{E} \left[ \exp \left( \frac{\langle \lambda, S_{t-1} \rangle}{R} + \frac{\langle \lambda, \eta_t X_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{V_{t-1} - V_0}^2 - \frac{1}{2} \|\lambda\|_{X_t X_t^\top}^2 \right) \mid F_{t-1} \right] \\
 &= M_{t-1}^\lambda \mathbb{E} \left[ \exp \left( \frac{\langle \lambda, \eta_t X_t \rangle}{R} - \frac{1}{2} \|\lambda\|_{X_t X_t^\top}^2 \right) \mid F_{t-1} \right] && M_{t-1}^\lambda \text{ is } F_{t-1}\text{-measurable} \\
 &= M_{t-1}^\lambda \mathbb{E} \left[ \exp \left( \frac{\langle \lambda, X_t \rangle}{R} \eta_t \right) \mid F_{t-1} \right] \exp \left( -\frac{1}{2} \|\lambda\|_{X_t X_t^\top}^2 \right) && X_t \text{ is } F_{t-1}\text{-measurable} \\
 &\leq M_{t-1}^\lambda \exp \left( \frac{\langle \lambda, X_t \rangle^2}{2} \right) \exp \left( -\frac{1}{2} \|\lambda\|_{X_t X_t^\top}^2 \right) && \eta_t \text{ is } R\text{-sub-Gaussian} \\
 &= M_{t-1}^\lambda.
 \end{aligned}$$

$$\mathbb{E} [M_t^\lambda \mid F_{t-1}] \leq M_{t-1}^\lambda \implies \{M_t^\lambda\}_{t=0}^\infty \text{ is a supermartingale} \implies \mathbb{E} [M_t^\lambda] \leq \mathbb{E} [M_0^\lambda] = 1 \quad \square$$

# Proof of Self-Normalized Concentration

*Proof.*

$$\text{Lemma 4. } \mathbb{E} [M_\tau^\lambda] \leq 1$$

$$\Rightarrow \mathbb{E} [M_\tau^\lambda] \leq 1$$

$$\int_{\mathbb{R}^d} \mathbb{E} [M_\tau^\lambda] dh(\lambda) \leq \int_{\mathbb{R}^d} 1 dh(\lambda) \quad \text{Jensen's inequality}$$

$$\mathbb{E} \left[ \int_{\mathbb{R}^d} M_\tau^\lambda dh(\lambda) \right] \leq 1$$

**Lemma 3** (Method of Mixture).

$$\exp \left( \frac{1}{2R^2} \|S_t\|_{V_t^{-1}}^2 \right) = \frac{\det(V_t)}{\det(V_0)} \int_{\mathbb{R}^d} M_t^\lambda dh(\lambda), \text{ where } h(\lambda) = \mathcal{N}(0, V_0^{-1})$$

$$\Rightarrow \mathbb{E} \left[ \frac{\det(V_0)}{\det(V_t)} \exp \left( \frac{1}{2R^2} \|S_\tau\|_{V_\tau^{-1}}^2 \right) \right] \leq 1$$



# Proof of Self-Normalized Concentration

*Proof.*

$$\tau = \min \left\{ t \geq 0, \|S_t\|_{V_t^{-1}}^2 \leq 2R^2 \log \left( \frac{\det(V_t)^{\frac{1}{2}} \det(V_0)^{-\frac{1}{2}}}{\delta} \right) \right\}$$

$$\begin{aligned} \mathbb{P} \left( \exists t \geq 0, \frac{\det(V_0)}{\det(V_t)} \exp \left( \frac{1}{2R^2} \|S_t\|_{V_t^{-1}}^2 \right) \geq \frac{1}{\delta} \right) &= \mathbb{P} \left( \tau < \infty, \frac{\det(V_0)}{\det(V_\tau)} \exp \left( \frac{1}{2R^2} \|S_\tau\|_{V_\tau^{-1}}^2 \right) \geq \frac{1}{\delta} \right) \\ &\leq \mathbb{P} \left( \frac{\det(V_0)}{\det(V_\tau)} \exp \left( \frac{1}{2R^2} \|S_\tau\|_{V_\tau^{-1}}^2 \right) \geq \frac{1}{\delta} \right) \\ &\leq \mathbb{E} \left[ \frac{\det(V_0)}{\det(V_\tau)} \exp \left( \frac{1}{2R^2} \|S_\tau\|_{V_\tau^{-1}}^2 \right) \right] \delta \\ &\leq \delta. \end{aligned}$$

$$\Rightarrow \mathbb{P} \left( \forall t \geq 0, \|S_t\|_{V_t^{-1}}^2 \leq 2R^2 \log \left( \frac{\det(V_t)^{\frac{1}{2}} \det(V_0)^{-\frac{1}{2}}}{\delta} \right) \right) \geq 1 - \delta.$$

□

# Proof of Estimation Error Bound

**Proof.**  $\left| \mathbf{x}^\top (\hat{\theta}_t - \theta_*) \right| \leq \|\mathbf{x}\|_{V_{t-1}^{-1}} \left( \left\| \sum_{s=1}^{t-1} \eta_s X_s \right\|_{V_{t-1}^{-1}} + \|\lambda \theta_*\|_{V_{t-1}^{-1}} \right)$

**Theorem 4** (Self-normalized concentration). *For any  $\delta \in (0, 1)$ , with probability at least  $1 - \delta$ , for all  $t \geq 0$ ,*

$$\|S_t\|_{V_t^{-1}}^2 \leq 2R^2 \log \left( \frac{\det(V_t)^{\frac{1}{2}} \det(V_0)^{-\frac{1}{2}}}{\delta} \right).$$

$$\text{Tr}(V_t) = \text{Tr}(\lambda I) + \text{Tr} \left( \sum_{s=1}^t X_s X_s^\top \right) \leq \lambda d + tL^2 \qquad V_t = \lambda I + \sum_{s=1}^t X_s X_s^\top$$

$$\det(V_t) = \prod_{i=1}^d \lambda_i \leq \left( \frac{\sum_{i=1}^d \lambda_i}{d} \right)^d = \left( \frac{\text{Tr}(V_t)}{d} \right)^d \leq \left( \frac{\lambda d + tL^2}{d} \right)^d$$

$$\det(V_0) = \det(\lambda I) = \lambda^d \qquad V_0 = \lambda I$$

# Proof of Estimation Error Bound

**Proof.**  $\left| \mathbf{x}^\top (\hat{\theta}_t - \theta_*) \right| \leq \|\mathbf{x}\|_{V_{t-1}^{-1}} \left( \left\| \sum_{s=1}^{t-1} \eta_s X_s \right\|_{V_{t-1}^{-1}} + \|\lambda \theta_*\|_{V_{t-1}^{-1}} \right)$

$$\begin{aligned} \left\| \sum_{s=1}^{t-1} \eta_s X_s \right\|_{V_{t-1}^{-1}} &\leq \sqrt{2R^2 \log \left( \frac{\det(V_t)^{\frac{1}{2}} \det(V_0)^{-\frac{1}{2}}}{\delta} \right)} \leq \sqrt{2R^2 \log \left( \frac{1}{\delta} \left( \frac{\lambda d + (t-1)L^2}{\lambda d} \right)^{\frac{d}{2}} \right)} \\ &= R \sqrt{2 \log \left( \frac{1}{\delta} \right) + d \log \left( 1 + \frac{tL^2}{\lambda d} \right)} \quad \det(V_t) \leq \left( \frac{\lambda d + tL^2}{d} \right)^d \\ &\quad \det(V_0) = \lambda^d \\ \|\lambda \theta_*\|_{V_{t-1}^{-1}} &\leq \frac{1}{\sqrt{\lambda_{\min}(V_{t-1})}} \|\lambda \theta_*\|_2 \leq \frac{1}{\sqrt{\lambda}} \|\lambda \theta_*\|_2 \leq \sqrt{\lambda} S \\ \left| \mathbf{x}^\top (\hat{\theta}_t - \theta_*) \right| &\leq \|\mathbf{x}\|_{V_{t-1}^{-1}} \left( R \sqrt{2 \log \left( \frac{1}{\delta} \right) + d \log \left( 1 + \frac{tL^2}{\lambda d} \right)} + \sqrt{\lambda} S \right) \end{aligned}$$

□

# LinUCB: Regret Bound

**Theorem 5.** Let  $\lambda = d$ , the regret of LinUCB is bounded with probability at least  $1 - 1/T$ , by

$$\mathbb{E}[\text{Regret}_T] \leq \tilde{O}\left(d\sqrt{T}\right)$$

**Proof.** Let  $X_* \triangleq \arg \max_{\mathbf{x} \in \mathcal{X}} \mathbf{x}^\top \theta_*$ , each of the following holds with probability at least  $1 - \delta$ ,

$$\forall t \in [T], X_*^\top \theta_* \leq X_*^\top \hat{\theta}_t + \beta_{t-1} \|X_*\|_{V_{t-1}^{-1}}$$

$$\forall t \in [T], X_t^\top \theta_* \geq X_t^\top \hat{\theta}_t - \beta_{t-1} \|X_t\|_{V_{t-1}^{-1}}$$

With probability at least  $1 - 2\delta$ ,

$$\begin{aligned} \forall t \in [T], X_*^\top \theta_* - X_t^\top \theta_* &\leq X_*^\top \hat{\theta}_t - X_t^\top \hat{\theta}_t + \beta_{t-1} \left( \|X_*\|_{V_{t-1}^{-1}} + \|X_t\|_{V_{t-1}^{-1}} \right) \\ &\leq 2\beta_{t-1} \|X_t\|_{V_{t-1}^{-1}}, \quad X_*^\top \hat{\theta}_t + \beta_{t-1} \|X_*\|_{V_{t-1}^{-1}} \leq X_t^\top \hat{\theta}_t + \beta_{t-1} \|X_t\|_{V_{t-1}^{-1}} \end{aligned}$$

# LinUCB: Regret Bound

**Proof.** With probability at least  $1 - 2\delta$ ,  $\forall t \in [T]$ ,  $X_*^\top \theta_* - X_t^\top \theta_* \leq 2\beta_{t-1} \|X_t\|_{V_{t-1}^{-1}}$ .

$$\mathbb{E}[\text{Regret}_T] = \sum_{t=1}^T (X_*^\top \theta_* - X_t^\top \theta_*) \leq 2\beta_T \sum_{t=1}^T \|X_t\|_{V_{t-1}^{-1}} \leq 2\beta_T \sqrt{T \sum_{t=1}^T \|X_t\|_{V_{t-1}^{-1}}^2}$$

**Lemma 4** (Elliptical Potential Lemma). For any sequence  $\{X_1, \dots, X_T\} \in \mathbb{R}^{d \times T}$ , suppose  $V_0 = \lambda I$ ,  $V_t = V_{t-1} + X_t X_t^\top$ , and  $\|X_t\|_2 \leq L$ , then

$$\sum_{t=1}^T \|X_t\|_{V_t^{-1}}^2 \leq d \log \left( 1 + \frac{L^2 T}{\lambda d} \right) \quad \text{proofed in Lecture 6}$$

$$\mathbb{E}[\text{Regret}_T] \leq 2\beta_T \sqrt{T \sum_{t=1}^T \|X_t\|_{V_{t-1}^{-1}}^2} \leq 2\beta_T \sqrt{T d \log \left( 1 + \frac{L^2 T}{\lambda d} \right)}$$

# LinUCB: Regret Bound

**Proof.** With probability at least  $1 - 2\delta$ ,  $\mathbb{E}[\text{Regret}_T] \leq 2\beta_T \sqrt{Td \log \left(1 + \frac{L^2 T}{\lambda d}\right)}$

$$\begin{aligned} \mathbb{E}[\text{Regret}_T] &\leq 2\beta_T \sqrt{Td \log \left(1 + \frac{L^2 T}{\lambda d}\right)} \quad \beta_t = R \sqrt{2 \log \left(\frac{1}{\delta}\right) + d \log \left(1 + \frac{tL^2}{\lambda d}\right)} + \sqrt{\lambda} S \\ &\leq 2 \left( R \sqrt{2 \log \left(\frac{1}{\delta}\right) + d \log \left(1 + \frac{TL^2}{\lambda d}\right)} + \sqrt{\lambda} S \right) \sqrt{Td \log \left(1 + \frac{L^2 T}{\lambda d}\right)} \end{aligned}$$

Let  $\delta = 1/2T$ , then with probability at least  $1 - 1/T$ ,

$$\begin{aligned} \mathbb{E}[\text{Regret}_T] &\leq 2 \left( R \sqrt{2 \log \left(\frac{T}{2}\right) + d \log \left(1 + \frac{TL^2}{\lambda d}\right)} + \sqrt{\lambda} S \right) \sqrt{Td \log \left(1 + \frac{L^2 T}{\lambda d}\right)} \\ &= \tilde{\mathcal{O}}(d\sqrt{T}) \end{aligned}$$

□

## Improved Algorithms for Linear Stochastic Bandits

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### Abstract

We improve the theoretical analysis and empirical performance of algorithms for the stochastic multi-armed bandit problem and the linear stochastic multi-armed bandit problem. In particular, we show that a simple modification of Auer's UCB algorithm (Auer, 2002) achieves with high probability constant regret. More importantly, we modify and, consequently, improve the analysis of the algorithm for the linear stochastic bandit problem studied by Auer (2002), Dani et al. (2008), Rusmevichientong and Tsitsiklis (2010), Li et al. (2010). Our modification improves the regret bound by a logarithmic factor, though experiments show a vast improvement. In both cases, the improvement stems from the construction of smaller confidence sets. For their construction we use a novel tail inequality for vector-valued martingales.

### 1 Introduction

Linear stochastic bandit problem is a sequential decision-making problem where in each time step we have to choose an action, and as a response we receive a stochastic reward, expected value of which is an unknown linear function of the action. The goal is to collect as much reward as possible over the course of  $n$  time steps. The precise model is described in Section 1.2.

Several variants and special cases of the problem exist differing on what the set of available actions is in each round. For example, the standard stochastic  $d$ -armed bandit problem, introduced by Robbins (1952) and then studied by Lai and Robbins (1985), is a special case of linear stochastic bandit problem where the set of available actions in each round is the standard orthonormal basis of  $\mathbb{R}^d$ . Another variant, studied by Auer (2002) under the name "linear reinforcement learning", and later in the context of web advertisement by Li et al. (2010), Chu et al. (2011), is a variant when the set of available actions changes from time step to time step, but has the same finite cardinality in each step. Another variant dubbed "sleeping bandits", studied by Kleinberg et al. (2008), is the case when the set of available actions changes from time step to time step, but it is always a subset of the standard orthonormal basis of  $\mathbb{R}^d$ . Another variant, studied by Dani et al. (2008), Abbasi-Yadkori et al. (2009), Rusmevichientong and Tsitsiklis (2010), is the case when the set of available actions does not change between time steps but the set can be an almost arbitrary, even infinite, bounded subset of a finite-dimensional vector space. Related problems were also studied by Abe et al. (2003), Walsh et al. (2009), Dekel et al. (2010).

In all these works, the algorithms are based on the same underlying idea—the *optimism-in-the-face-of-uncertainty* (OFU) principle. This is not surprising since they are solving almost the same problem. The OFU principle elegantly solves the exploration-exploitation dilemma inherent in the problem. The basic idea of the principle is to maintain a confidence set for the vector of coefficients of the linear function. In every round, the algorithm chooses an estimate from the confidence set and an action so that the predicted reward is maximized, i.e., estimate-action pair is chosen optimistically. We give details of the algorithm in Section 2.

1

## Improved algorithms for linear stochastic bandits

Authors Yasin Abbasi-Yadkori, Csaba Szepesvári, David Pal

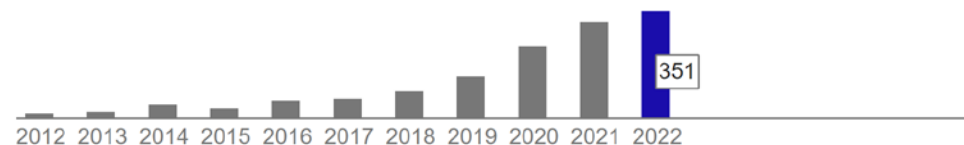
Publication date 2011

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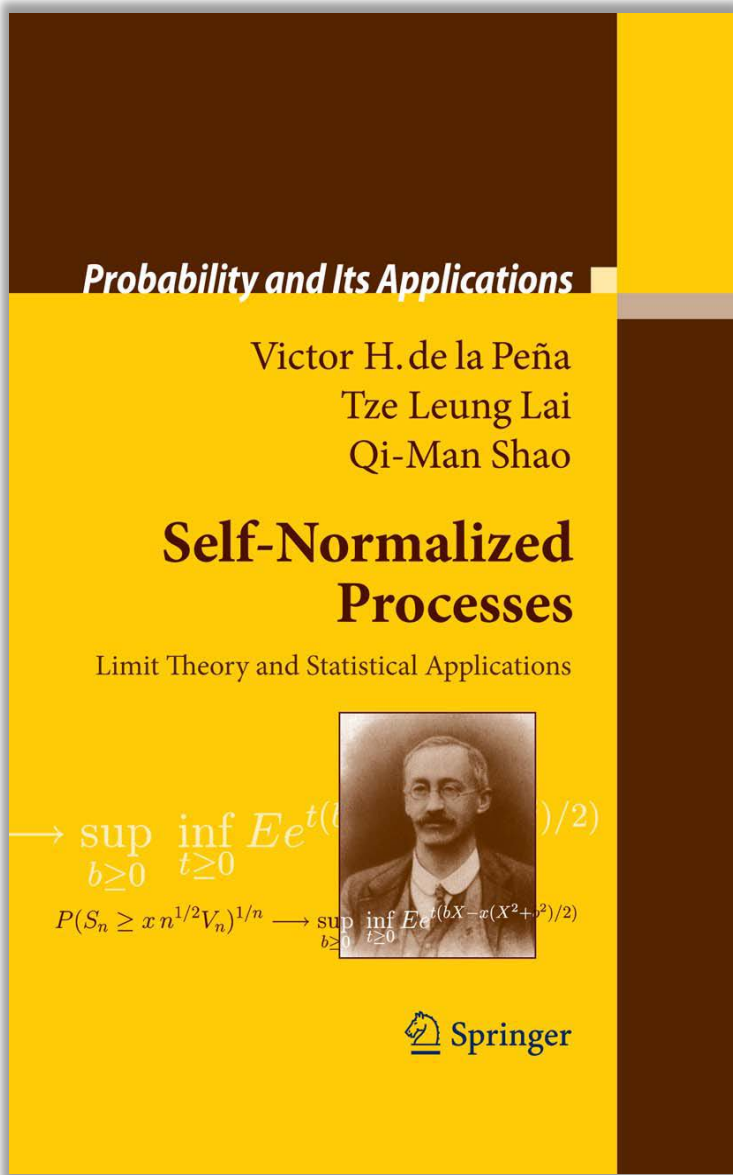
Description We improve the theoretical analysis and empirical performance of algorithms for the stochastic multi-armed bandit problem and the linear stochastic multi-armed bandit problem. In particular, we show that a simple modification of Auer's UCB algorithm (Auer, 2002) achieves with high probability constant regret. More importantly, we modify and, consequently, improve the analysis of the algorithm for the linear stochastic bandit problem studied by Auer (2002), Dani et al. (2008), Rusmevichientong and Tsitsiklis (2010), Li et al. (2010). Our modification improves the regret bound by a logarithmic factor, though experiments show a vast improvement. In both cases, the improvement stems from the construction of smaller confidence sets. For their construction we use a novel tail inequality for vector-valued martingales.

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Improved algorithms for linear stochastic bandits.  
In Advances in Neural Information Processing Systems  
24 (NIPS), pages 2312–2320, 2011.





## Self-Normalized Processes: Limit theory and Statistical Applications

Victor H. de la Peña, Tze Leung Lai,  
and Qi-Man Shao

Probability and Its Applications  
Series. Springer. 2009.



# Generalized Linear Bandits (GLB)

Extension: want to model Non-linear reward.

- Generalized linear model:  $r_t = \mu(X_t^\top \theta_*) + \eta_t$ 
  - Link function  $\mu : \mathbb{R} \mapsto \mathbb{R}$   $k_\mu$ -Lipschitz

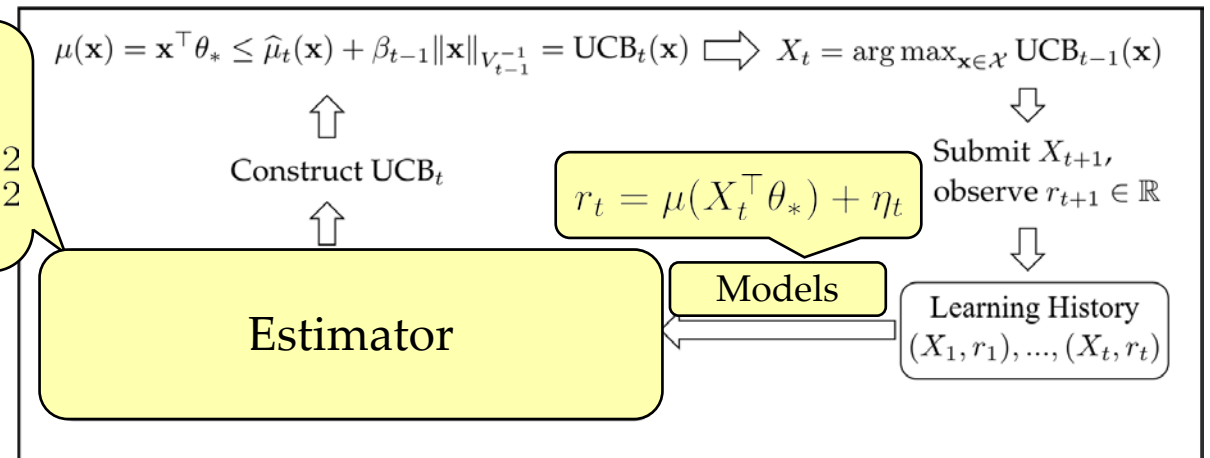
$$c_\mu = \inf_{\{\|\theta\|_2 \leq S, \mathbf{x} \in \mathcal{X}\}} \dot{\mu}(\theta^\top \mathbf{x}) > 0$$

Special cases: linear model:  $\mu(x) = x$ , logistic model:  $\mu(x) = \frac{1}{1 + \exp(-x)}$

- GLM-UCB

Maximum quasi-likelihood estimator

$$\hat{\theta}_t = \arg \min_{\theta \in \mathbb{R}^d} - \sum_{s=1}^{t-1} \log \mathbb{P}_\theta(r_{s+1} | X_s) + \frac{\lambda}{2} c_\mu \|\theta\|_2^2$$



# Advanced Topic: Best of Both Worlds

- Best of adversarial MAB:  $\mathbb{E}[\text{Regret}_T] = \max_{a \in [K]} \mathbb{E} \left[ \sum_{t=1}^T r_t(a) - \sum_{t=1}^T r_t(a_t) \right] \leq \mathcal{O} \left( \sum_{a: \Delta_a > 0} \frac{\ln T}{\Delta_a} \right)$
- Best of stochastic MAB:  $\mathbb{E}[\text{Regret}_T] = \mathbb{E} \left[ \sum_{t=1}^T \ell_t(a_t) \right] - \min_{a \in [K]} \sum_{t=1}^T \ell_t(a) \leq \mathcal{O} \left( \sqrt{TK} \right)$

Can one algorithm achieve the *best of both worlds*, without knowing whether the world is stochastic or adversarial?

- UCB: can get almost linear regret under adversarial setting.
- Exp3: can't have adaptive regret bound in stochastic case.1

⇒ Using OMD with *Tsallis entropy* regularizer.

Reference: Julian Zimmert, Yevgeny Seldin. [An Optimal Algorithm for Stochastic and Adversarial Bandits](#). AISTATS 2019.

# Advanced Topic: Bayesian Optimization

## Gaussian Process Optimization in the Bandit Setting: No Regret and Experimental Design

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### Abstract

Many applications require optimizing an unknown, noisy function that is expensive to evaluate. We formalize this task as a multi-armed bandit problem, where the payoff function is either sampled from a Gaussian process (GP) or has low RKHS norm. We resolve the important open problem of deriving regret bounds for this setting, which imply novel convergence rates for GP optimization. We analyze GP-UCB, an intuitive upper-confidence based algorithm, and bound its cumulative regret in terms of maximal information gain, establishing a novel connection between GP optimization and experimental design. Moreover, by bounding the latter in terms of operator spectra, we obtain explicit sublinear regret bounds for many commonly used covariance functions. In some important cases, our bounds have surprisingly weak dependence on the dimensionality. In our experiments on real sensor data, GP-UCB compares favorably with other heuristic GP optimization approaches.

### 1. Introduction

In most stochastic optimization settings, evaluating the unknown function is expensive, and sampling is to be minimized. Examples include choosing advertisements in sponsored search to maximize profit in a click-through campaign (Lizotte et al., 2007) or learning to rank (Chaloner et al., 2008). In this paradigm, the goal is to maximize the expected value of the unknown function by exploring and exploiting the available information.

as possible, for example by maximizing information gain. The challenge in both approaches is twofold: we have to estimate an unknown function  $f$  from noisy samples, and we must optimize our estimate over some high-dimensional input space. For the former, much progress has been made in machine learning through kernel methods and Gaussian process (GP) models (Rasmussen & Williams, 2006), where smoothness assumptions about  $f$  are encoded through the choice of kernel in a flexible nonparametric fashion. Beyond Euclidean spaces, kernels can be defined on diverse domains such as spaces of graphs, sets, or lists.

We are concerned with GP optimization in the multi-armed bandit setting, where  $f$  is sampled from a GP distribution or has low “complexity” measured in terms of its RKHS norm under some kernel. We provide the first sublinear regret bounds in this nonparametric setting, which imply convergence rates for GP optimization. In particular, we analyze the Gaussian Process Upper Confidence Bound (GP-UCB) algorithm, a simple and intuitive Bayesian method (Auer et al., 2002; Auer, 2002; Dani et al., 2008). While objectives are different in the multi-armed bandit and experimental design settings, our work generalizes stochastic linear optimization in a bandit setting, where the unknown function comes from a finite-dimensional linear space. GPs are nonlinear random functions, which can be represented in an infinite-dimensional linear space. For the standard linear setting, Dani et al. (2008)

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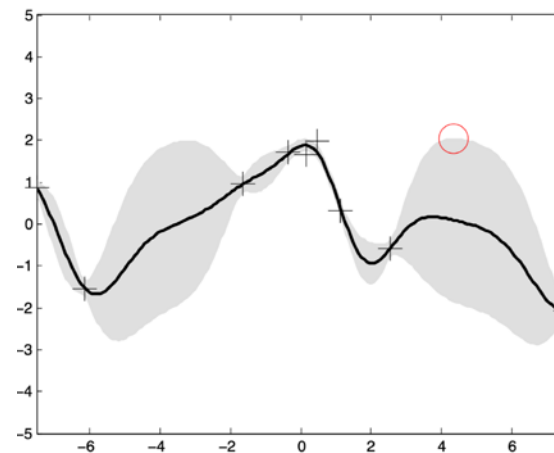
Reward function:  $r_t = f(X_t) + \eta_t$

$f(\mathbf{x})$  belongs to RKHS with  $k(\mathbf{x}, \mathbf{x}') = \sum_{m=1}^{|\mathcal{H}|} \varphi_m(\mathbf{x}) \varphi_m(\mathbf{x}')$

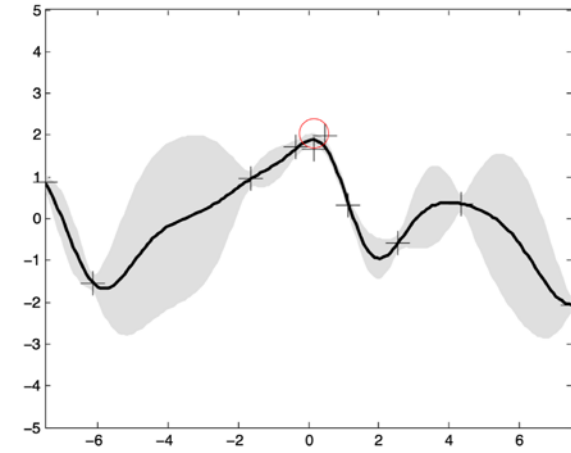
Rewrite  $f(x) = \sum_{m=1}^{|\mathcal{H}|} \theta_m \varphi_m(x) = \varphi(x)^\top \theta$

$\Rightarrow r_t = \varphi(X_t)^\top \theta + \eta_t$

*Linear bandits in RKHS*



Iteration  $t$



Iteration  $t + 1$

[Gaussian Process Optimization in the Bandit Setting: No Regret and Experimental Design](#). ICML 2010.

# Advanced Topic: Linear MDPs

## Linear MDPs

- Exists feature map  $\phi : \mathcal{S} \times \mathcal{A} \mapsto \mathbb{R}^d$ 
  - Such that:

$$r_h(s, a) = \theta_h^\star \cdot \phi(s, a), \quad P_h(\cdot | s, a) = \mu_h^\star \phi(s, a), \forall h$$

- Implies a low-rank assumption in large-MDP case

(Jin et al., 2020) Provably efficient reinforcement learning with linear function approximation

## UCB-VI for Linear MDPs

- In every round:
  1. Run Ridge regression for estimating the model
$$\hat{\mu}_h^n = \operatorname{argmin}_{\mu \in \mathbb{R}^{|\mathcal{S}| \times d}} \sum_{i=0}^{n-1} \|\mu \phi(s_h^i, a_h^i) - \delta(s_{h+1}^i)\|_2^2 + \lambda \|\mu\|_F^2.$$
$$\hat{\mu}_h^n = \sum_{i=0}^{n-1} \delta(s_{h+1}^i) \phi(s_h^i, a_h^i)^\top (\Lambda_h^n)^{-1}$$
  2. Construct the exploration bonuses
$$b_h^n(s, a) = \beta \sqrt{\phi(s, a)^\top (\Lambda_h^n)^{-1} \phi(s, a)},$$
  3. Run optimistic value iterations, and update greedy policy

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[Yu-Xiang Wang's course CS292F Lecture 10 Exploration IV: Linear MDP](#)

# History bits

- Bandit problems were introduced for the clinical trial design by **William R. Thompson** in an article published in 1933 [[Thompson, 1933](#)].

ON THE LIKELIHOOD THAT ONE UNKNOWN  
PROBABILITY EXCEEDS ANOTHER IN VIEW  
OF THE EVIDENCE OF TWO SAMPLES.

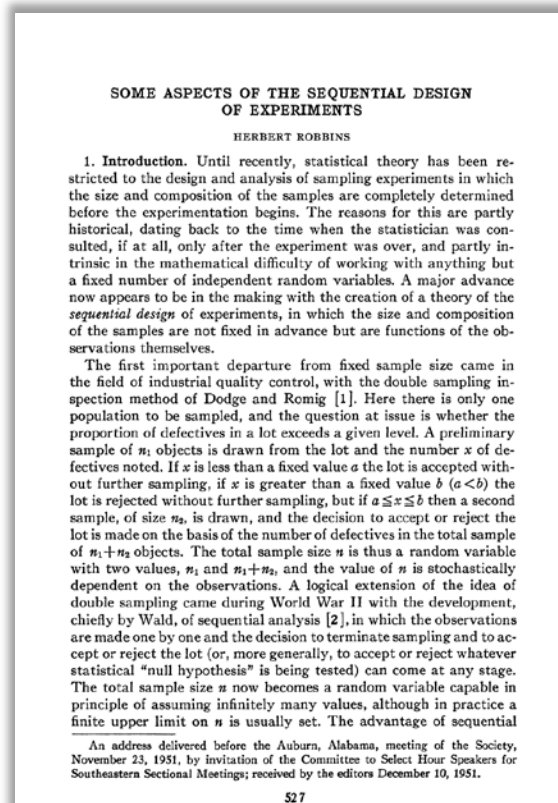
By WILLIAM R. THOMPSON. From the Department of Pathology,  
Yale University.



- Thompson Sampling (TS) was originally described in this paper but has been largely ignored by the artificial intelligence community.
- TS was subsequently rediscovered numerous times independently in the context of reinforcement learning.

# History bits

- Bandit problems were later formally restated in a short but influential paper [Robbins, 1952] and further developed in [Lai and Robbins, 1985].



H. Robbins. Some aspects of the sequential design of experiments. *Bulletin of the American Mathematical Society*, 58(5):527–535, 1952.

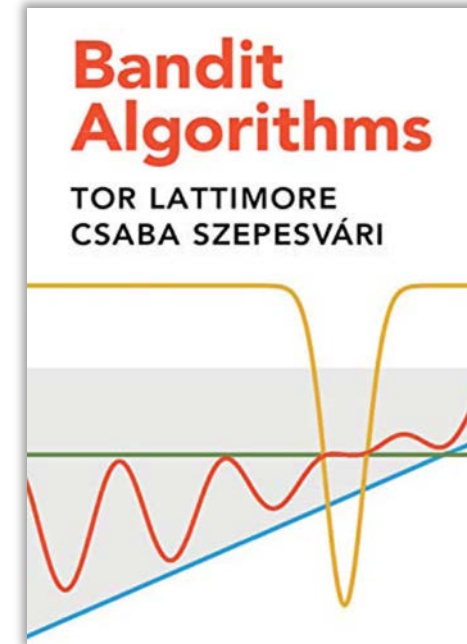
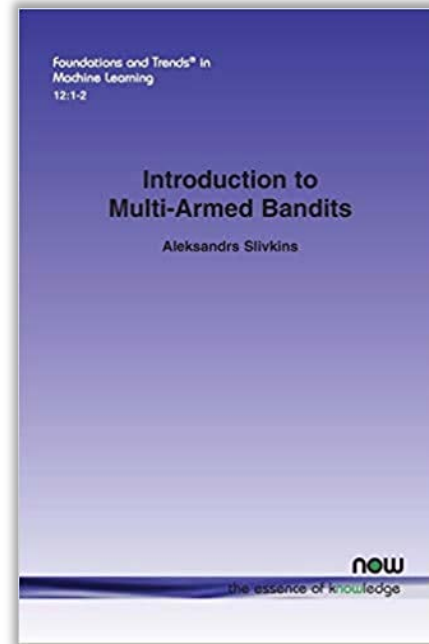
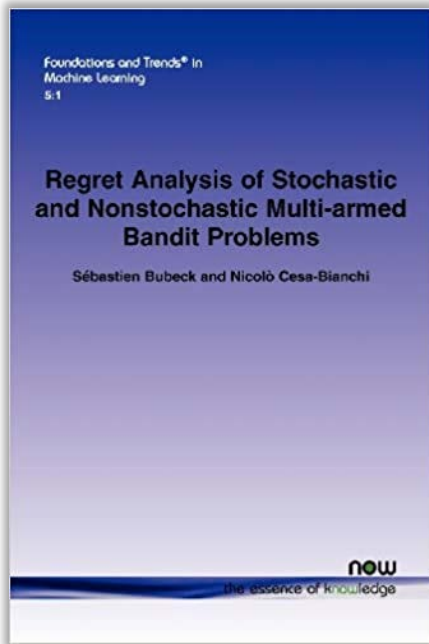


Herbert Ellis Robbins (1915 -- 2001)



# History bits

- Techniques developed in bandit problems have been applied in many areas, including machine learning, statistics, operational research, and information theory [Bubeck and Cesa-Bianchi, 2012; Slivkins, 2019; Lattimore and Szepesvári, 2020].



# Summary

